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1 **Metal and proton toxicity to lake zooplankton: a chemical speciation based modelling**
2 **approach**

3

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26

27 **1. Introduction**

28 Current environmental regulations rarely require assessment of mixtures of chemicals.
29 However, research on metal mixtures in the environment, and the development of efficient
30 and economic science-based risk assessment approaches for mixtures, are vital for ensuring
31 adequate environmental protection and for being prepared for future regulatory demands (Van
32 Genderen et al., 2012). Several methods have been proposed to relate organism responses to
33 the concentrations of mixtures of potentially toxic metals in the field. A pragmatic means of
34 combining single-metal Environmental Quality Standards (EQS) for the field assessment of
35 mixtures is through the use of Cumulative Criterion Units (Clements et al., 2008), calculated
36 as the sum of the ratios of the concentration of each metal to its individual EQS. However,
37 this only predicts an acceptable limit, not a concentration-response relationship, and as
38 applied to date does not take full account of chemical speciation and bioavailability. The
39 concept of multi-substance Potentially Affected Fractions has been developed, in which
40 results from experiments with single toxicants are used to predict the proportion of species
41 affected by a chemical mixture in the field, assuming either concentration or response
42 additivity (De Zwart and Posthuma, 2005 and De Zwart et al., 2006). The Biotic Ligand
43 Model has been applied in a meta-analysis study of metal mixtures, by assessing competition
44 for binding at the biotic ligand by toxic metals (Zn, Cu, Cd), calcium and other major cations
45 (Kamo and Nagai, 2008). A ‘saturation’ model, where toxicity is assumed to be a function of
46 the amount of metal bound to a specific binding site on the organism, has also been tested for
47 mixtures (e.g. Borgmann et al., 2004 and Norwood et al., 2007). The toxic unit approach,
48 including incorporation of BLM modelling, has been applied to zooplankton communities of
49 smelter-affected lakes in Sudbury, Canada (Khan et al., 2012). However, rather than using
50 field biology data to define toxic effects, the majority of these approaches generally
51 incorporate laboratory toxicity data and then use a derived model to make predictions of
52 toxicity in field settings.

53 The chemical speciation code WHAM VI (Windermere Humic Aqueous Model; Model
54 VI; Tipping, 1998;<http://www.ceh.ac.uk/products/software/wham/>) calculates the distribution
55 of chemical species within waters, accounting for competition for binding sites on dissolved
56 organic matter between protons and multiple cationic metals. Relationships between organism
57 binding of metals and predictions of metal binding to humic acid (HA) using WHAM have
58 been shown for bryophytes (Tipping et al., 2008) and stream macroinvertebrates (Stockdale
59 et al., 2010). The WHAM- FTOX model (Stockdale et al., 2010), is based on this ability of

60 WHAM to yield results that are a proxy for metal binding to organisms. It is assumed that, at
61 steady-state, the WHAM calculated bound metal and proton correlate with the “metabolically
62 available” cations of the organisms, as defined by Rainbow (2002), i.e., variability in humic
63 acid bound metals and protons will reflect the variability in the metabolically available metal
64 and protons. The ability of WHAM (specifically the binding of metals and protons to HA) to
65 predict the accumulation of metals by biota is supported by recent work showing correlation
66 between measurements and modelling for several different taxa (Tipping et al.,
67 2008, Stockdale et al., 2010 and Tipping and Lofts, 2013). WHAM-FTOX quantifies the
68 combined toxic effects of protons and metal cations towards aquatic organisms through the
69 toxicity function (F_{TOX}), a linear combination of the products of organism-bound cation and a
70 toxic potency coefficient for each cation. Stockdale et al. (2010) parameterised WHAM-
71 FTOX using the combined species richness of the macroinvertebrate groups Ephemeroptera,
72 Plecoptera and Trichoptera (EPT) as the ecological indicator. In subsequent work Stockdale
73 et al. (2014) used the EPT version of the model to compare temporal predictions of maximum
74 macroinvertebrate species richness with field observations and found good agreement
75 between observed and predicted recovery rates from acidification. The use of cation binding
76 to HA as a proxy for accumulation by organisms, and consequent toxicity, has also been
77 reported by Antunes et al. (2012) and Iwasaki et al. (2013).

78 The WHAM- F_{TOX} model is a promising tool relating the chemical speciation of natural waters
79 to field biological responses. In order to increase the utility of the model for applications
80 related to environmental monitoring and protection there is a need to establish relationships
81 for additional biological taxa. Here, we apply WHAM- F_{TOX} to over 1000 observations of
82 pelagic zooplankton species richness and associated water chemistry from surveys of 184
83 individual lakes from across Europe and North America, which have been studied with
84 respect to either acidification as a result of atmospheric acid deposition or the impacts of
85 metals from historical and continuing smelting operations in the Sudbury region of Canada
86 (Keller et al., 2007). Pelagic zooplankton are good target organisms for this analysis as they
87 have a constant exposure (including via food) to the medium (water) where chemistry is
88 measured. Moreover, zooplankton populations have an important role in lake functioning, by
89 affecting water clarity through grazing, and providing food for fish and predatory
90 invertebrates (Makarewicz and Likens, 1979). The aims were (1) to obtain model parameters
91 applicable to lake crustacean zooplankton species richness and (2) to assess model utility by

92 comparing predictions with observations of temporal change, for cases where monitoring has
93 been performed for a period of 10 years or more.

94 Although analysis of field data is attractive because it addresses toxicity directly within the
95 given environment, a drawback is that species richness depends upon environmental factors
96 other than cation toxicity. Such factors may be biotic (trophic status, competition, predation
97 and food web structure), physical (temperature, lake area, stratification, suspended sediment),
98 or chemical (presence of anionic toxicants such as arsenic, organic pollutants). Examples of
99 these factors in the literature are provided by Keller and Conlon (1994; lake depth and
100 area); Stemberger et al. (1996; temperature); Kratz et al. (1997; elevation); Jeppesen et al.
101 (2000; productivity); Valois et al. (2010; predation); Nierzwicki-Bauer et al. (2010;
102 predation). Both chemical toxicity and adverse unknown factors are assumed to reduce
103 species numbers below a maximum that may be expected to be observed if all factors were at
104 their optimum. As factors other than cation toxicity are not universally quantified across all
105 surveys, we employed quantile regression, a statistical technique that provides a more
106 complete view of possible causal relationships between variables in ecological processes than
107 the simple use of linear response variables (Cade and Noon, 2003). This is achieved by fitting
108 the data with a ‘roof’ that seeks to capture variations in zooplankton species richness that
109 are due solely to chemical (cation) factors. This approach has previously been applied to
110 assessing organism responses to metals in freshwaters, including (Stockdale et al.,
111 2010 and Schmidt et al., 2012) and disregarding (Pacheco et al., 2005, Crane et al.,
112 2007 and Linton et al., 2007) chemical speciation. The ‘roof’ fitted by quantile regression
113 is in essence a prediction of the species numbers in the absence of adverse unknown factors
114 but in the presence of varying degrees of chemical toxicity. This is under the assumption that
115 the theoretical maximum has no spatial variability.

