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Phenotypes of *Pinus sylvestris* are more coordinated under local harsher conditions across Europe

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ABSTRACT

1. Plant species that grow across environmental gradients show a range of trait expression, but traits do not vary independently. In fact, phenotypes are integrated expressions of multiple traits that covary due to trade-offs among functions and processes. Understanding trait covariation structures will ultimately help with predicting species' responses to change and developing management actions.

2. We measured variation and covariation (a proxy of phenotypic integration) among functional traits of *Pinus sylvestris* from paired populations across its European distribution. Populations within a pair were close enough to be in gene flow contact but located in contrasting environmental conditions across a local gradient. Measured traits represented three axes of variation (groups of traits) related to a tree's competitive ability and the trade-off between resource acquisition and conservation, namely plant size measures and stem and foliar traits.

3. Results revealed important intra- and inter-population trait variability. In particular, at the population level, trait means shifted across the climatic gradient mainly described by mean annual temperature. Moreover, we found a higher degree of trait covariation in populations under harsher environments (i.e. lower environmental suitability for the species). This pattern was consistent within population pairs, suggesting that higher trait covariation may be adaptive, being more coordinated in sites with harsher conditions. At larger spatial scales, we found a less conclusive pattern with a trend of increasing covariation at the northern edge of the species distribution. This result suggests that at larger scales different processes may be involved in the trade-off between the adaptive value of phenotypic covariation vs. its constraints on trait combinations that may limit plant's response capability.

4. Synthesis: trait covariation varies at different spatial scales, increasing under harsher conditions, and the robustness and repeatability of this pattern suggests its adaptive role for the species' responses to different environments.

RESUMEN

1. Las especies vegetales muestran diferencias en la expresión de sus rasgos a lo largo de gradientes ambientales; sin embargo, éstos no varían de manera independiente. El fenotipo es la expresión integrada de rasgos que covarían debido a los distintos procesos bajo las

condiciones a los que están sometidos. Comprender la estructura de covariación entre rasgos puede ayudar a predecir la respuesta de las especies frente a los cambios y a anticipar medidas de gestión frente a los mismos.

2. En este trabajo, hemos medido la variación y covariación (aproximación a la integración fenotípica) entre rasgos funcionales en individuos de *Pinus sylvestris* procedentes de pares de poblaciones a lo largo de su área de distribución europea. Las poblaciones dentro de una pareja estaban lo suficientemente cerca para asumir flujo genético entre ellas, pero localizadas en condiciones contrastadas en un gradiente local. En cada población medimos rasgos funcionales pertenecientes a tres ejes de variación relacionados con la capacidad competitiva del individuo y de la economía de los recursos (el equilibrio entre la adquisición y conservación de recursos), en concreto medidas de tamaño de la planta y rasgos medidos en tronco y hojas.

3. Los resultados revelaron una importante variabilidad de rasgos a nivel intra- e inter-poblacional. En concreto, la media poblacional de la mayoría de los rasgos varió a lo largo del gradiente climático definido principalmente por la temperatura media anual. Además, encontramos un grado de covariación de rasgos mayor en poblaciones bajo condiciones ambientales más duras (i.e. con menor idoneidad ambiental para la especie). Este patrón fue consistente entre parejas de poblaciones, sugiriendo que la covariación de rasgos puede ser adaptativa apareciendo fenotipos más coordinados en ambientes menos favorables. A escala espacial mayor, encontramos una tendencia con mayor covariación en el extremo norte del área de distribución. Este resultado asimétrico entre los extremos de la distribución, aunque no concluyente, indica que a grandes escalas otros procesos pueden intervenir en el balance entre el valor adaptativo de la covariación y la limitación que impone en la combinación de rasgos que pueden mermar la capacidad de respuesta frente a cambios.

4. Síntesis: la covariación de rasgos varía a diferentes escalas espaciales, incrementándose en condiciones estresantes. La robustez de dicho patrón sugiere su papel adaptativo en la respuesta de las especies frente a diferentes ambientes.

Palabras clave: covariación de rasgos, gradientes multi-escala, integración fenotípica, interacción planta-ambiente, pino silvestre, rasgos funcionales, respuesta adaptativa, variabilidad intraespecífica.

INTRODUCTION

Plant functional traits are an expression of a combination of individual use of resources and adaptation to a given environment (Grime, 1977; Westoby et al., 2002; Leimu & Fischer, 2008). Thus, within a species, traits may vary geographically as a response to environmental gradients at different scales, capturing individual (e.g. Benavides, Scherer-Lorenzen et al., 2019) and population responses (e.g. Abdala-Roberts et al., 2017; Luo et al., 2019). Previous works have studied intraspecific trait variability across wide study areas with changes in altitude and latitude (Fajardo & Piper, 2011; Laforest-Lapointe et al., 2014; Umaña & Swenson, 2019), demonstrating that trait–environment relationships are excellent candidates to predict species responses to environmental changes (Albert et al., 2010; Jung et al., 2014). They also showed that, although plant species may have the potential for highly plastic responses, strong environmental gradients impose ecological limits on these responses, which are specific for each trait (Valladares et al., 2007; Umaña & Swenson, 2019). In fact, most of these patterns come from studying responses of traits in isolation, or multiple traits but from a univariate perspective (but see e.g. Laforest-Lapointe et al., 2014; Umaña & Swenson, 2019; Benavides, Valladares et al., 2019).

Phenotypes are much more than the sum of multiple traits. Their integration reflects trade-offs among functions (Pigliucci, 2003; Bonser, 2006) and life history strategies (Chave et al., 2009; Westoby, 1998; Wright et al., 2004), and their responses under different environmental conditions (Maire et al., 2013). Thus, great interest is emerging to understand how traits covary to shape a phenotype, how these relationships vary across gradients and scales, and the adaptive role of trait combinations in different environments. Phenotypic integration is not a new concept within evolutionary ecology (Olson & Miller, 1958; Berg, 1960), although it has only been addressed intermittently due to the lack of a clear conceptual framework for empirical and theoretical studies (Pigliucci, 2003). Since the 2000s, a generation of studies has developed new state-of-the-art techniques, amassed evidence and identified challenges for future research on the topic (Pigliucci, 2003; Pigliucci, 2004; Hallgrímsson et al., 2009; Armbruster et al., 2014; Klingerber, 2014). Thereafter, phenotypic integration, strictly speaking, has been defined as the functional, genetic or developmental disposition or propensity to produce trait covariation, i.e. the correlation structure measured directly among traits (also known as statistical integration *sensu* Armbruster et al., 2014). The trait covariation structure (statistical integration) is then necessary to assess phenotypic integration, which is an intrinsic characteristic of organisms, influencing species evolvability by

producing correlated responses in other traits under selection (Wagner et al., 2007; Hallgrímsson et al., 2009).

The analysis of trait covariation is currently used in other disciplines within ecology to infer underlying mechanisms of species coexistence (Kraft et al., 2015; Dwyer & Laughlin, 2017a; Benavides, Scherer-Lorenzen et al., 2019), to describe the adaptive strategies of species (Boucher et al., 2013; Messier et al., 2018; Anderegg et al., 2018; Rosas et al., 2019; Carvalho et al., 2020; Damián et al., 2020) or to describe phenotypic diversity across gradients (between and within species), i.e. specific syndromes in response to the environmental variation (Wright et al., 2004; Laforest-Lapointe et al., 2014; Dwyer & Laughlin, 2017a; Umaña & Swenson, 2019). From these studies, a couple of general patterns have arisen. First, increasing stress usually leads to strong directional selection pressures (Boucher et al., 2013) that filter out unsuitable trait combinations and favour higher trait covariance (Westoby & Wright, 2006; Dwyer & Laughlin 2017a) and functional convergence (i.e. trait similarity) among coexisting species (Boucher et al., 2013; Dwyer & Laughlin, 2017a). Second, trait covariation structures are not always conserved across scales because responses to environmental drivers across gradients are trait-specific. For instance, the well-known global leaf economic spectrum (leaf traits closely correlated along a continuum, involving a trade-off between conservation and acquisition of resources, Wright et al., 2004; Reich et al., 2014) or the trade-off between hydraulic safety and efficiency (conduits' characteristics fostering either xylem transport or protection against embolism) are not always reproduced at fine spatial scales or within communities/populations (Gleason et al., 2015; Messier et al. 2017; Anderegg et al., 2018; Rosas et al., 2019; Carvalho et al., 2020).

