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Bumblebee family lineage survival is enhanced in high quality landscapes

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Summary paragraph

Insect pollinators such as bumblebees (*Bombus* spp.) are in global decline^{1,2}, a major cause of which is habitat loss due to agricultural intensification³. A range of global and national initiatives aimed at restoring pollinator habitats and populations have been developed⁴⁻⁶. However, the success of these initiatives depends critically upon understanding how landscape change affects key population-level parameters, such as survival between lifecycle stages⁷, in target species. Such understanding is lacking for bumblebees because of the difficulty of

systematically finding and monitoring colonies in the wild. We used a novel combination of habitat manipulation, land-use and habitat surveys, molecular genetics⁸ and demographic and spatial modelling to examine between-year survival of family lineages in field populations of three bumblebee species. Here we show that the survival of family lineages from the summer worker to the spring queen stage in the following year increases significantly with the proportion of high-value foraging habitat, including spring floral resources, within 250-1000 m of the natal colony. This is the first evidence of a positive impact of habitat quality on survival and persistence between successive colony cycle stages in bumblebee populations. The findings provide strong support for conservation interventions that increase floral resources at a landscape scale and throughout the season having positive effects on wild pollinators in agricultural landscapes.

Main text

The loss of semi-natural habitats and floral resources within intensively managed agricultural landscapes has been identified as a major driver of declines in insect pollinators^{3,9,10}, with negative consequences for crop pollination¹¹. Habitat restoration (e.g. the planting of flowering hedgerows, meadows or flower strips along field margins under agri-environment schemes¹²) can mitigate these effects, increasing local pollinator abundance and species richness¹³⁻¹⁵ and enhancing rates of persistence and colonization at the community level¹⁶. However, we lack understanding of the effects of restoration on key aspects of pollinator biology that may explain the mechanisms behind these responses. In particular, improving habitat quality might be expected to enhance the prospects of successful reproduction and between-year survival in targeted areas, but whether this occurs is unknown.

Bumblebees (*Bombus* spp.) are key pollinators of wild flowers and commercial crops^{17,18}. Following a eusocial, annual colony cycle, new queens enter hibernation in the autumn and emerge in spring to search for a nest site and found a colony. Each colony may produce up to several hundred ‘daughter’ workers, which forage from spring to summer at flowers for nectar and pollen to rear new daughter queens and males¹⁹. The survival and dispersal patterns of bumblebee queens during hibernation and nest-searching periods are critical to overall population persistence, but remain undescribed in wild populations^{8,20,21}. In addition, although the availability of floral resources within foraging distance of the nest has been shown to increase numbers of workers and males produced per colony, effects on queen production have been less clear²² and there is no evidence regarding how queen production, survival and dispersal may be linked with underlying habitat quality and land-use²³.

Here, we investigated the effects of habitat quality and land-use on bumblebee survival and dispersal between colony cycle stages across two years. We first tested whether colonies located within or near high-value foraging habitats had a greater probability of producing daughter queens that survive the winter hibernation and spring emergence stages, henceforth termed ‘family lineage survival’. Second, we tested whether the distances travelled by queens between hibernation and nest-searching periods (as a measure of minimum relative queen dispersal distances within our study landscape) were affected by the proportion of high quality habitat surrounding their natal colony. We sampled DNA non-lethally from 537 spring queens (in 2011 and 2012) and 2,101 workers (in 2011) of three widespread species (*Bombus terrestris*, *B. lapidarius* and *B. pascuorum*) at a fine spatial scale across a 20-km² agricultural landscape in southern England, UK. The landscape was dominated by arable fields and permanent intensive grassland but also included a range of habitat restoration measures for pollinators²⁴, which resulted in a heterogeneous matrix with areas of high and low proportions

of flower-rich habitat¹⁵. Sampled queens and workers were genotyped at 13-14 microsatellite loci per species and family relationships between them were estimated using maximum likelihood sibship reconstruction⁸. This revealed a total of 1,665 family lineages within our sample. Relationships were detected within and between generations in the colony cycle: between spring (2011) queens and their daughter (2011) workers (i.e. within a single colony cycle), and between spring (2011) queens and their daughter (2012) queens sampled the following year or summer (2011) workers and their sister (2012) queens sampled the following year (i.e. between two successive colony cycles, Fig. 1; Extended Data Table 1). These relationships were used to estimate family lineage survival in terms of rates of ‘apparent survival’²⁵, i.e. the probability that a family lineage survives and remains available for capture within the study landscape (see Methods and Extended Data Tables 2, 3). Family lineage survival between the summer (2011) worker and spring (2012) queen stages was then related to measures of habitat quality and land-use at four spatial scales (relative to colony locations estimated from the distributions of sampled sister workers²⁴). We focused on the family lineage relationship between the summer worker and spring queen stage, since only data from summer workers permitted the estimation of colony locations (see Methods).