116 **2. Methods**

117 **2.1. Data sources and data processing**

118 We required datasets that included both water chemistry and zooplankton species richness for
119 the same lakes. We collated data from several studies reported in peer reviewed literature
120 from both time series and spatial studies. Detailed location descriptions, including number of
121 lakes and the years over which sampling was undertaken, are included in Table 1. Data were
122 obtained for the Bohemian Forest region, Czech Republic (Vrba et al., 2003); ranges of

123 locations in Norway (the same sites and years as reported in Stockdale et al., 2014;
124 zooplankton data previously reported in Walseng et al., 2006 and Skjelkvåle, 2010) and
125 Sweden (Persson, 2008); the English Lake District region of the UK; Lake Plessa 107, formed
126 as a result of lignite mining in Germany (Lessmann et al., 1999); lakes in the Adirondacks
127 region of New York state, USA (Nierzwicki-Bauer et al., 2010); and two surveys of lakes in
128 the Sudbury region of Ontario, Canada, one spatial (Valois et al., 2010) and one temporal
129 (Khan et al., 2012). The Norwegian and Swedish sites are geographically spread across the
130 respective countries but are not necessarily representative of either country. The Sudbury data
131 include a number of sites that are known to have historical metal pollution, the majority of
132 other sites are, or were, affected by acidification.

133 Crustacean zooplankton data were from summer collections and represent the total species
134 richness (number of species; SR_{Czo}) of the pelagic zone. Data are either from a single sampling
135 period or are averaged over several sampling periods. When littoral species were found within
136 the water column during sampling they were included in the value for species richness for all
137 study regions except the Bohemian Forest lakes. The Supporting information provides details
138 of sampling methodology for each region and species lists for each regional group of lakes.

139 Chemistry data were either means of multiple samples taken during the year (Bohemian
140 Forest region, Czech Republic, Norway, Sweden and Adirondacks, USA), or from a single
141 sample (Lake Plessa 107, English Lake District and Sudbury, Canada datasets). Details of the
142 sampling and analysis methods are given in the site-specific references above, except for the
143 English Lake District (Maberly et al., 2011 and Tipping et al., 1988). The 10th and 90th
144 percentiles of chemical data are provided in the Supporting Information. Not all datasets
145 included all of the required input data for trace metals and in these cases we used assumed
146 'background' values in the modelling. These assumed data are italicised in the Supporting
147 information.

148 **2.2. Chemical speciation**

149 The WHAM model (<http://www.ceh.ac.uk/products/software/wham/>) calculates the
150 equilibrium distributions of dissolved chemical species based on the following input
151 parameters; temperature, pH, pCO_2 , and the concentrations of DOC, major ions and trace
152 metals. The model incorporates the Humic Ion-Binding Model VI (Tipping, 1998), which
153 uses a structured formulation of discrete, chemically plausible, binding sites for protons,
154 allowing the creation of regular arrays of bidentate and tridentate binding sites for metals.

155 Metal aquo ions (Al^{3+} , Fe^{3+} , etc.) and their first hydrolysis products (AlOH^{2+} , FeOH^{2+} , etc.)
 156 compete with each other, and with protons, for binding. The same intrinsic equilibrium
 157 constant is assumed to apply to the aquo ion and its first hydrolysis product (Tipping et al.,
 158 2008). In calculating aqueous chemical speciation, the concentrations of Na, Mg, K, Ca, Cl,
 159 NO_3 and SO_4 were assumed to represent dissolved components, as were concentrations of
 160 filterable trace metals (Ni, Cu, Zn, Cd and Pb). We used reactive Al for the calculations when
 161 it was available, but where filtered Al data were provided, the activity of Al was calculated
 162 from the measured total filtered concentration and from the generalised equation derived
 163 by Tipping (2005), with the lower of the two values being adopted. This avoided over-
 164 estimation of Al activity in filtrates containing colloidal forms of the element. The activity of
 165 Fe(III) was calculated from the empirical equation of Lofts et al. (2008). Concentrations of
 166 dissolved organic matter (DOM) were based on measured [DOC], assuming DOM to be 50%
 167 carbon and that 65% of it is active with respect to cation binding and can be represented by
 168 fulvic acid (Tipping et al., 2008). For example, for a DOC concentration of
 169 5 mg L^{-1} corresponds to a fulvic acid concentration of 6.5 mg L^{-1} for modelling. For each
 170 speciation calculation, cation binding to humic acids (the proxy for accumulation by
 171 organisms) is computed, by assuming the presence of $10^{-6} \text{ gHA L}^{-1}$, which is sufficiently low
 172 not to affect the solution distribution of species (i.e. HA is not affecting the overall speciation,
 173 FA is present as the dominant DOM component).

174 **2.3. The WHAM- F_{TOX} model**

175 The combined toxicity of the cations is taken to be the sum of the products of (a) bound
 176 proton or metal cation (v_i , mmol gHA^{-1}) and (b) a cation-specific toxicity coefficient (α_i) that
 177 characterises the relative potency of the individual cation. Thus we obtain a linear toxicity
 178 function;

$$179 \quad F_{\text{TOX}} = \sum \alpha_i \cdot v_i \quad (1)$$

180 Note that the values of v_i are taken to be proportional, but not equal, to accumulation by the
 181 organism, since only relative binding is required. We describe the dataset with a lower
 182 threshold, $F_{\text{TOX-LT}}$, below which no toxic effects occur and an upper threshold, $F_{\text{TOX-UT}}$, above
 183 which all crustacean zooplankton species are absent. Thus we have the following conditions,
 184 in terms of species richness (SR_{Czo} ; Eqs. (2), (3) and (4));

$$185 \quad F_{\text{TOX}} \leq F_{\text{TOX-LT}} \quad \text{SR}_{\text{Czo}} = \text{SR}_{\text{Czo-max}} \quad (2)$$

$$186 \quad F_{\text{TOX-LT}} < F_{\text{TOX}} < F_{\text{TOX-UT}} \quad (3)$$

$$SR_{\text{Czoo}} = SR_{\text{Czoo-max}} - \left(SR_{\text{Czoo-max}} \times \frac{F_{\text{TOX}} - F_{\text{TOX-LT}}}{F_{\text{TOX-UT}} - F_{\text{TOX-LT}}} \right)$$

$$187 \quad F_{\text{TOX}} > F_{\text{TOX-UT}} \quad SR_{\text{Czoo}} = 0 \quad (4)$$

188 **2.4. Data analysis**

189 When examining field biological data, a direct correlation between predicted and actual
 190 species richness is expected only when modelled chemical behaviour is the sole factor that
 191 reduces species richness. To account for other unknown factors that also reduce richness, we
 192 used quantile regression. In this approach, quantiles of the response variable are employed,
 193 rather than the conditional mean (Cade and Noon, 2003). Quantiles were estimated with an
 194 optimisation function that minimises the sum of weighted absolute deviations, using the
 195 Solver function in Microsoft Excel[®].

196 The use of a structured dataset, comprising individual subsets of spatially and/or temporally
 197 linked datasets, introduces the possibility of bias into the estimation of regression parameter
 198 variance. In conventional (least squares) regression, this would be alleviated by the use of a
 199 linear mixed modelling approach. In this work, however, the use of quantile regression
 200 minimises the risk of introducing bias in parameter variance due to the relaxation of the
 201 assumptions regarding the distribution of residuals about the regression line, therefore we
 202 have elected not to consider the data subset structure in parameterising the model.