Although emerging knowledge is helping to re-construct the phenotypic integration concept by including patterns of trait variation, it remains almost unknown whether trait covariation within species is maintained along gradients and scales, or whether it varies according to the specific selective pressure (see for intraspecific trait covariation across gradients at a single scale Anderegg et al., 2018 and Umaña & Swenson, 2019). Conserved patterns may reveal genetic and developmental constraints or consistent past selection towards certain trait combinations (Westoby et al., 1995). Alternatively, different trait means and covariation patterns may indicate the role of divergent selection (Ambruster et al., 2004), boosting or discarding certain trait combinations in given environments (Berg, 1960; Cheverud, 1984; Wagner et al., 2007; Damián et al., 2018) through the *plasticity of integration* (*sensu* Pigliucci, 2003), i.e. the integration's ability to vary under different conditions.

In this study, for the first time we analysed the phenotypic response of a tree species, *Pinus sylvestris*, in terms of variation and covariation of functional traits across a large part of its range and at different spatial scales. *Pinus sylvestris* is the most widely distributed pine species in the world and grows from Mediterranean mountains up to boreal forests in Fennoscandia and East Siberia, encompassing a wide range of climatic conditions (Carlisle & Brown, 1968). Given the range of environmental variation it experiences, the evaluation of changes in the phenotypic space and trait covariation for this species offers great potential for insight into the role played by trait covariation in the responses of the species. Specifically, we surveyed more than 500 trees from 20 populations across its European distribution (from Spain to Finland, and from Greece to the United Kingdom). Most populations were sampled in paired sites occupying locally contrasting conditions (such as elevation or water availability), but with significant potential for gene flow among them. This design followed Lotterhos and Whitlock (2015), which suggested that sampling populations in pairs across a species' range minimises the neutral genetic differences between sampled populations and therefore improves the power to detect signatures of selection. In our study, this novel design combining a wide spatial scale approach (European distribution) with local environmental gradients (pairs) allowed us to explore the effect of the environment on trait covariation. Thus, we analysed the distribution of trait variability, sought the main environmental drivers of this variability, and studied trait covariation (statistical integration) at two scales, i.e. across local gradients and species range. Our aim was to understand trait responses to environmental variation at both scales, and to search for trait covariation patterns indicative of adaptive strategy in this pine species. We assumed that the environmental suitability decreases towards the edges of the species distribution and, therefore, peripheral populations face more stressful conditions with respect to the species' optimum condition in central populations (Soulé, 1973; Sexton et al. 2009). Under this assumption, we hypothesised that trees in central populations would have more acquisitive phenotypes, i.e. larger trees with more productive needles (e.g. larger with higher N content, Wright et al., 2004; 2017), while those in peripheral populations would probably exhibit a more conservative strategy (e.g. smaller, less productive but more long-lasting needles). We also expected that more coordinated phenotypes would be expressed in harsher conditions for the species (reduced environmental suitability), where populations would have experienced more severe filters, and only individuals with a given set of trait values would have survived. Moreover, we hypothesised that in the core of the distribution local environmental heterogeneity (e.g. soil characteristics, high altitudes or water availability) may impose harsher

conditions on some populations in a similar way to which it occurs in peripheral populations (Soulé, 1973), driving convergence of conservative traits and more coordinated phenotypes.

MATERIAL AND METHODS

Study design and sites

Twenty populations of *Pinus sylvestris* L. from 11 regions (countries) across Europe were surveyed by partners in the GenTree consortium (http://www.gentree-h2020.eu/; Opgenoorth et al., 2021) (Fig. 1; Appendix A. Fig. S1, Table S1). Populations were selected from across the distribution, representing the range of environmental variation encountered by the species, but excluding heavily managed stands, or those disturbed by intense, very recent and obvious anthropogenic actions. Populations were representative of local environmental conditions, without being dominated by particular extremes, and included populations at the southernmost edge of the distribution (in Spain, Italy and Greece), and populations in the northernmost edge in Fennoscandia (Norway, Sweden and Finland) and the United Kingdom.

Except for two sites (Lithuania and Sweden), populations were sampled in local pairs, i.e. two distinct populations across a local gradient (such as elevation, water availability or day length), but with significant potential for gene flow among them based on previous studies. The geographic distances among populations were variable across the distribution, i.e. in southern sites topographic relief allowed strong environmental contrasts over short spatial scales; while in the north large spatial scales were required, particularly in Fennoscandia. However, F_{ST} values (measure of population differentiation due to genetic structure) among populations across the *P. sylvestris* range have been shown to be extremely low (Robledo-Arnuncio et al., 2004; Wachowiak et al., 2014; Pyhäjärvi et al., 2019). Thus, we assumed that our paired populations are likely in gene flow contact based on these studies that demonstrated the lack of genetic structure at local scales.

We classified each population within the local gradient (pair); distinguishing each as being under milder (M) vs. harsher (H) conditions (Fig. 1, Table S1). Local gradient selection sought contrasting conditions, but was not defined *a priori* by concepts of relative environmental suitability. Therefore, we combined two methods to achieve a robust classification: a climate-based criterion, which coarsely defines suitability based on the climatic space of the species distribution, and a tree-centric criterion based on tree performance (secondary growth data). The

latter determined which population within a pair performed better and we accordingly assigned this as the milder (M) site (Anderegg & HilleRisLambers, 2019).

We assessed the climatic space of the species by retrieving from CHELSA V1.2 (Karger et al., 2017a, b) the mean annual temperature (T) and annual precipitation (P) of each location obtained from the high-resolution tree occurrence records from the EU-Forest dataset (Mauri et al., 2017). We then defined density levels of *P. sylvestris* occurrence using a two-dimensional Kernel density estimate, and re-scaled the density levels to the probability density (between 0 and 1) and positioned the paired populations within the climate space. We assumed that the population under harsher conditions was the one with lower probability density, i.e. the population within a pair whose climatic conditions were less frequent within the entire distribution (Fig. 2a). With this criterion, some paired populations had similar density levels and could not be differentiated (i.e. Norwegian and Italian pairs).

In addition, we obtained the tree-ring width series (2006-2015) of surveyed trees from Martínez-Sancho et al. (2019, 2020). We selected the last 10 years when the competitive environment was closer to the sampling year (Anderegg & HilleRisLambers, 2019). We transformed tree-ring widths into annual basal area increments (BAI) using the *dplR* R package (Bunn et al., 2016), which is a better estimate of the overall tree growth than raw data (Biondi & Qeadan, 2008), and averaged the annual figures to get the mean $BAI_{2006-2015}$ for each tree. Then, we compared mean $BAI_{2006-2015}$ of paired populations using the *t-test*, after estimating age-, size- and local competition standardised growth values (Anderegg & HilleRisLambers, 2019), i.e. we used the residuals from the relationship $\log(\text{meanBAI}_{1985-2015}) \sim DBH + \text{tree age} + CI$ (Fig. 2b). *CI* is the competition index assessed at the tree level, which describes the competitive environment (see details in the following section *Environmental data*). Mean growth was significantly different for German and Finnish populations. For the rest, mean growth differences were more subtle but matched the classification derived from the climate-based criterion. The only exception was the Spanish pair, where the population with harsher climatic conditions performed slightly better, indicating that climate may not be the sole driver of environmental suitability here.

Field sampling and trait collection

Within each population, at least 25 adult trees were randomly selected and georeferenced, totalling 511 trees. Selected trees were dominant or co-dominant, healthy, and at least 30 m from the next selected tree (Opgenoorth et al., 2021). We chose ten of the most commonly used plant traits

(Table 1) from three trait dimensions important to define the global plant spectrum (Díaz et al., 2016) in order to capture comprehensive species' responses across the distribution (Westoby & Wright, 2006). The three trait dimensions included plant size, which reflects the ability to pre-empt light resources and disperse seeds; stem traits related to hydraulic safety and plant protection; and leaf traits, which balances the acquisition and conservation of resources, i.e. construction costs and growth potential (Wright et al., 2004; Chave et al., 2009; Díaz et al., 2016).