We found that, across all three bumblebee species, habitat quality and land-use variables were significantly positively correlated with the between-year survival of family lineages (Extended Data Tables 4-6). Mixed semi-natural vegetation cover, queen-visited spring flower cover (as provided, for example, by flowering trees and hedgerow plants) and summed spring and summer flower cover for bumblebee-visited species had highly significant positive effects on family lineage survival within radii of 250 m, 500 m and 1,000 m from estimated colony locations (Fig. 2; Extended Data Table 4). While the individual colony locations estimated using our approach may be subject to some error, we found no evidence to suggest any

systematic bias²⁴ which would affect relationships with habitat across the large sample sizes that were analysed. The strongest relationship was for mixed semi-natural vegetation cover (including sown field margins) within a radius of 1,000 m (slope = 26.17; LRT statistic = 11.34; $P < 0.001$; Fig. 2a). We also found significant positive effects of high-value foraging habitats (queen-visited spring flower cover and summed spring and summer flower cover) within radii equal to the estimated colony-specific foraging distance (Fig. 2; Extended Data Table 4). Given that colony-specific foraging distances are reduced as floral resources increase²⁴, this result suggests that when workers were able to forage closer to their colonies, greater numbers of queens from those colonies survived the winter and spring emergence periods. Our findings also suggest that family lineage survival is particularly sensitive to small changes in landscape composition (Fig. 2).

Family lineage survival was not significantly influenced by all spring or all summer flower cover (flower cover of all surveyed plant groups in either season), worker-preferred summer flower cover or nesting habitat cover within the ranges of variation tested across the study landscape at any modelled radius (Extended Data Tables 4 and 5). This suggests that summer-flowering resources at the worker stage alone are not sufficient to sustain colonies throughout their cycle. Moreover, it suggests that flower cover of particular plant groups utilised as forage resources (Extended Data Table 6) is more important than overall flower cover within the habitats surrounding a colony. We did, however, find a weak but significant positive effect on family lineage survival of arable field cover within 1,000 m of estimated colony locations (Extended Data Table 4); this is most likely due to the presence of spring-flowering oilseed rape and summer-flowering field bean crops across the landscape (Extended Data Table 5).

Across all three bumblebee species, the mean (\pm s.e.) distance of sampled 2012 spring queens from their estimated natal colony location was 1227 ± 125 m. This distance, taken as a measure of minimum relative queen dispersal distance, is between two to three times greater than the typical foraging distances estimated for workers within the same landscape²⁴. Mean minimum relative queen dispersal distances were greatest for *B. terrestris* (1553 ± 223 m, from $n = 15$ colonies), and lower for *B. pascuorum* (1149 ± 273 m, $n = 12$) and *B. lapidarius* (980 ± 148 m, $n = 16$). Although these differences were not statistically significant (ANOVA, $F_{2,42} = 2.07$; $P = 0.14$), the three bumblebee species differed significantly in the proportion of 2012 spring queens from family lineages not represented in the previous year's dataset (82% in *B. terrestris* and 56% in both *B. pascuorum* and *B. lapidarius*, $\chi^2_2 = 31.06$; $P < 0.01$) (Extended Data Table 1). Taken together, these findings suggest that *B. terrestris* is the most widely dispersing of the three species.

Despite this variation in minimum relative queen dispersal distance within species and, potentially, between species, there were no significant correlations at any spatial scale between this distance and any seasonal flower cover variable, mixed semi-natural vegetation cover or arable field cover (Extended Data Table 5). However, there were significant positive correlations between queen dispersal distance and nesting habitat cover at radii of 250 m, 500 m and 1,000 m (Kendall's tau correlation coefficients = 0.309, 0.308 and 0.331, respectively; $df = 41$; $P < 0.05$). This suggests that while high quality foraging habitats may not extend the dispersal distances of queens, non-crop habitats suitable for nesting may facilitate queen movement into the wider landscape.