203 We confirmed the significance of the contribution to F_{TOX} of proton and individual metals by
 204 performing t -tests. Uncertainties in the model parameters $SR_{\text{Czoo-MAX}}$, $F_{\text{TOX-UT}}$, $F_{\text{TOX-}}$
 205 $_{\text{LT}}$, α_{Al} , α_{Ni} , α_{Cu} and α_{Zn} were estimated by bootstrapping, sampling the dataset with
 206 replacement 2000 times. The median, 15.9%ile and 84.1%ile of the distribution of each
 207 parameter were calculated; the range enclosed by these percentiles is a nonparametric
 208 equivalent of taking ± 1 standard deviation of a normal distribution.

209 Regional Kendall testing (Helsel and Frans, 2006) was used to determine whether there were
 210 unidirectional trends over time and if a consistent trend occurred across an entire area at
 211 multiple locations. We analysed observed and modelled time series data of 10 years or more
 212 (Bohemian Forest, Norway, Adirondacks and Sudbury). Previous work compared temporal
 213 trends in predicted and observed macroinvertebrate species richness (Stockdale et al., 2014).

214 This work used WHAM-FTOX parameters derived from independent data (Stockdale et al.,
215 2010) and data distribution across the prediction 'roof' was consistent with the independent
216 fitting. Here we are performing temporal analysis for the same data as used for fitting.

217 **3. Results and discussion**

218 **3.1. WHAM- F_{TOX} fitting to crustacean zooplankton data**

219 Values of the amount of proton or metal bound to humic acid extracted from the WHAM
220 model (Table 2) give an overall indication of the various cation chemical stressors within each
221 of the survey areas. Values for protons are always above 1 mmol g⁻¹ but reach over 4 for the
222 extremely acid site Lake Plessa 107. Aluminium has a large range, from low values in
223 circumneutral environments (e.g., English Lake District, <1 $\mu\text{g L}^{-1}$), to high in many of the
224 acidification affected regions (e.g., Lake Plessa 107, >1 mg L⁻¹). Trace metals are generally
225 present at bound concentrations an order of magnitude or more lower than H⁺ or Al. The
226 ranges for all cations reveal the large inter-regional and geographical variability in the
227 chemical measure.

228 All datasets were combined for fitting, in order to obtain global parameters (Table 3). Species
229 richness was related to water chemistry by fitting the model to the 0.9th quantile, selected
230 because higher quantiles may result in greater errors in fitted parameters (Cade and Noon,
231 2003). Toxicity coefficients (α_i), the F_{TOX} thresholds ($F_{\text{TOX-LT}}$ and $F_{\text{TOX-UT}}$) and the maximum
232 value for species richness ($\text{SR}_{\text{CZOO-MAX}}$) were optimised. Different starting values were
233 applied to ensure that the correct optimisation minimum was reached. Provided that (a) the
234 values were not initially set to zero, (b) $F_{\text{TOX-UT}}$ was forced to exceed $F_{\text{TOX-LT}}$ and (c) the
235 initial estimate for $\text{SR}_{\text{CZOO-MAX}}$ was no higher than the maximum observed value of 16, the
236 model always converged to give consistent values.

237 Values of the fitted toxicity coefficients (α_i) were statistically significant for Al, Ni, Cu and
238 Zn, but not for either Cd or Pb. Toxic potencies increase in the order H⁺ < Al < Cu < Zn < Ni,
239 which is broadly similar to WHAM- F_{TOX} derived values for stream macroinvertebrates
240 (Stockdale et al., 2010). However, α_{Cu} and α_{Zn} for the zooplankton (Table 3) are greater than
241 the values of 3.50 and 2.69 respectively obtained for the macroinvertebrates, implying
242 different sensitivities to trace metal toxicity between the two groups of organisms. However,

243 we cannot rule out the possibility that physiological differences between taxa mean that using
244 HA as a proxy for organism binding may not be equally appropriate for zooplankton and
245 macroinvertebrates. The toxic contribution of each cation (metal or proton) depends on the
246 cation's toxicity coefficient, combined with its bound concentration. Thus, a metal that
247 exhibits weak binding but has a comparatively high toxicity coefficient can exert the same
248 toxic effect as a strongly binding metal with a comparatively low toxicity coefficient. Thus, in
249 the present study, the high value of α_{Ni} arises because the speciation model predicts weak
250 binding of Ni by HA. The values of $F_{\text{TOX-LT}}$ and $F_{\text{TOX-UT}}$ for zooplankton (Table 3) are similar
251 to those estimated for stream macroinvertebrates (2.33 ± 0.48 and 5.20 ± 0.65 respectively).

252 The “broken stick” line in Fig. 1 shows the modelled effects of water chemistry on species
253 richness in terms of F_{TOX} . Points falling below the line indicate that additional un-modelled
254 (i.e. non-chemical) factors have reduced species numbers. Overall, these effects are less
255 marked when compared to previous findings for stream macroinvertebrates, suggesting that
256 zooplankton diversity in lakes is subject to fewer or weaker non-chemical disturbances than
257 macroinvertebrates in streams.

258 Plots of F_{TOX} against pH, with data points separated according to the contribution of trace
259 metals (Ni, Cu and Zn) to F_{TOX} , show that toxicity due to trace metals is encountered mainly
260 in the pH region 6 to 7.5 (Fig. 2), and these data largely reflect the temporally studied lakes at
261 Sudbury. The most extreme case is Lake Plessa 107 which has high concentrations
262 ($>10 \mu\text{mol L}^{-1}$) of both Ni and Zn. However, the very low pH of 2.3, and thus high
263 competition for humic binding sites from protons, means that neither of these metals are
264 calculated to be significantly bound by humic acid (the proxy for non-specific sites on the
265 organism) and so neither contributes to F_{TOX} . In the absence of significant contributions from
266 trace metals, F_{TOX} exceeds the lower threshold for all sites at acid pH, although the pH above
267 which exceedance occurs varies with water chemistry (see below). The same effect was found
268 for macroinvertebrates (Stockdale et al., 2010).

269 **3.2. Results for individual datasets**

270 **3.2.1. Bohemian Forest lakes (Czech Republic)**

271 The lakes included in this study were strongly acidified by acid deposition during the 1960s
272 and 1970s (Vrba et al., 2003, Veselý et al., 1998 and Oulehle et al., 2013). The pH increased
273 by between 0.5 and 0.9 units since 1984. Predicted and observed recovery in species richness
274 was ~ 0.5 species per decade over the study period (Table 4). The constituent bound metals
275 that make up the F_{TOX} value (mainly Al for this region) can partly explain the slow recovery
276 rates for this region. Comparing data from 1984 to 2007-9 (three year average), we predicted
277 only a small decrease in F_{TOX} of between 0.13 and 0.19 for all lakes except Prášilské
278 (decrease of 0.4). The contribution of v_H to F_{TOX} has decreased, but this has partly been offset
279 by an increase in v_{Al} .

280 Regional Kendall tests (Table 4) show that recovery in crustacean zooplankton species
281 richness was both predicted from chemistry and observed in the Bohemian Forest lakes. There
282 is fair agreement between recovery trends predicted from maximum modelled species
283 richness (0.52 species per decade) and observations (0.43 species per decade). For these sites
284 there is a comparatively large gap between the modelled maximum species richness and the
285 observed values (Fig. 3), which may be explained partly by the absence of littoral species
286 from this dataset (numbering up to 6; see the Supporting information for details).