Specifically, in the field we measured tree height (H, m); diameter at breast height (DBH, cm); crown size, calculated with two perpendicular diameters of the crown projection (CP, m^2); bark thickness (Brk, mm), as the average of three to five measurements using a bark thickness gauge (Haglöf Barktax, Sweden); and trunk straightness (TSt) following the scale: 1) moderate or strong bends, 2) slight to moderate bend in different directions, 3) fairly straight (in one direction slightly crooked) and 4) absolutely straight (Opgenoorth et al., 2020). In addition, one wood core was extracted at breast height (1.3 m), perpendicular to the slope direction to avoid reaction wood, to assess wood density (WD, g/cm^3) (Martínez-Sancho et al., 2019, 2020) and one branch from the top of the crown to assess leaf traits (Benavides et al., 2021). For the latter, we collected ten needles of the last complete growing season from each branch (totalling 5110 needles), scanned them and measured their projected area (LA, mm^2) using WinFOLIA (Regent Instruments Inc., Canada), considered a proxy of the total leaf area (Pérez-Harguindeguy et al., 2013). They were then oven-dried at 60 °C for 72 h and weighed for dry mass, and we estimated specific leaf area (SLA, mm^2 mg^{-1}). We obtained leaf morphological trait per tree averaging the figures of the ten individual needles (Benavides et al., 2020).

Finally, for a subset of at least 14 individuals in each population (284 in total), dried needles were ground and analysed for nitrogen content (LNC; %) and isotope ^{13}C content in plant material (δ^{13}C, reported relative to V-PDB, ‰). Leaf collection, storage, processing and morphological trait measurement followed Pérez-Harguindeguy et al. (2013), while the chemical analyses were carried out using gas chromatography-combustion isotope ratio mass spectrometry (GC-C-IRMS) at the UC Davis Stable Isotope Facility (https://stableisotopefacility.ucdavis.edu/) (Benavides et al., 2020, 21). Two of the leaf traits, SLA and LNC, are directly related to the leaf economics spectrum (acquisition and conservation of resources trade-off), while the other two, LA and δ^{13}C, reflect aspects related to gas exchange in leaves and water use efficiency, respectively. Finally, we also compiled tree age data from Martínez-Sancho et al. (2019, 2020).

Environmental data

At different spatial scales, we selected environmental variables relevant for tree performance (Opgenoorth et al., 2020). At the population scale, 26 climatic variables were retrieved from CHELSA (resolution 30 arcsec, c.a. 1 km^2); and 16 topographic variables derived from the European digital elevation model with 25 m spatial resolution (EU-DEM v. 1.1 from the Copernicus program; https://land.copernicus.eu/).

Moreover, we estimated a competition index (CI) at the tree level (Lorimer, 1983), which represents the degree of competition for available resources with surrounding neighbours. It was calculated using the five nearest trees within a maximum radius of 15 m around each surveyed tree as $CI = \sum_{i=1}^{5} (dbh_i/dbh)/dist_i$, where dbh is the diameter at breast height of the subject tree, dbh_i the diameter at breast height of the competitor tree i and $dist_i$ the distance between the subject tree and competitor tree i. This index assumes that the net effects of neighbouring trees vary as a direct function of the size of the neighbours and as an inverse function of the distance.

Statistical analyses

We explored changes in the phenotypic space of *P. sylvestris*, analysing trait shifts and drivers across scales, and implemented a multi-level approach to phenotypic integration (*sensu* Armbruster et al., 2014; Klingenberg, 2014) analysing the trait covariation structure across scales within its distribution.

Trait variance partitioning and factors affecting trait variation

First, we ran linear random models for each trait to examine the distribution of the variance of each trait across scales (Zuur et al., 2009). We included *population* as a random effect that accounts for the variability among populations, representing the residual variance the variability within populations. We repeated the analysis with *pair* nested within *region* (country) for those populations sampled in pairs to partition the variability among populations driven by the local gradient. Trait values were Box-Cox transformed (including log-transformation as a particular case of Box-Cox transformation), as appropriate, to optimise normality of the residuals.

Next, we tested the effect of environmental factors on trait values (eq. 1). Previously, we ran two principal component (PCAs) and correlation analyses, one for climatic and another for topographic variables, to select those variables that orthogonally explained more variance (Appendix A, Figs. S2, S3; Tables S2, S3). The final selection included mean annual temperature

(T, ºC); annual precipitation (P, mm); temperature of the wettest quarter (Twet, ºC), potential total solar radiation (rad, GJ m^{-2}), which varied among populations; slope (sl, º); tree age (years), which varied both among and within populations; and the competition index (CI), for which more than 75% of the variance was within populations (Appendix A, Fig. S4). We also included the quadratic term of T to consider a common non-monotonic response typical of physiological processes along the thermal gradient.

$$Trait \sim age + CI + sl + rad + T + T^2 + P + Twet + (1|population) \text{ (eq1)}$$

We fitted all possible linear mixed models (eq. 1) for each trait, and ranked them using the Akaike Information Criterion (AIC), where 'better' models achieved an improvement of at least two units AIC over the next one (Burnham & Anderson, 2002), and followed the principle of parsimony to prioritise the simplest model (Appendix C, Table S4). We assessed the significance of factors by comparing models with and without each factor selected in the best model using the *likelihood ratio test* (Zuur et al., 2009).

Trait covariation across gradients

Then, we characterised overall and intrapopulation structures of trait covariation and analysed its variation across the local (mild *vs*. harsh) and rangewide (climate) gradients. Previous analyses showed an important effect of tree age on size plant and stem traits (Table 2). Thus, we removed the effect of age prior to trait covariation analysis by using the residuals from the relationship between each trait and age (trait ~ age). In doing so, we accounted for the potential confounding effect of ontogeny on covariation patterns. Then, we obtained the trait correlation matrices with Pearson correlation analyses (except for trunk straightness, for which we used Spearman correlations, as it is an ordinal variable). We used permutation tests to evaluate the statistical significance of each pairwise correlation at the population and species levels. Traits within individuals were untied and shuffled 1000 times. From each randomisation, we assessed pairwise trait correlations, generating a null distribution, and we extracted a p-value associated with the observed correlations. At the species level (i.e. with pooled data), in each randomisation, traits from 25 individuals (instead of the complete sample size, i.e. > 500) were extracted to generate the null distribution and get significant coefficients comparable to those obtained at the population level. Moreover, very small, meaningless but highly significant correlation coefficients, typically obtained from large-sized samples, were discarded (Aggarwal & Ranganathan, 2016). Thus, all significant pairwise correlations (significantly stronger from what is expected by chance) had

values over |ρ| ≥ 0.3. We graphically represented the covariation structures running a network analysis with the correlation coefficients for each population. We generated undirected circle networks, with traits as nodes, and the significant correlations amongst them as edges.

Next, we calculated two quantitative measures of trait covariation degree, i.e. the *edge density* (ED), assessed as the ratio between the number of significant correlations (edges) and all possible pairwise trait combinations, and the *functional variability shape* (FS) (*sensu* Boucher et al 2013), i.e. the variance of the eigenvalues of the trait correlation matrix which considered a phenotypic integration index (Wagner, 1984; Cheverud et al., 1989). If traits are uncorrelated within populations, eigenvalues will be similar and have low variance, while eigenvalues will show high variance when traits are correlated (Dwyer & Laughlin, 2017b). Finally, we fitted linear models to analyse the effects of the climate (T, T^2, P and Twet) and the local gradient (H vs M) on these two metrics. We fitted all possible models, ranked them using the Akaike Information Criterion corrected for small sample size (AICc), and assessed the significance of selected factors by comparing the model with and without factors selected in the optimal model using the *likelihood ratio test* (Zeileis & Hothorn, 2002).We also explored ED and FS variation across latitude (see Supporting Information).

All statistical analyses were performed in R (version 3.6.1, R Core Team, 2019), using the *stats* package for correlation analysis and linear models, *lme4* package (Bates et al*.,* 2015) and *MuMIn* (Barton, 2020) for linear mixed models, *lmtest* package for likelihood ratio test in linear models (Zeileis & Hothorn, 2002), *AID* package for Box-Cox transformation (Asar et al*.,* 2017), *igraph* package (Csardi & Nepusz, 2006) for the network analysis, and *ggplot2* package (Wickham, 2016) for graphics.

RESULTS

Distribution of trait variation

Trait values varied not only across regions but also between pairs and within populations (Fig. 3; Appendix B, Fig. S5). Plant size traits were the most variable with the highest coefficients of variation (Fig. 3), while WD and leaf isotopic signature were the most stable traits.

The distribution of trait variance was also trait-dependent (Fig. 3). Plant size differed more among populations (over 60% of trait variance for the three traits) than within populations, leaf traits varied similarly at both scales (among and within populations), while stem traits had higher

variance within populations that among them. All traits showed non-negligible variance between pairs (at local gradients), except for $\delta^{13}C$ (Fig. 3).