Our study is the first to demonstrate a positive impact of habitat quality on survival and persistence between successive colony cycle stages in wild pollinators. Given that two of the

habitat variables having the strongest influence, mixed semi-natural vegetation and summed spring and summer flower cover, included sown field margins (Fig. 2; Extended Data Table 4), this impact includes effects of habitat restoration via the implementation of agri-environment schemes. The study provides evidence that floral resources in spring-flowering trees, hedgerow plants and crops are particularly vital for bumblebee queens during their emergence and colony founding, and, in combination with summer floral resources, enhance the representation of colonies in the population the following year. It therefore adds to previous studies in highlighting the importance of temporally sustained floral resources within 1 km of nests for within-season survival and performance of bumblebee colonies^{22,26,27}, especially spring-flowering resources, which are often overlooked in conservation intervention options²⁸. Furthermore, our findings suggest that, as well as sustaining colonies, appropriately managed non-crop areas can act as a source of queens to the wider landscape²⁹. While there is an urgent need for systematic monitoring of pollinator populations to provide more robust data on patterns and causes of decline³⁰, our study provides strong support for conservation interventions targeted at a landscape scale having a positive impact on wild pollinators in agricultural landscapes.

Supplementary Information is linked to the online version of the paper at www.nature.com/nature.

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Author Contributions. C.C., M.H., A.B., S.S. and W.J. (the late Bill Jordan) conceived the project and designed the study. C.C., M.H. and S.H. coordinated the fieldwork and modelling elements and C.C. prepared the manuscript. S.D. carried out the molecular genotyping and sibship assignments with guidance from J.W. J.R. developed and applied the spatial analyses and S.F. designed and undertook the statistical analyses. All authors contributed to writing and critiquing the manuscript.

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METHODS

Mapping the study landscape. The study was conducted across an agricultural landscape covering 20 km², centred on the Hillesden Estate, Buckinghamshire, UK (1°00'01"W; 51°57'16"N)⁸. The Estate consists of a c. 1,000 ha intensive arable farm on which a number of experimental targeted habitat restoration options (including the sowing of wildflower mixtures for pollinators along field margins) have been established and managed since 2005 under the English agri-environment scheme¹². These 'sown field margins' comprised 2% of the total area of the study landscape, although their density was manipulated spatially so that, per 50-60 ha of land, the area taken out of production varied systematically from 0-8%¹⁵. The landscape surrounding the Estate was predominantly arable, with some areas of permanent intensive grassland, woodland and small villages.

Detailed habitat maps of the study landscape were generated using a land use/land cover (LULC) map derived from two airborne remote sensed sources - Light Detection and Ranging (LiDAR) and hyperspectral imaging³¹. These data were combined and classified to form a high-resolution (0.5 x 0.5 m pixels) LULC map with each pixel assigned to one of 9 land cover classes (arable; short grass; non-woody semi-natural mixed vegetation; agri-environmental field margin; garden and urban vegetation; woody vegetation; road and building; water; and bare soil). The study landscape was surveyed systematically in terms of its value for bumblebees in both spring and summer (during and immediately following our sampling of queens and workers, respectively). The LULC map was converted to vector format and every resultant LULC polygon which could be distinguished as a discrete habitat parcel in the field was surveyed in July and August 2011, to estimate a) the percentage cover of each plant species, family or group (hereafter plant groups), and b) the proportion of that plant group in flower at the time of the survey. These variables (a and b) were multiplied for each plant group to give a measure of the proportion of each habitat parcel covered with flowers. To construct a

spring habitat map, we surveyed a stratified random sub-sample of parcels across all habitat types in April 2011 and 2012. These data were then used to estimate plant and floral cover values in unsampled parcels by adjusting species cover estimates from the full summer dataset.

An assessment of the suitability of each land parcel as nesting habitat for bumblebees was made by estimating average vegetation height (m) across the whole parcel, whether tussocky vegetation was present, the extent of plant litter or moss within the sward and whether there were signs of small mammal activity such as the presence of burrows, runs or faeces. Species-specific nesting requirements (classified a priori using expert knowledge and published sources^{19,32}) were then used to categorise each parcel as being of high, medium or low nesting habitat suitability for each of the three study species.