287 **3.2.2. Norway**

288 Regional Kendall testing (Table 4) shows actual recovery (0.7 species per decade) slightly
289 higher than the predicted recovery of 0.5 species per decade. The discrepancy may be related
290 to regional factors such as ongoing natural slow colonisation of isolated lakes, which would
291 still occur if chemical conditions remained constant. Data from this survey show how such
292 biogeographic differences between lakes (considered as un-modelled factors in our analysis)
293 can additionally affect the species richness. Observations and model calculations of maximum
294 species richness for the Norwegian lakes are shown in Fig. 4. Highland lakes all exhibited
295 observed species richness significantly less than the maximum values predicted by the
296 modelling. In contrast boreal and lowland sites exhibited some sites with close agreement. In
297 addition to slow colonisation, the impact of dispersal barriers such as mountain ridges
298 separating adjacent lowland areas may affect the species richness across some lakes within

299 this study. We compared the mean difference in F_{TOX} between observations and modelling
300 and lake altitude (Fig. 5). These results give further evidence that those lakes in the highland
301 climate region show larger differences between modelled maximum values and observations.
302 Additional analysis of this type demonstrates the potential for future statistical models to
303 include additional drivers.

304 **3.2.3. Sweden**

305 The Swedish lakes differ from other acid lakes in that they had high dissolved organic matter
306 concentrations (mean of 10 mg DOC L⁻¹), which altered the calculated binding of Al to humic
307 acid (i.e. our proxy of organism binding), through chemical speciation. Differences from other
308 lakes can be seen in Fig. 2 where these data have F_{TOX} values largely below the lower
309 threshold, but at a much lower pH compared with other data, particularly over the pH range
310 6.0-6.6. This shows the predicted protective effects that can be exhibited by DOM, as higher
311 binding to fulvic acid (directly related to measured DOM in our modelling) results in lower
312 binding to the humic acid (fixed concentration as the proxy for organism binding).

313 The data for Swedish lakes fall appreciably below the F_{TOX} prediction line, and to a greater
314 degree than some other sets. This may be due to lower richness values caused by
315 biogeographic differences, similar to the Norwegian lakes, as discussed above. Lower
316 taxonomical resolution (many taxa identified only to genus level) may also be an
317 explanation.

318 **3.2.4. English Lake District (Cumbria, United Kingdom)**

319 No toxic effects are predicted for these lakes and therefore it appears that non-chemical
320 environmental influences constrain species richness below the maximum predicted by the
321 model. Previous work has shown that variation in zooplankton species richness among these
322 lakes is at least partly related to variations in productivity, as indicated by proxies such as
323 concentrations of chlorophyll *a* and soluble reactive phosphorus (Thackeray, 2007). Enhanced
324 lake productivity may reduce species richness due to correlated changes in resource quality,
325 predation pressure and physico-chemical factors, e.g. oxygen concentrations (Dodson et al.,
326 2000). The English Lake District case study therefore highlights a need for future work to
327 evaluate the influence of metal toxicity on community structure, within the context of other
328 measured environmental drivers.

329 3.2.5. Sudbury (Ontario, Canada)

330 The model predicts recovery of 1.6 species per decade, in good agreement with the observed
331 value of 1.4 (Table 4). Recovery trends can be seen clearly for the individual lakes (Fig. 6).
332 Recovery in species richness was more marked than in other regions owing to the combined
333 effects of increased pH and DOC concentrations and decreasing concentrations of Ni, Cu, and
334 Zn. Changes reflect both reductions in inputs and manipulative liming. Middle, Hannah and
335 Lohi lakes were all limed in the 70s. Lohi Lake was limed in 1973, 1974 and 1975, but it re-
336 acidified, before slowly naturally recovering, Hannah Lake was limed in the spring of 1975,
337 after two years of monitoring. The pH of Middle and Hannah remained high because their
338 watersheds were later also limed, while the watershed of Lohi Lake was not. Furthermore,
339 Hannah flows into Middle, while the historically acidic Clearwater (not limed) feeds Lohi
340 Lake.

341 Due to chemical recovery from low pH values (<5) to higher pH (>6) over the monitoring
342 period, results from many lakes in the Sudbury region follow the trend of Fig. 2, where
343 toxicity as a result of bound metal is more evident at pH values greater than ~6. The complex
344 interplay among chemical variables needs to be considered when considering trends in F_{TOX} .
345 For several sites in this dataset (Clearwater, Hannah and Lohi) decreased Ni and Cu
346 concentrations (also Zn in Hannah) do not themselves lead to a decrease in F_{TOX} driven by
347 decreased in bound metals. The v_{Ni} , v_{Cu} , and v_{Zn} values actually increased over the period, as a
348 result of decreased competition from H and Al. However, this increase was not enough to
349 offset the decreasing trend in v_H and v_{Al} . Middle Lake was monitored in the spring and
350 summer of 1973 (pH of ~4.6) before being subjected to liming in 1973, the pH was ~6.5
351 during the latter part of the 1970s, rising to ~7.1 in the early 2000s due to watershed liming.
352 Following initial lime treatment, further recovery in Middle Lake was largely due to lower
353 concentrations of total metals driving lower amounts for modelled bound Ni, Cu and Zn (Cu
354 being the most important). Nelson and Joe Lakes approximately follow the trends of Middle
355 Lake although they had higher initial species richness and lower total metal levels. Whitepine
356 Lake differs from many others in Sudbury as F_{TOX} never exhibits large contributions from
357 bound metals and was not subject to the same degree of chemical changes as other locations.
358 The recovery in F_{TOX} was due to decreased metal and increased DOC concentrations and a
359 ~0.8 pH increase over the 27 years studied. Laundrie and Wavy Lakes remain acidified and
360 have only limited recovery.

361 **3.2.6. Adirondacks lakes region (New York State, United States)**

362 Many of the sites have observed species richness values close to those predicted by the model
363 (Fig. 7). For the region an actual recovery of 0.7 species per decade was observed, whereas
364 the model predicts no change. Thus our modelling suggests that on a regional basis, recovery
365 in Adirondack lakes was not related to changes in chemistry. This and the Sudbury dataset
366 tended to exhibit more sites that had actual species richness values that were closer to the
367 maximum values predicted by the model. Fewer or weaker un-modelled factors were
368 therefore indicated. For example species richness may have been less affected by low
369 dispersal (as indicated in Fig. 5).

370 **3.3. Dietary metal uptake**

371 WHAM- F_{TOX} makes no explicit assumptions about the role of dietary uptake as a vector of
372 toxic metal exposure. Diet may be a significant uptake route, as emphasised by Luoma and
373 Rainbow (2008). Previous work has shown good correlations between metal accumulation by
374 organisms at a variety of trophic levels and WHAM-predicted humic acid metal binding
375 (Tipping et al., 2008, Stockdale et al., 2010 and Tipping and Lofts, 2013). Therefore, it is
376 reasonable to assume that dietary metal uptake by zooplankton will exhibit a similar
377 dependence on water chemistry as direct uptake. We thus assume that trends in both dietary
378 and direct uptake correlate with trends in binding to the HA proxy used for the organism. The
379 validity of this assumption and the effect of the dietary metal uptake is worthy of further
380 study. There may be intervening factors such as relative metal accumulation potential of
381 different phytoplankton species, selectivity of feeding zooplankton to different phytoplankton,
382 digestibility and uptake potential of different metal binding components in the ingested
383 phytoplankton, which may cause deviation from a direct relationship between dissolved
384 speciation and ultimate ecotoxicological effects on local zooplankton.