Similarly to trait variance partitioning, linear models showed that there are factors acting on trait means at different scales (Table 2; Fig. 4; Appendix B Figs. S4, S6). Age and competition (CI) significantly affected traits at the tree level, accounting for variance within populations for stem and plant size traits (Table 2). We also found a significant climatic signal for all traits except for trunk straightness and SLA (Table 2; Figs. 4, S6), with a non-linear effect of mean annual temperature on plant size, stem and leaf traits. They all presented lower means at both edges, and the highest value in the centre-warmer half of the thermal range (Figs. 4, S6). Moreover, the mean temperature in the wettest quarter of the year had a positive effect on H, while precipitation had a negative effect on H and LA. Potential solar radiation derived from the topography had a positive effect on $\delta^{13}C$, Brk and LNC, but a negative effect on WD (Table 2).

Integration of traits along gradients

Both metrics, edge density (ED) and functional variability shape (FS), were correlated in our study populations ($\rho = 0.67$), showing the same trends. Considering all populations, the phenotypic space of *P. sylvestris* was poorly coordinated, with an edge density of only 6.67% and functional variability shape of 0.49 (Appendix B, Fig. S7). Only three traits covaried (DBH, CP and Brk), three pairwise correlations that appeared in most of the populations (Fig. 5, Fig. S8). However, this figure increased when we calculated ED within each population, ranging from 17.8% up to 63.9% (FS from 0.62 to 2.2). This range reflected important differences in the trait correlation structure that differed not only quantitatively (number of significant correlations) (Fig. 5), but also qualitatively (traits more frequently involved in trait covariations) (Fig. S8). In fact, only six trait pairs were frequently correlated and showed a mean correlation coefficient over $|\rho| > 0.3$ (significance threshold), namely correlations between DBH and CP, Brk and H, between CP and Brk, and between SLA and $\delta^{13}C$ and LA.

The variation of edge density and functional variability shape across the rangewide climatic gradient (and latitudinal) showed significantly greater values in northern populations where conditions were cooler (significance levels $p<0.01$ for T (ED and FS), $p<0.1$ for T^2 (ED), Fig. 6; see Fig. S9 results in terms of latitude). This asymmetry was nonetheless driven by the values obtained in the northernmost population in Finland, diminishing the effect of temperature or latitude when this population was omitted (Appendix B, Fig. S10). Moreover, we found a

significant effect of the local gradient with consistently higher trait covariation in populations under harsher conditions (estimated marginal means ED=29.9%, log(FS)= -0.0173) compared to those under milder conditions (estimated marginal mean ED=21.5%, log(FS)= -0.2186) (p<0.05 for ED and p<0.1 for FS, Fig. 6). We did not find a significant interaction among the local gradient and mean annual temperature (or latitude).

DISCUSSION

Phenotypic variation within the widely distributed conifer *Pinus sylvestris* has an important intrapopulation component that, irrespective of other factors such as genetic variability among individuals or differences in their microhabitat, was significantly driven by ontogeny (age) and competition with neighbours. At the interpopulation level, temperature strongly influenced mean trait values, with non-linear effects that showed smaller trees with smaller and less productive (less N content) needles at both edges of the thermal range of the distribution, i.e. close to the southern and northern edges of the distribution. We also found a systematically higher degree of trait covariation (higher statistical phenotypic integration) in populations under harsher conditions in local gradients, and an increasing trend at the northern edge of the distribution, which showed that integration varies at different spatial scales (Fig. 6). Despite finding qualitative differences among populations without same patterns in covarying traits, the consistently higher trait covariation under harsher conditions across the distribution, entailing 'tighter' phenotypes, suggests that it has an adaptive value.

Trait covariation across local gradients

At the beginning of this century, the concept of plasticity of integration was coined (Murren, 2002; Pigliucci, 2003; Piggliucci, 2004), recognising that, as for individual traits, trait correlation could vary dramatically in response to environment or spatial scale (Schlichting, 1986; Nicotra et al*.,* 1997; Anderegg et al*.,* 2018; Rosas et al*.,* 2019). In fact, some experimental studies connected stress to 'tighter' phenotypes probably seeking higher coherence or optimality under adverse conditions (Schlichting, 1986; 1989; Chapin, 1991; Damián et al*.,* 2020, but see Pigliucci & Kolodynska, 2002a,b; Pigliucci & Kolodynska, 2006). Our sampling design, with population pairs assumed to be genetically connected due to spatial proximity and highly efficient long-distance gene flow by wind dispersion (a robust assumption given the many studies on the topic in this widely studied species, e.g. Robledo-Arnuncio et al*.,* 2004; Wachowiak et al*.,* 2014; Pyhäjärvi et

al*.,* 2019), drew outcomes in agreement with the expected positive relationship between level of stress and phenotypic coordination. High levels of gene flow between paired populations mean that both share a common gene pool, and so divergence in trait means can be assumed to reflect differences in local adaptive pressures (Conner & Hartl, 2004; Holderegger et al., 2006). Therefore, the increasing edge density and functional variability shape in harsher conditions suggest that trait covariation responds to different environments. It is remarkable that this pattern is found across the entire distribution, between paired populations subjected to different gradients (elevation, water availability, day length), and is a strong indication that higher trait covariation may be adaptive, in accordance with recent work exploring covariation in leaf traits (Damián et al*.,* 2020). In other words, 'tighter' phenotypes appear in more stressful sites to avoid the high cost of maintaining certain and rare trait combinations (Westoby & Wright, 2006; Dywer & Laughlin, 2017a). However, our study lacks the means to discern the underlying mechanism, i.e. whether these 'tighter' phenotypes are expressing local adaptation (Conner & Hartl, 2004; Kawecki & Ebert, 2004) or adaptive plasticity is operating (Pigliucci, 2001; Stamp & Hadfield, 2020). Regardless, the consistency and robustness of this pattern across contrasting climatic conditions and selective pressures clearly suggest an adaptive value for trait covariation in this species (Damián et al*.,* 2020).

Rangewide trait covariation

We hypothesised that the pattern of response to stress at relatively small spatial scales might also occur at larger scales among climatically different regions. We based this on the central-periphery hypothesis (Soulé, 1973; Abeli et al., 2014) that suggests that at the extreme edges of a species distribution, the number of trait combinations should be limited (higher trait covariation) due to the high cost of certain values, resulting in more integrated phenotypes (Westoby & Wright, 2006; Dywer & Laughlin, 2017a). It is notable that connection between increasing trait covariation with stress has not always been captured (Pigliucci & Kolodynska, 2002a, b; 2006; Bouchar et al*.,* 2013), showing the complexity of the effect. Our results showed a greater trait covariation at the northern edge of the species distribution (cold edge) supporting apparently our hypothesis (Fig. 6). However, this outcome should be taken with caution as this latitudinal and climatic pattern is mainly driven by the trait covariation obtained in the northernmost population in Finland (Table S1; Fig. S10). Whether these values represent outliers or the threshold that triggers higher

coordination of traits would need further investigation with more surveys in peripheral populations.

At the warm (southern) edge we did not find this increasing pattern opposing our expected positive relationship between phenotypic integration and less suitable conditions at the margins (Fig. 6a). Trait covariation is considered to be a trade-off between adaptation and constraints (Pigliucci, 2003; Merilä & Björklund, 2004). As previously mentioned, natural selection may favour certain combinations of traits and, therefore, trait coordination can be adaptive (Damián et al., 2020). However, certain relationships among traits may limit their plasticity (Gianoli & Palacio-López, 2009) becoming a threat for adaptation under unpredictable conditions (Jernigan et al., 1994). In Mediterranean mountains, the climate frequently encompasses extreme events of summer drought and cold winters (Pereira & Chaves, 1995), hence populations should have a higher disposition for phenotypic integration (Dwyer & Laughlin, 2017a; Michelaki et al.,2019) compared to those thriving on summer drought-free, more benign areas (Baraloto et al.,2010; Fortunel et al., 2012). However, the unpredictable nature of the Mediterranean climate and a complex topography, which creates a high heterogeneity of environmental conditions over short distances, can underlie lower trait covariation than expected, where trait covariation can be considered a disadvantage rather than a benefit (Jernigan et al., 1994). Also, our southern populations belong to fragmented areas of the distribution southern edge (Fig. 1). This means that they could be affected by pernicious effects related to both genetic drift and genetic isolation, which may in turn cause a selection inefficiency (Sexton et al., 2009). In contrast, northern populations lie environmentally closer to the centre of the distribution and likely have much larger population sizes. Consequently, they may exhibit better responses to natural selection (Kirkpatrick & Barton, 1997), producing more 'optimised' phenotypes based on higher degree of trait covariation (Bouchar et al., 2013).