The vector LULC map was updated using the digitized, completed survey maps, and linked to the floral and nesting data. In total, 18.7 km² of the study landscape were surveyed in this way. Where parcels were not surveyed (due to access restrictions), plant cover values were estimated by taking the average value of covers from parcels of the same LULC class within a 500 m radius. The final habitat dataset thus consisted of all discrete parcels with information on cover and floral cover of all surveyed plant groups³¹. Floral cover of surveyed plant groups was further summed in terms of the plants' relative value as forage resources for bumblebees, according to whether they had been observed to be visited or not visited by foraging queens or workers during bee sampling, and, if visited, whether they were classified as 'preferred' in the worker dataset (Extended Data Table 6). Handling of the LULC map and survey data was performed in ArcMAP v10.0 (© ESRI, Redlands, CA, USA).

Study species. We studied three social bumblebee (*Bombus*) species that are common and widespread across much of the UK but vary in their forage plant choice and nesting behaviour. *Bombus terrestris* L. and *B. lapidarius* L. typically nest underground in large colonies (reaching up to 300 workers at maturity) and have shorter-tongued workers that visit a wide range of

flowers. *B. pascuorum* (Scopoli) nests on the ground surface, usually within mossy grass tussocks tending to form smaller colonies (reaching between 100 and 200 workers) with longer-tongued workers that specialize in foraging at flowers with long corolla tubes¹⁹. There is limited evidence regarding numbers of queens produced in wild colonies in the UK, and these are likely to vary widely within and between species but, in *B. terrestris*, studies suggest averages of 4, 14 and 35 queens produced per colony³³⁻³⁵.

Bumblebee sampling and genotyping. The study area was divided into 250 × 250 m grid cells and within every cell sampling intensity (i.e. search effort) for both queens and workers was proportional to the relative cover of suitable habitats present. Hence searches were more focussed on field boundaries and other non-crop habitat parcels (defined areas of continuous land use) but did include field interiors. All female individuals of the three study species encountered were caught for DNA sampling. At the same time, their locations were recorded using a GPS device and their behaviours noted (whether nest-searching, in flight or foraging and, if foraging, which plant species was being visited). Sampling was performed at this fine spatial scale to maximise the likelihood of detecting sister workers at multiple sites³⁶ and to ensure a high proportion of colonies were sampled across the landscape.

We obtained DNA samples non-lethally by clipping the tarsal tip of a mid-leg of each bee and preserving it in 100% ethanol³⁷. If a bee was encountered that had already been sampled, which occurred in three queens and <10 workers, we identified it to individual level by taking a second DNA sample from the basitarsus of the same mid-leg that had already been clipped and seeking a match for its multilocus genotype in the dataset. Sampling was carried out between 09:00h and 17:00h during dry weather when ambient temperature was above 11°C with at least 60% clear sky, or above 15°C under any sky conditions.

Queens were sampled across the study landscape from 21 March to 18 April 2011 and from 19 March to 2 May 2012, reflecting the main periods of spring emergence and nest founding

activity observed for the three study species. Workers of the three species were sampled continuously (during 4-5 days per week) between 20 June and 5 August 2011 across all the habitats in the study landscape. We used PCR-based molecular identification of mtDNA markers to differentiate *B. terrestris* workers from any *B. lucorum* workers present in our sample (H.M.G. Lattorff, personal communication), since workers of the two species are difficult to separate reliably in the field. DNA was isolated from each tarsal sample using the HotSHOT protocol, and all individuals were genotyped at 14 (*B. terrestris* and *B. pascuorum*) or 13 (*B. lapidarius*) microsatellite loci (see Dreier et al., 2014^{8,38}). Missing data were rare across both queen and worker genotypes, with all except a single individual worker of *B. terrestris* included in analyses being successfully typed at 12 or more loci. The mean missing genotype frequency per individual across all species was 0.002.

Assigning family relationships. COLONY version 2.0³⁹ was used to reconstruct family relationships between the bees in our sample for each species (Fig. 1a). COLONY implements a full-likelihood approach to sibship analysis, and yields the best (maximum likelihood) and possible alternative (high likelihood) estimates of family relationships with corresponding estimated posterior probabilities. The inferred relationships with an estimated posterior probability of 0.8 or higher were accepted and used in downstream analyses⁸. Given previous work⁴⁰ we assumed a monogamous mating system for males and females, therefore allowing the assignment of full-siblings, mothers and daughters. We carried out a medium-length run with medium-likelihood precision, using genotyping error rates of 0-5% based on results of re-genotyping and re-scoring 10% of randomly selected individuals⁸. Two replicate COLONY runs were conducted across the full sample of queens and workers, each with a different random number seed but with all other parameters kept equal (membership of sibship families was identical in both COLONY runs and the variance between the estimated probabilities of inference was small (<0.002)). Reconstructed relationships included those between sister

workers sampled in summer 2011 and between: a) founding queens sampled in spring 2011 and their daughter workers sampled in summer 2011; b) queens sampled in spring 2011 and their daughter queens sampled in spring 2012; and c) workers sampled in summer 2011 and their sister queens sampled in spring 2012 (Fig. 1a).