385 **4. Conclusions**

386 1. The WHAM- F_{TOX} model relates the complex chemical interplay of pH, mixtures of
387 potentially toxic cations, DOC and major ions, with field ecological responses. By
388 using quantile regression and fitting the model on field-derived crustacean
389 zooplankton data responses can be predicted independently of other factors that may
390 influence species richness. The model gives regional predictions for recovery that are

391 consistent with observations for 3 of the 4 datasets. We have demonstrated that
392 recovery in crustacean zooplankton species richness is both predicted from chemistry
393 and observed across the geographical regions covered in the Bohemian Forest,
394 Norwegian and Sudbury surveys (0.4-1.6 species per decade).

395 2. The failure of the model to predict the observed recovery for the Adirondacks region
396 implies that non-chemical factors are responsible for the observed changes.

397 3. Future work could focus on examining the causes of the differences between
398 observations and predicted maximum species richness. The analysis of chemical toxic
399 effects provided by WHAM- F_{TOX} opens the possibility of identifying presently un-
400 modelled factors, as demonstrated by the relationship with altitude for lakes in the
401 Norwegian survey.

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535

536 **Tables**

537 Table 1. Mean, 10%ile and 90%ile values of v_i (the proxy for organism binding, the amount of metal¹ or protons bound per gram of humic acid) obtained from the WHAM
 538 modelling for each data set (units of mmol g⁻¹ for H and Al, and $\mu\text{mol g}^{-1}$ for Ni, Cu and Zn).

Survey	<i>n</i>	H			Al			Ni			Cu			Zn		
		10%ile	Mean	90%ile	10%ile	Mean	90%ile	10%ile	Mean	90%ile	10%ile	Mean	90%ile	10%ile	Mean	90%ile
Norway	231	1.35	1.68	2.17	0.63	4.46	10.6	0.06	0.38	0.95	0.63	4.46	10.64	0.41	4.63	13.18
Sudbury (temporal)	172	1.16	1.56	2.01	3.96	56.0	133	0.43	14.00	36.88	3.96	56.00	133.07	0.54	7.94	16.21
Sudbury (spatial)	116	1.21	1.42	1.68	2.34	16.6	47.4	0.39	7.70	23.24	2.34	16.61	47.45	0.83	7.87	20.28
L. Plessa 107	1	-	4.09	-	-	0.00	-	-	0.00	-	-	0.00	-	-	0.02	-
English Lake District	20	1.12	1.18	1.22	6.02	8.53	10.0	0.17	0.33	0.49	6.02	8.53	10.04	12.24	24.77	46.58
Sweden	80	1.15	1.37	1.65	0.48	3.72	10.6	0.01	0.05	0.11	0.48	3.72	10.60	0.32	1.54	2.44
Adirondacks	431	1.32	1.73	2.33	7.14	22.7	41.5	3.66	15.77	30.32	7.14	22.66	41.53	1.02	11.26	26.62
Bohemian Forest	28	1.12	1.54	1.99	0.19	1.06	2.27	0.00	0.00	0.00	0.19	1.06	2.27	0.24	0.71	1.55
All	1078	1.21	1.62	2.14	1.20	21.1	44.9	0.05	9.45	29.48	1.20	21.15	44.89	0.54	8.19	22.16

539 ¹Cd data not included as all values are $\leq 0.001 \mu\text{mol g}^{-1}$.

Table 2. WHAM- F_{TOX} zooplankton parameters from the best fit using the 90th quantile, with results of the significance testing (p values) for the α_i values and errors (one standard deviation) estimated by bootstrapping.

Parameter	Values	p
$\text{SR}_{\text{Czoo-MAX}}$	12.91 (\pm 0.50)	-
$F_{\text{TOX-LT}}$	2.43 (\pm 0.24)	-
$F_{\text{TOX-UT}}$	4.64 (\pm 0.10)	-
α_{H}	1 ^a	< 0.0001
α_{Al}	1.74 (\pm 0.15)	< 0.0001
α_{Ni}	14.62 (\pm 2.43)	< 0.0001
α_{Cu}	6.31 (\pm 0.69)	< 0.0001
α_{Zn}	12.08 (\pm 2.47)	< 0.0001
α_{Cd}	0.00	n/s
α_{Pb}	34.82	n/s

^aThe α_{H} value is fixed. n/s = not significant

Table 3. Temporal trends in actual and modelled species richness ($SR_{C_{zoo}}$) assessed using the Regional Kendall test. Values in parentheses show the significance (p), the lower the value the smaller the likelihood that there is no trend. Data from L. Plessa 107, English Lake District and Sweden are not included in the temporal analyses due to limited temporal data (1, ≤ 6 and 5 years, respectively).

Region	Number of sites	Data time span (years)	Change in actual $SR_{C_{zoo}}$ per decade	Change in predicted $SR_{C_{zoo}}$ per decade
Sudbury	9	34	1.40 (<0.0001)	1.57 (<0.0001)
Norway	19	14	0.68 (0.0001)	0.50 (<0.0001)
Adirondacks	35	17	0.74 (<0.0001)	0.03 (0.68)
Bohemian Forest	5	25-26	0.43 (0.0001)	0.52 (<0.0001)

Figure Captions

Figure 1. Relationship between zooplankton species richness and F_{TOX} using quantile regression. The solid line represents the best fit for the 90th quantile. The top left panel shows all data used for the WHAM- F_{TOX} fitting, other panels are included to show how each dataset relates to the overall fit.

Figure 2. Variations in F_{TOX} with pH. The lower horizontal dashed lines represents $F_{\text{TOX-LT}}$ and the upper dashed line $F_{\text{TOX-UT}}$. Shown are the 1078 data separated based upon the fraction of F_{TOX} that is comprised of contributions from trace metals (Cu, Ni and Zn). As there will always be some contribution to F_{TOX} from protons we chose to show data where there is at least a 20% contribution from bound trace metals. The Sweden lakes data are show separately as filled triangles.

Figure 3. Temporal trends in field observed species richness (filled circles) and modelled maximum species richness (open circles) based on the WHAM- F_{TOX} results for Bohemian Forest lakes. The horizontal dashed grey line denotes the global maximum (90th quantile) species richness that is predicted by the model. Some years include only observed data as chemistry data was not available.

Figure 4. Temporal trends in field observed species richness (filled circles) and modelled maximum species richness (open circles) based on the WHAM- F_{TOX} results for Norway lakes. The horizontal dashed grey line denotes the global maximum (90th quantile) species richness that is predicted by the model. The lakes designated as highland climate region are designated with (H), lowland sites with (L) and boreal with (B).

Figure 5. Comparison of the mean differences of observed and calculated species richness ($\text{SR}_{\text{OBS}} - \text{SR}_{\text{MOD}}$), and lake altitude for sites in the Norway study. See Figure 4 for details of which individual lakes are classified as which climate region.

Figure 6. Temporal trends in field observed species richness (filled circles) and modelled maximum species richness (open circles) based on the WHAM- F_{TOX} results for Sudbury lakes . The horizontal dashed grey line denotes the global maximum (90th quantile) species richness that is predicted by the model.