Trait variation across environmental gradients

Most of the trait clines showed a very similar pattern across the species distribution (Fig. 4; Fig S6). In particular, we detected non-linear responses along the temperature gradient that reflects lower climatic suitability at both edges of the range, which translated into bigger trees with larger and more nutrient-rich needles in the central-warmer part of the range (central Europe). At the northern edge, *P. sylvestris* is normally limited by low temperatures and a shorter length of the growing season; while at the southern edge, it is mainly limited by a combination of high summer

temperature and drought (Carlisle & Brown, 1968; Castro et al., 2004). In our study populations, the pattern along the thermal range matched expectations: shorter trees with smaller leaves with lower nutrient content at both extremes (Reich et al., 2014; Wright et al. 2017). In parallel, the precipitation range translated into wetter locations being less suitable, with decreasing tree size and smaller leaf area. This trend was mainly driven by trait values in populations in Scotland, which represents the north-western edge of the range. There, *P. sylvestris* populations are considered to be in long-term decline, preceding human influence (Bennet et al., 1984), which is thought to be partially due to increasing precipitation leading to the formation of bogs.

Traits related to plant size were most tightly correlated throughout the range; and they also correlated with bark thickness, which agrees with previous work in two populations of *P. sylvestris* (Carvalho et al., 2020). Allometric constraints determine that taller trees normally have a thicker trunk and larger crown, associated with thicker bark that confers higher mechanical support and defence (Niklas, 1999; Paine et al., 2010). Among the leaf traits, the most widespread covariations were the negative correlations between SLA and LA and between SLA and $\delta^{13}C$. The latter was also found in Carvalho et al. (2020) and reflects that individuals investing more in long-lasting, larger and denser needle tissue (lower SLA) have higher water use efficiency. The leaf economic spectrum (LES) describes gradual strategies for plants running from low-cost, short-lived leaves with rapid return of carbon and nutrients (i.e. high SLA and N content) to costly, long-lived leaves with slow returns (low SLA and N content) holding low photosynthetic rates (Wright et al., 2004). However, our data do not fit this pattern, as we found no clear relationship between N content and SLA, outcome that agrees with the different sensitivity to climate shown by both traits in this study and elsewhere (Rosas et al., 2019; Umaña & Swenson, 2019). In addition, we found a negative covariation between SLA and LA, which disagrees with studies showing that larger leaves are normally less dense at the intraspecific level (Martin et al., 2017; Benavides, Scherer-Lorenzen et al., 2019). Variation of leaf size in our climatic gradient followed the global pattern described by Wright et al. (2017), i.e. smaller leaves at higher latitudes and in drier and warmer places. However, SLA is a rather complex trait, especially in evergreens (Lusk et al., 2008). It integrates the trade-off between reduction costs and prolongation of leaf lifespan in resource-limited environments determined by multiple factors, including climate (precipitation, drought or minimum temperatures), light availability or soil characteristics (Wright et al., 2004; Gonzalez-Zurdo et al., 2016; Gong & Gao, 2019). Therefore, this mismatch (the unexpected negative correlation between SLA and LA) may reflect a complex response to different factors or

scales involving other associated traits that we did not measured (e.g. leaf thickness, leaf lifespan) (Wilson et al., 1999).

CONCLUSIONS

The multivariate approach to the phenotype, recognising the importance of trait covariation, is emerging as a necessary milestone to comprehend species responses to environmental variation. Here, we analysed the distribution of trait mean and variability across the distribution of the widely distributed *Pinus sylvestris*, and found idiosyncratic covariation structures with traits covarying differently across scales. More importantly, we found a systematic increase of trait covariation in more stressful sites at a regional scale, suggesting that regardless of the traits involved, there is a pattern towards more coordinated and more efficient responses –'tighter phenotypes'- in such conditions, likely related to adaptive processes.

AUTHORS' CONTRIBUTIONS

RB, FV, BC, AE, SM, EMS conceived the ideas and designed methodology. SC led the selection of sites. RB, BC, CCB, SM, FV, PF collected field data. PF and EMS provided wood density and age data. RB, BC, CCB assessed leaf traits. RB analysed the data, led the writing of the manuscript. All authors contributed significantly with comments to the different versions of the manuscript and gave the approval for publication.

DATA AVAILABILITY

Wood density, tree age (Martínez-Sancho et al. 2019), leaf data (Benavides et al. 2020) and the rest of phenotypic and environmental data (Opgenoorth et al. 2020) are all available in figshare.

SUPPLEMENTARY INFORMATION

Appendix A. Information about the study sites.

Table S1. Description of the study sites

Fig. S1. Variation in mean annual temperature (T), precipitation (P) and mean temperature of the wettest quarter (Twet) of the study populations across latitude

Fig. S2. PCA biplots (left) and loadings (right) based on analysis of data for 26 climatic variables from each population.

Fig. S3. PCA biplot (left) and loadings (right) based on analysis of data for 16 topographic variables from each population.

Table S2. Spearman correlation coefficients among climatic variables in the study sites

Table S3. Spearman correlation coefficients among topographic variables.

Fig. S4. Distribution of trait variance of environmental variables.

Appendix B. Surveyed traits: distribution, mean values and covariation.

Fig. S5. Boxplots of traits in each population, a) plant size and stem traits, b) leaf traits.

Fig. S6. Traits versus climatic variables.

Fig. S7. Trait correlations of *Pinus sylvestris* phenotype.

Fig. S8. Distribution of pairwise correlation coefficients.

Fig. S9. Effect of local gradient and latitude on a) edge density (ED); and on b) (log) functional variability shape (FS).

Fig S10. Effect of local gradient and temperature/latitude on a) edge density (ED) and b) functional variability shape (FS) without northern-most population (FI_18).

Appendix C. Model selection

Table S4. Model comparison of variables affecting trait variation for each trait.

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Figures captions

Fig. 1. **Location of study populations of Pinus sylvestris across Europe**. In most regions, pairs of populations located in contrasting conditions were surveyed. Black circles represent the population within a pair that was classed as occupying harsher conditions; triangles represent the population within a pair that was classed as occupying milder conditions, asterisks represent unpaired locations. Grey area represents the species distribution obtained from EUFORGEN (www.euforgen.org).

Fig. 2. Local gradient definition. a) Distribution in a climatic space of occurrence data of Pinus sylvestris across Europe using 2D kernel density estimates to define the probability density. Grey points were obtained from Mauri et al. (2017) and red line contains 95% of the observations. The probability density of each population in the climatic space is shown in the table on the right, and determines the classification within pairs, assigning M (populations under more suitable conditions) to those located in a climate with a higher probability density of occurrence, and H (populations under harsher conditions) to those with smaller probability density of occurrence. b) Mean basal area increments -BAIst- (and standard deviation) standardised by age, size and competitive environment (unitless) of each population (x axis, ordered by latitude) for the period 2006-2015. Triangles represent the higher values within pairs (populations under more suitable conditions, M), and circles the lower values (populations within pair under harsher conditions, H). Populations in red represent a classification mismatch between both criteria.

Fig. 3. Distribution of trait variability. The partitioning of variance revealed the component within populations (dotted purple bar) and among populations (green bars), which in turn can be partitioned into variance among regions (green striped bars) and paired populations in a local gradient (solid green bars). The coefficients of variation (CV with pooled data, CVm within-population mean) of each trait are shown. H: tree height, DBH: diameter at breast height, CP: crown projection size, Brk: bark thickness, WD: wood density, TSt: trunk straightness, LNC: leaf N content, SLA: specific leaf area, LA: leaf area, δ13C: isotopic signature of 13C. Variance partitioning analysis was performed on both raw and transformed data, which resulted in very close outcomes. This figure shows the results only obtained with raw data.

Fig. 4. Traits versus climatic variables. Plots show linear mixed models as described in Table 2, and selected according to AIC and parsimony criteria. a) Traits related to plant size, b) stem traits, and c) leaf traits. Only significant effects are shown. T: mean annual temperature (ºC), P: mean annual precipitation (mm), Twet: temperature of the wettest quarter (ºC), H: height (m), DBH:

diameter at breast height (cm), CP: crown projection area (m2), Brk: bark thickness (mm), WD: wood density (g cm-3), LNC: leaf N content (%), LA: leaf projected area (mm2). Grey area represents the 95% confident intervals, which reflect only the variance of the fixed effects, not the random effects (populations), and points indicate population means (see Fig. S6 with tree level observations).