Estimating colony locations and surrounding landscape quality. The location of each sampled queen and worker was mapped from GPS locations in ArcGIS. We estimated the location of each colony from the distributions of full-sib sister workers, using a mean centre approach²⁴. Model simulation showed that this approach was not biased by either outlying worker locations or clusters of workers at particular foraging patches, and it required no additional parameters or prior assumptions (e.g. regarding likely foraging distances)^{24,36}. Colony locations were estimated only for colonies represented by inferred sibships of two or more workers, since it is not possible to assign a biologically meaningful nest location for colonies represented by single workers.

Mean centre locations were ‘snapped’ (ie. moved to coincide exactly with the coordinates of another feature) to the nearest LULC class that might have formed suitable nesting habitat for bumblebees, thus avoiding cropped arable fields, roads, buildings and water¹⁹. Overall, 214 estimated colony locations required snapping (47% of the sample); 208 of these were from cropped arable fields (mean snapping distance, i.e. distance between original location and ‘snapped’ location, = 47.2 m) and 6 were from the ‘roads and buildings’ land-use category (mean snapping distance = 4.9 m). The straight-line distance of each worker from its capture location to estimated colony location was calculated and the mean of these distances for all workers in a sibship was used as a measure of ‘colony-specific foraging distance’²⁴. The snapping process made a mean difference to colony-specific foraging distances of only 4.3 m. Furthermore, random resampling of pairs of sister workers from sibships with more than two workers showed that sibship size had only a minor influence on estimated colony locations and

foraging distances; hence there was no evidence to suggest a systematic bias that might have affected relationships with habitat²⁴.

Measures of habitat quality and land-use at different spatial scales surrounding each colony were made by creating a series of buffers with the following radii around the colony location: colony-specific foraging distance; 250 m; 500 m; and 1,000 m. Within each buffer, the proportion of each LULC class, floral cover of different subsets of plant groups in spring and summer (including queen-visited and worker-preferred plant groups), and cover of suitable nesting habitats were then determined. Mixed, non-woody, semi-natural vegetation, agricultural environmental sown field margins and other linear habitats were combined to a single 'mixed semi-natural vegetation' class (Extended Data Table 5). The range of variation in these variables across our study landscape, and correlations between them, are shown in Extended Data Figure 1 for the proportion of each variable within 1,000 m of colony locations.

Estimating queen dispersal distance. The distances travelled by queens of *B. terrestris*, *B. lapidarius* and *B. pascuorum* between departure from their natal nest in late summer and post-hibernation colony foundation in the spring were estimated by measuring geographic distances between the 2011 colony locations and their inferred sister queens sampled in spring 2012. For families in which two or more queens were sampled, the average distance between queens and the natal colony location was calculated. We did not estimate the nest-site locations of founding queens at the time of spring sampling, since it was not possible to accurately separate queens that had already founded nest-sites and were foraging from them from queens that were actively searching for nest sites i.e. still in their dispersal phase.

Statistical analyses. We developed a novel extension of the standard Cormack-Jolly-Seber (CJS) mark-recapture model²⁵ to estimate survival rates and 'recapture' probabilities (i.e. the probability of a daughter worker or queen being sampled from a given family lineage) using data on the numbers of individuals sampled at every stage for every family lineage of each

species (Figure 1 and Extended Data Table 1). This allowed us to estimate family lineage survival, corrected for under-detection of individuals arising, for example, from long-distance dispersal or non-capture of queens. Where the standard CJS model has recaptures relating to individual animals, our modified model related recaptures to different individuals from the same family lineage and incorporated counts of the number of individuals (full-sib workers or sister queens) sampled at each stage. A full step-by-step account of model construction is given in Supplementary Information (SI) under section ‘A novel mark-recapture model for colonial species’, along with the code used for model fitting.