Figure 7. Temporal trends in field observed species richness (filled circles) and modelled maximum species richness (open circles) based on the WHAM- F_{TOX} model results for Adirondacks lakes. The horizontal dashed grey line denotes the global maximum (90th quantile) species richness that is predicted by the model.

Figure 1

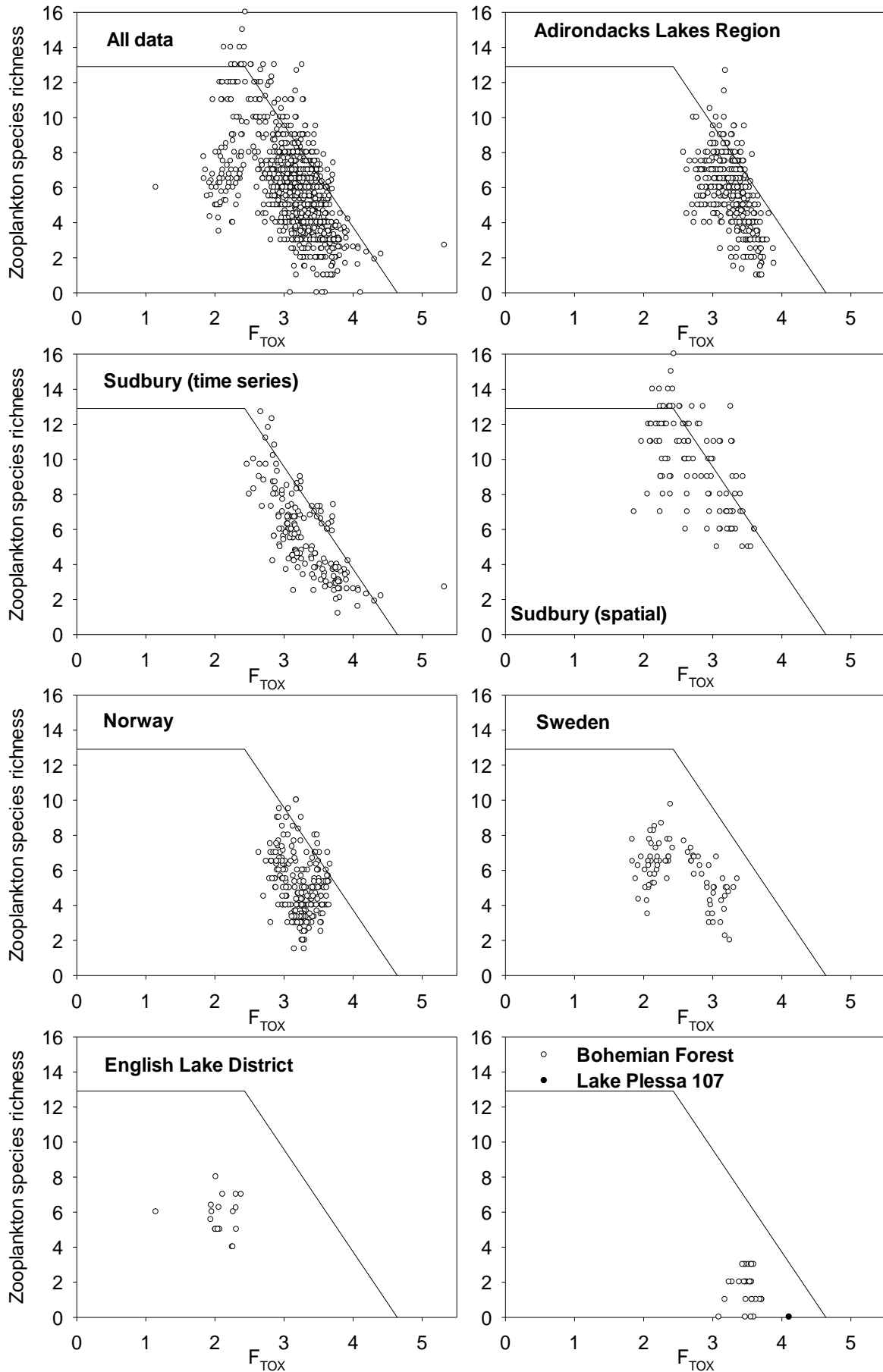


Figure 2

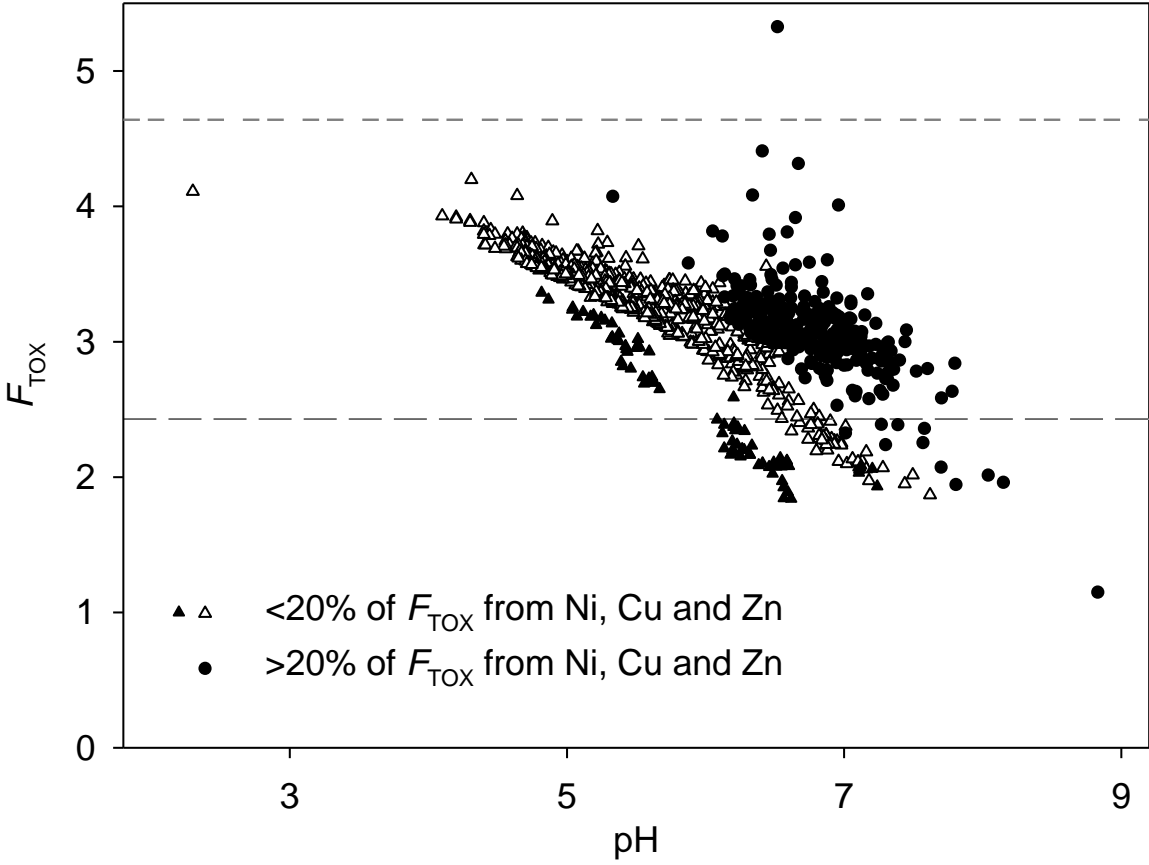


Figure 3

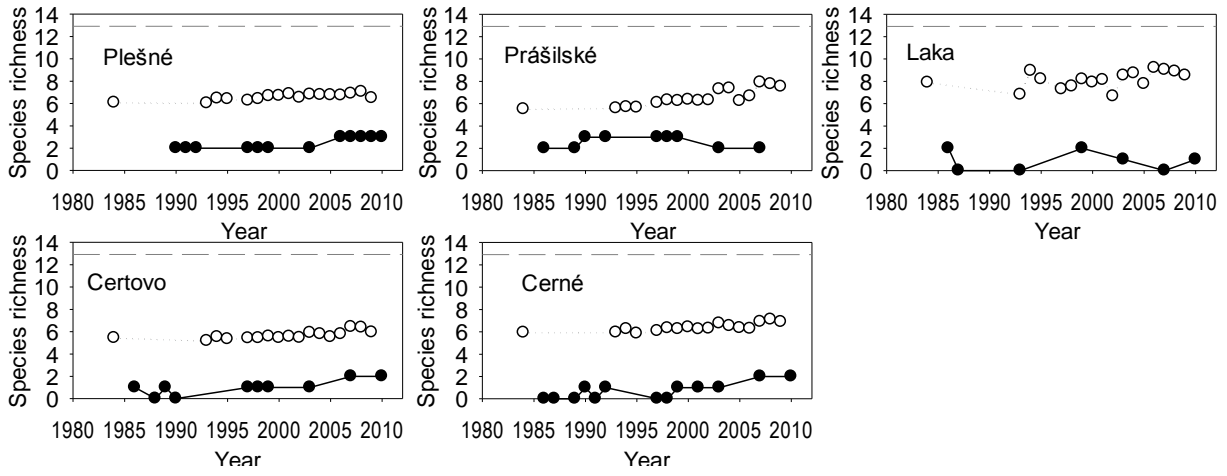