Fig. 5. Trait correlation networks in 20 *Pinus sylvestris* populations across 11 regions. Traits belong to three dimensions: plant size (blue), stem traits (brown) and leaf traits (purple). Local gradient is given as M (milder environment) vs H (harsher environment). Black lines represent positive correlations, while red ones are negative. Line colour intensity shows the strength of the correlation. H: tree height, DBH: diameter at breast height, CP: crown projection area, Brk: bark thickness, TSt: trunk straightness, WD: wood density, LNC: leaf N content, SLA: specific leaf area, $\delta^{13}C$: isotopic signature of ^{13}C, LA: leaf area. Where 'WD' is given in light grey means that we lack this trait in those populations. Plant size and stem traits were standardised by age before the correlation analysis. Correlation analyses were conducted with data surveyed in 25 trees per population

Fig. 6. Effect of local gradient and mean annual temperature (T) on a) edge density (ED); and on b) (log) functional variability shape (FS). ED is the ratio between the number of significant correlations and all possible pairwise trait combinations, and FS is the variance of the eigenvalues of the trait correlation matrix. On the top, the expression of the optimal models and their R2. Whiskers and grey areas represent the 95% confidence intervals. Points represent the observed values. M: populations under milder conditions; H: populations under harsher conditions.

Table 1. Overview of functional traits measured in each individual.

	Trait	Units	Trait description	Trait functions
Plant size traits	Height (H)	m	Distance from the soil surface to the top end of the crown	Competitive vigour to capture light, competing either in the vertical or horizontal plane
	Diameter at breast height (DBH)	cm	Diameter of the stem at breast height	
	Crown projection area (CP)	m^2	Area estimation of the crown projection, assuming elliptical areas and using two perpendicular diameters of the crown	
Stem traits	Bark thickness (Brk)	mm	Average of 3-5 measures of bark thickness using a bark gauge	Cambium protection and mechanical support
	Trunk straightness (TSt)	unitless	Categorical variable, from 1 to 4 meaning absolutely straight	Competitive vigour to capture light and mechanical support
	Wood density (WD)	g cm^{-3}	Ratio of wood dry mass (g) per displaced volume (cm^3)	Mechanical resistance, water storage in the trunk, hydraulic safety, growth-survival trade-off
Leaf traits	N content	%	N concentration in leaf	Trade-offs between investment in support and photosynthetic structures
	Specific leaf area (SLA)	mm^2 mg^{-1}	Ratio between leaf area* and dry mass	
	Leaf area (LA)	mm^2	One-side leaf lamina area*	Gas exchange
	δ^{13}C	‰	Ratio of stable isotopes 13C:12C, reported in parts per thousand (per mil, ‰)	Water use efficiency

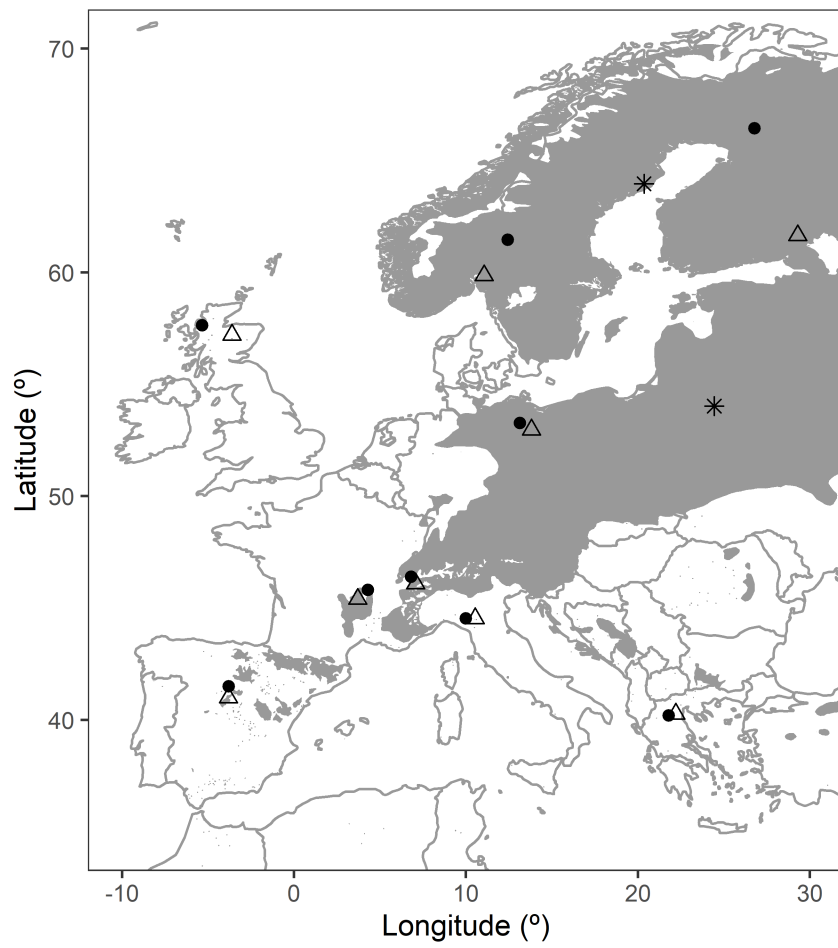* In this study projected leaf area is used.

Table 2. Summary of the optimal linear mixed models for individual trait values. Estimates (standard errors) and significance assessed with likelihood-ratio tests are shown.

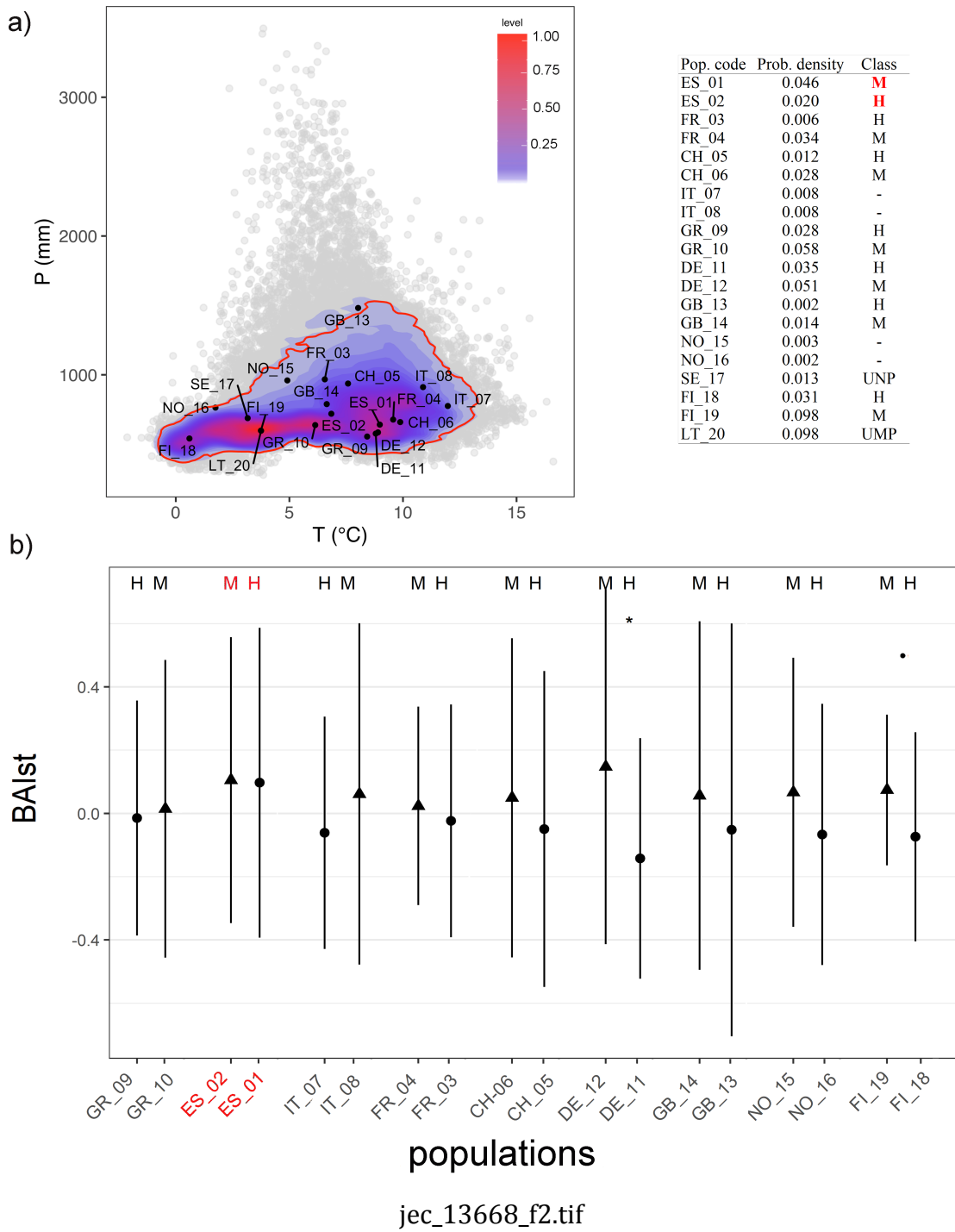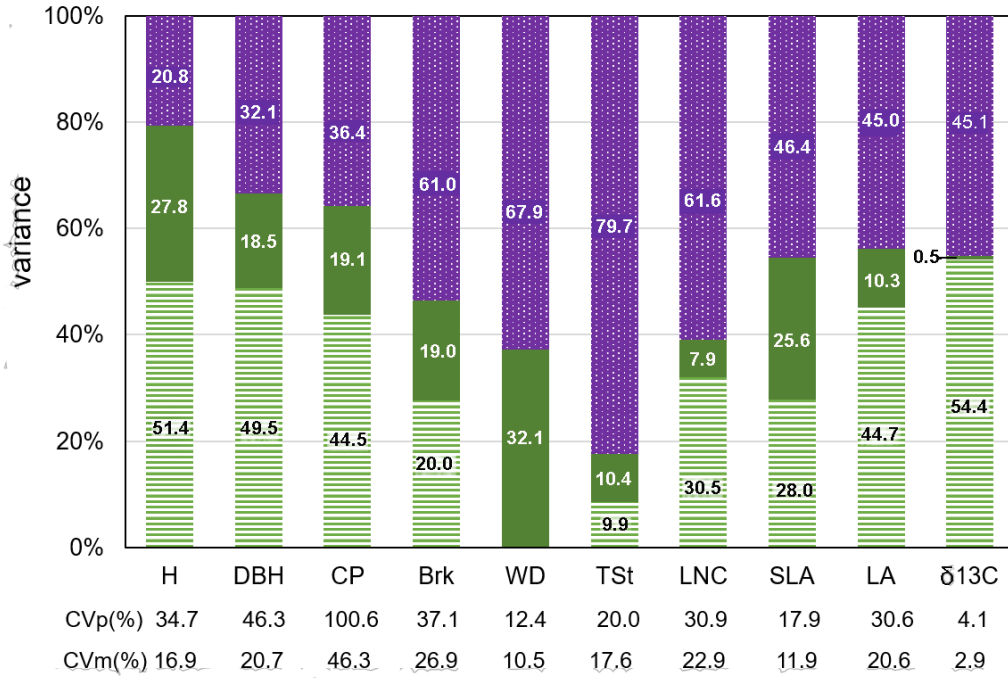|		Age	CI	slope	rad	T	T^2	P	Twet	n	R^2m	R^2c
Plant size traits	H	0.025*** (0.005)				4.417*** (1.061)	-0.338*** (0.008)	-0.017*** (0.004)	0.375*** (0.152)	495	0.53	0.80
	log(DBH)	0.003*** (0.0003)	-0.204*** (0.189)			0.264*** (0.060)	-0.018*** (0.080)			497	0.54	0.79
	log (CP)	0.005*** (0.0007)	-0.456*** (0.044)	-0.014** (0.004)		0.543*** (0.131)	-0.034** (0.010)			497	0.49	0.75
Stem traits	BxCx (Brk)	0.010*** (0.002)	-0.514*** (0.122)		0.378. (0.2100)	0.707** (0.257)	-0.044* (0.020)			497	0.30	0.53
	BxCx (WD)	0.001** (0.004)		-0.006** (0.002)	-0.088* (0.038)	0.029* (0.012)				365	0.14	0.31
	TSt	-0.004*** (0.001)								497	0.05	0.24
Leaf traits	BxCx (LNC)				0.117* (0.051)	0.212* (0.053)	-0.017*** (0.004)			277	0.28	0.43
	log (SLA)		0.032** (0.012)							497	0.01	0.54
	log (LA)					0.186*** (0.0413)	-0.010** (0.003)	-0.0004** (0.0002)		497	0.39	0.56
	$\delta^{13}C$		-0.262*		0.350.					277	0.06	0.51

| | (0.105) | (0.183) |

CI: competition index; slope (º); rad: potential total solar radiation (GJ/m^2); T: mean annual temperature (ºC); P: mean annual precipitation (mm); Twet: temperature of the wettest quarter of the year (ºC) ; H: tree height; DBH: diameter at breast height; CP: crown projection area; Brk: bark thickness; TSt: trunk straightness; WD: wood density; LNC: leaf N content; SLA: specific leaf area; LA: leaf area; δ13C (‰). ΔAIC represents the improvement of the model removing non-significant variables compared to the saturated model. Marginal (R^{2}m) and conditional (R^{2}c) r-squared are shown, showing the variance explained by the model considering only fixed and both fixed and random effects. Significance: •10%, *5%, **1%, ***01%.
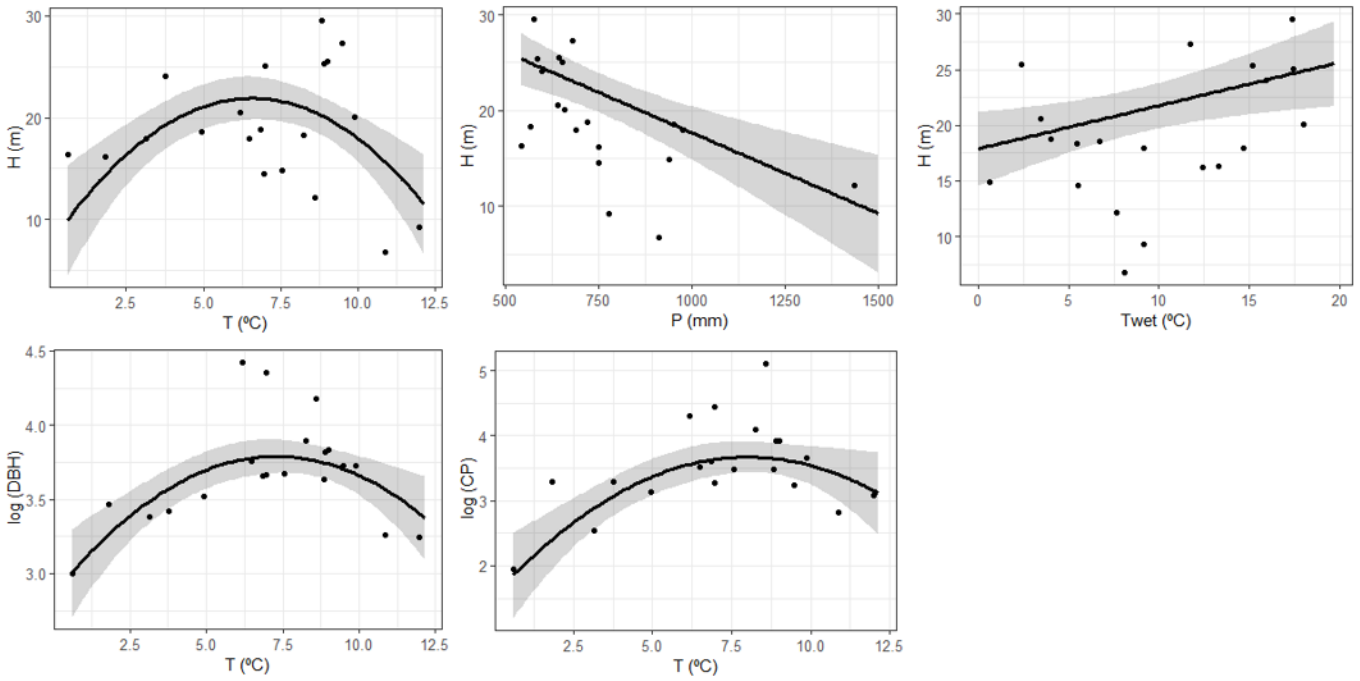
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a)