Figure 4

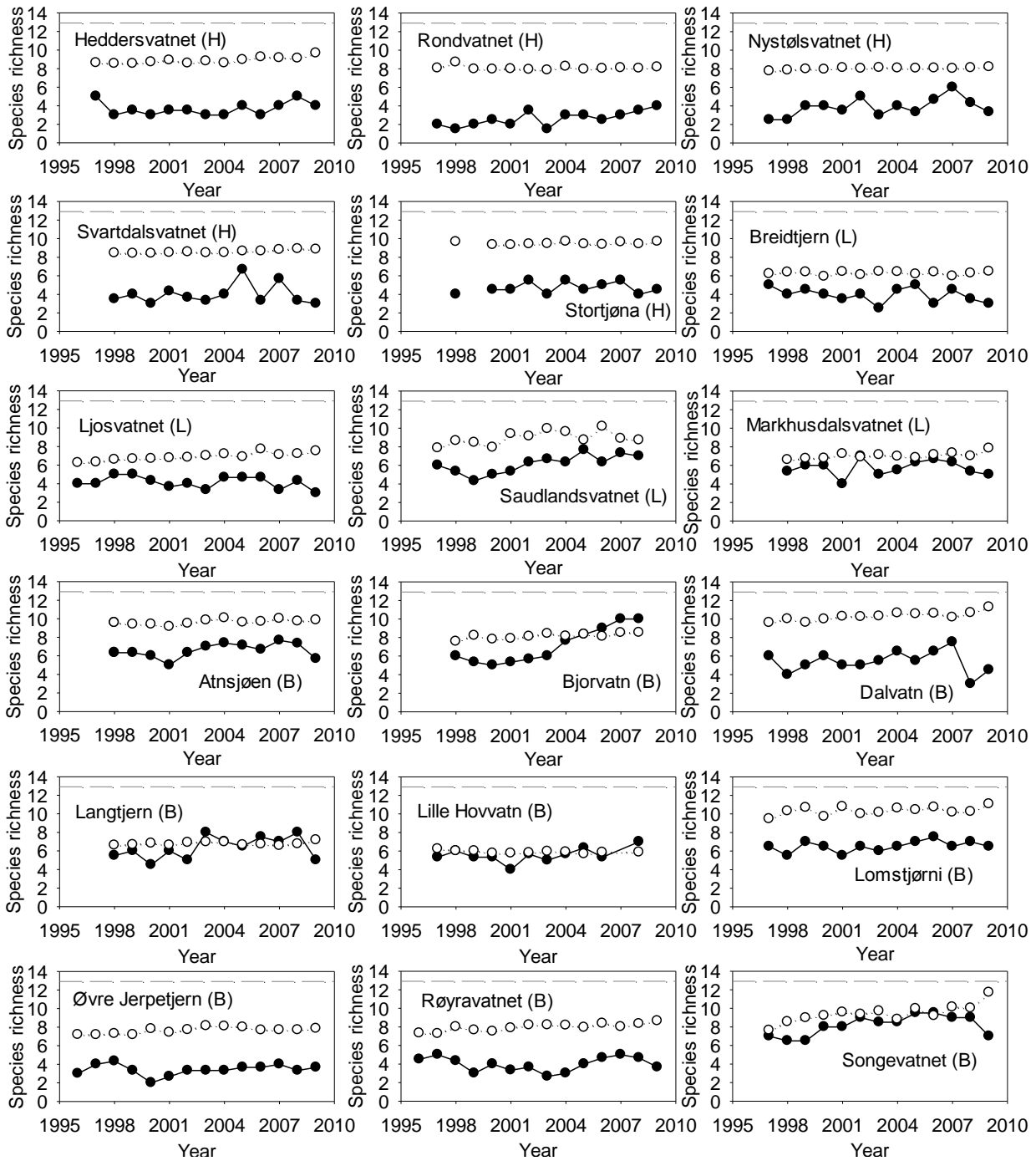


Figure 5

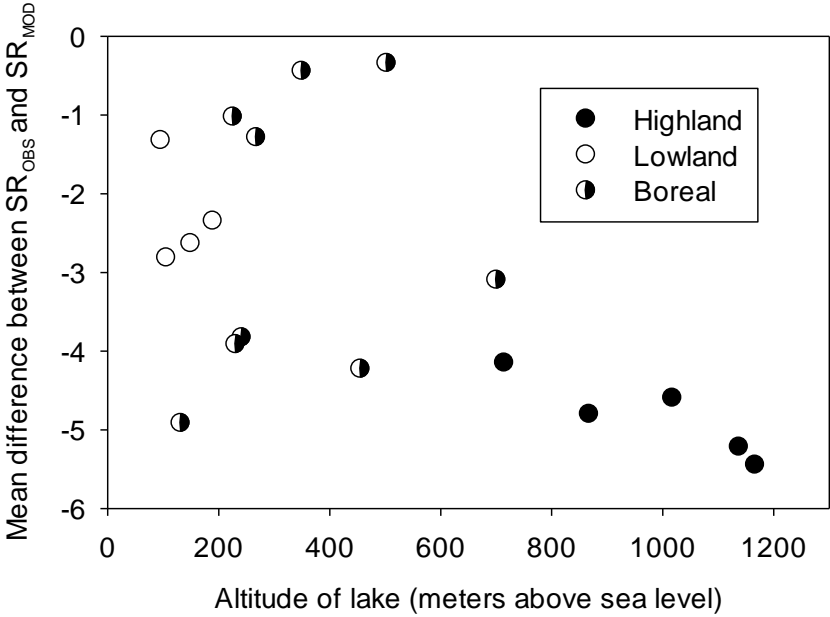


Figure 6

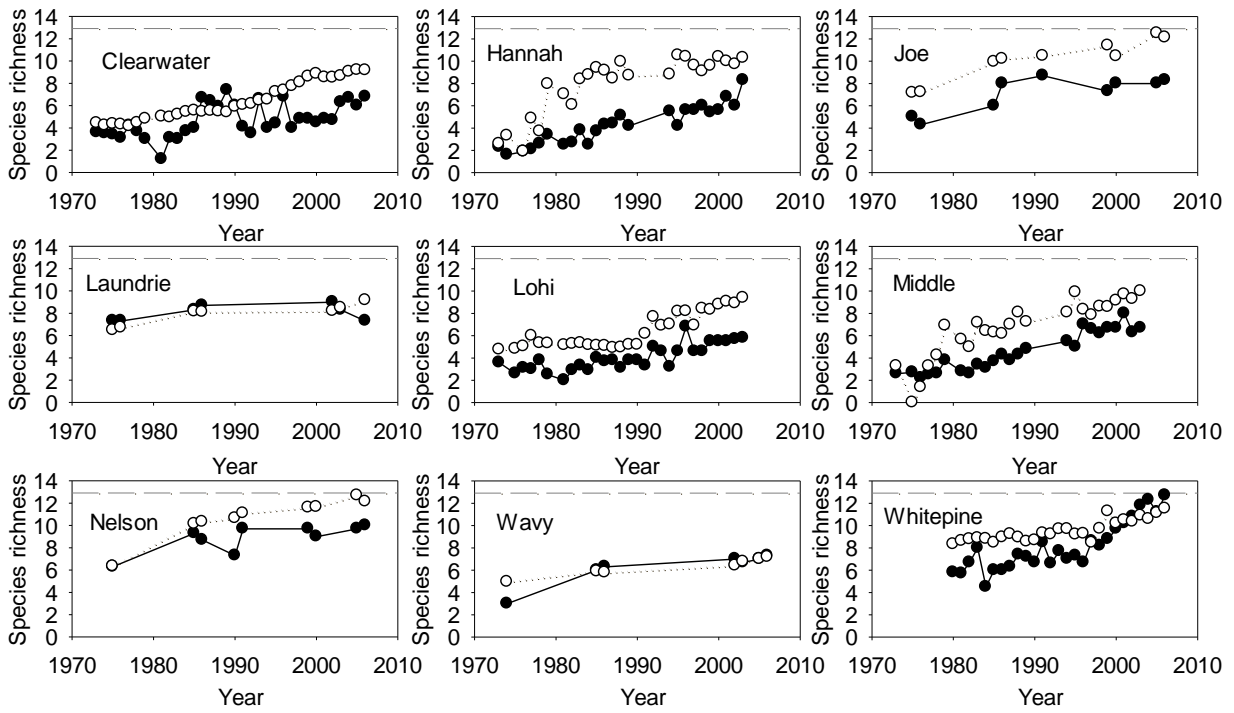


Figure 7 (part 1)

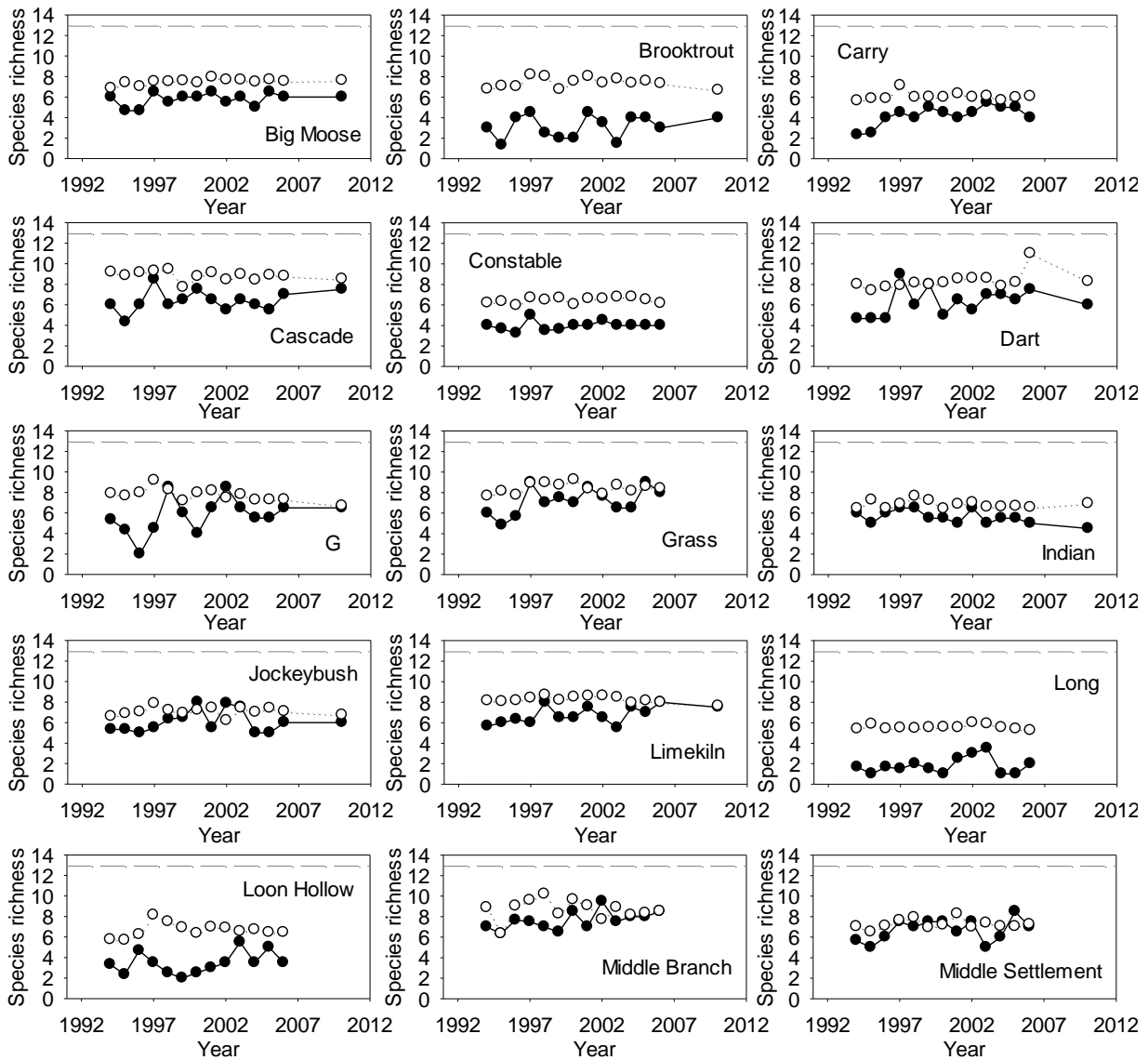


Figure 7 (part 2)