Pop. code	Prob. density	Class
ES_01	0.046	**M**
ES_02	0.020	**H**
FR_03	0.006	H
FR_04	0.034	M
CH_05	0.012	H
CH_06	0.028	M
IT_07	0.008	-
IT_08	0.008	-
GR_09	0.028	H
GR_10	0.058	M
DE_11	0.035	H
DE_12	0.051	M
GB_13	0.002	H
GB_14	0.014	M
NO_15	0.003	-
NO_16	0.002	-
SE_17	0.013	UNP
FI_18	0.031	H
FI_19	0.098	M
LT_20	0.098	UMP

b)

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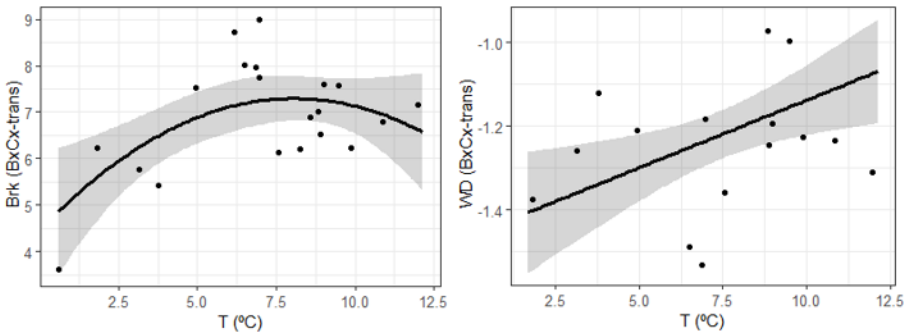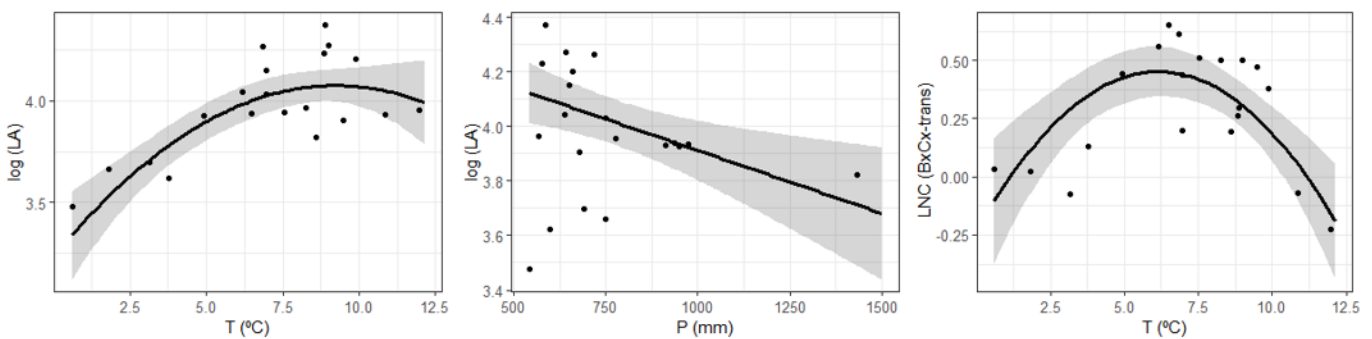|	H	DBH	CP	Brk	WD	TSt	LNC	SLA	LA	δ13C
CVp(%)	34.7	46.3	100.6	37.1	12.4	20.0	30.9	17.9	30.6	4.1
CVm(%)	16.9	20.7	46.3	26.9	10.5	17.6	22.9	11.9	20.6	2.9

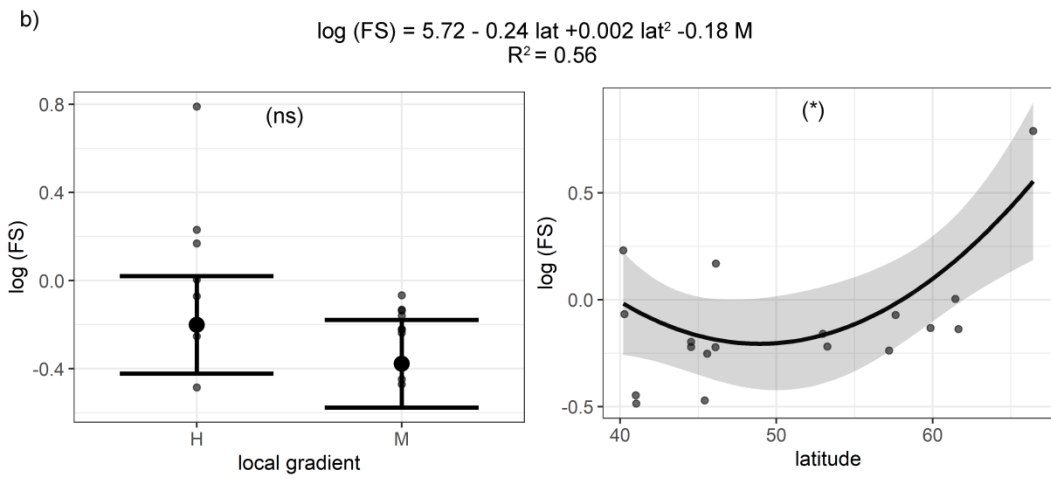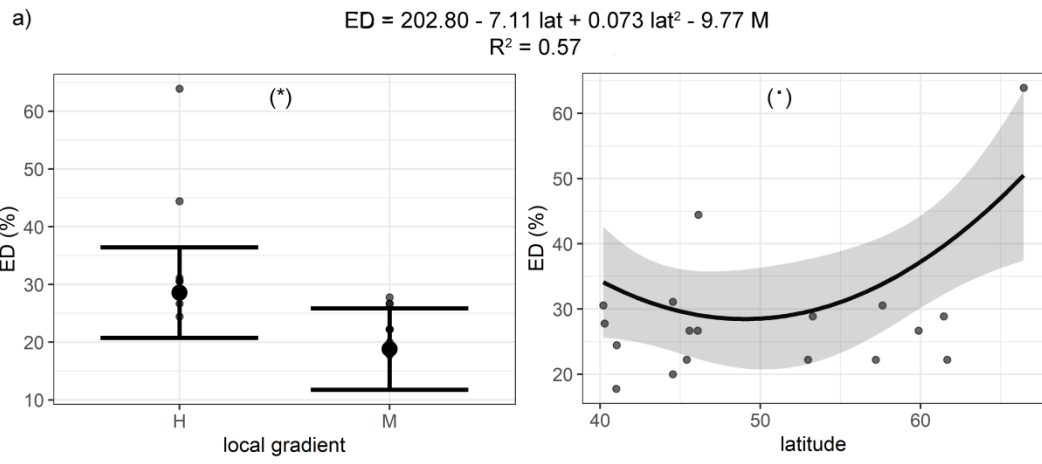jec_13668_f3.tif

a) Plant size traits

b) Stem traits

c) Leaf traits

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Greece
M — ED= 27.8%, FS= 0.94   H — ED= 30.6%, FS= 1.26

Spain
M — ED= 17.8%, FS= 0.64   H — ED= 24.4%, FS= 0.62

Italy
M — ED= 20%, FS= 0.80   H — ED= 31.1%, FS= 0.82

France
M — ED= 22.2%, FS= 0.62   H — ED= 26.7%, FS= 0.78

Switzerland
M — ED= 26.7%, FS= 0.80   H — ED= 42%, FS= 1.18

Germany
M — ED= 22.2%, FS= 0.85   H — ED= 28.9%, FS= 0.80

Lithuania
ED= 28.9%, FS= 1.00

United Kingdom
M — ED= 22.2%, FS=0.79   H — ED= 30.6%, FS= 0.93

Norway
M — ED= 26.7%, FS= 0.88   H — ED= 28.9%, FS= 1.00

Sweden
ED= 31.1%, FS= 0.82

Finland
M — ED= 22.2%, FS= 0.87   H — ED= 63.9%, FS= 2.20

jec_13668_f5.tif

a) ED = 202.80 - 7.11 lat + 0.073 lat^2 - 9.77 M
R^2 = 0.57

b) log (FS) = 5.72 - 0.24 lat +0.002 lat^2 -0.18 M
R^2 = 0.56

jec_13668_f6.tif