The standard CJS model of Equation (1) was first fitted for the three *Bombus* species separately and for all species combined ($n = 1,665$ family lineages) using the software package MARK⁴¹. For *B. lapidarius* no first-generation queens were captured and the model (1) collapsed to a simple binomial. We therefore estimated the single parameter for this species in isolation via a GLM routine. Performance of this simple, standard CJS model proved to be poor and only ϕ_1 and p_1 were uniquely estimable (Extended Data Table 2). We therefore improved the model using the modified form (SI Table 1; Extended Data Figure 2), which incorporated extra information contained in the colony counts of workers and queens the following spring and introduced the Poisson distribution to the model. All parameters were estimated with greater precision with data for all species combined, and with parameters common across all species (row 1, Extended Data Table 3), and the model captures the general variation in counts observed for the two castes (Extended Data Figure 3). Sampling correlations between the model parameters further demonstrate that survival and detection probability can be estimated separately (SI Table 2). Using the maximised log-likelihood values for these models ($-\log(L)$, Extended Data Table 3) suggested that there were between-species differences in the parameters (likelihood-ratio test statistic of $2 \times (2045.6 - 467.6 - 538.6 - 992.6) = 93.6$, which is significant on the basis of a χ^2 distribution with 7 degrees of freedom ($P < 0.01$)). However,

examination of the confidence limits suggested that the mean counts of workers (λ_1) and queens (λ_2) caught were estimated more precisely than the survival probabilities (ϕ_1 and ϕ_2). Thus data from all three species were combined for further analyses, with different mean counts but shared survival rates, to remove overparameterisation and increase precision.

Finally, relationships between habitat/land-use variables and survival (ϕ_2) were explored for families with $W_{1i} > 1$ for which colony locations could be estimated ($n = 456$). To further test for any species-specific differences in survival among this reduced sample, we established a baseline model (Model A) in which ϕ_2 was constant but, motivated by the above likelihood-ratio test statistic and confidence limits, λ_1 and λ_2 were permitted to vary between species. A formal comparison of Model A with a model including species-specific differences in survival (Model B) confirmed that these were not significant ($\chi^2_2 = 3.95$, $P = 0.14$). Therefore families of all three species ($n = 456$) were pooled for extensions of Model A with ϕ_2 regressed against each of the habitat or land-use variables at different radii from the colony locations in turn. Those models leading to a significant improvement in fit when compared to Model A were considered to demonstrate significant effects of habitat or land-use variables on the probability of family lineage survival (Fig. 2; Extended Data Table 4). These regressions were repeated using a complementary log-log link function which did not affect the model outcomes, with the sign of the slopes remaining the same and the maximum log-likelihood values remaining similar to those presented in Extended Data Table 4.

Data Availability: Datasets are available from the NERC Environmental Information Data Centre (EIDC) as follows: Family lineage and landscape quality data for wild bumblebee colonies across an agricultural landscape in Buckinghamshire, U.K. (<http://doi.org/10.5285/6be00174-6544-4156-b1df-8678f6df2034>); Map of land-use/land-cover and floral cover across an arable landscape in Buckinghamshire, UK (<http://doi.org/10.5285/0667cf06-f2c3-45c1-a80a-e48539b52427>); Microsatellite genotype

data for five species of bumblebee across an agricultural landscape in Buckinghamshire, UK. (<http://doi.org/10.5285/6a408415-0575-49c6-af69-b568e343266d>); Location data of worker bumblebees across an agricultural landscape in Buckinghamshire, UK. (<http://doi.org/10.5285/a60f52b8-0f9f-44f6-aca4-861cb461a0eb>). Reprints and permissions information is available at www.nature.com/reprints. The authors declare no competing financial interests. Correspondence and requests for materials should be addressed to ccar@ceh.ac.uk.

Code availability. The full R code for simulating a dataset and fitting the modified CJS model is available in Supplementary Information: Carvell_Nature_Supplementary Information_Msimulation.R.

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Figure 1. Bumblebee colony cycle stages (grey boxes) and family lineages sampled in the study with estimated survival parameters. **a)** Relationships (dotted lines, italics) assigned using sibship reconstruction from non-lethal DNA sampling. **b)** Family lineage survival (arrows: between the mother queen and daughter worker stage in spring and summer 2011 (ϕ_1), and between the summer worker and sister queen stage in spring 2012 (ϕ_2), with estimated asymptotic 95% confidence limits) and recapture probability (p_1, p_2) parameters (see Extended Data Table 3). Parameter values shown are (as examples) mean values across 1,665 family lineages of all species (*Bombus terrestris*, *B. lapidarius* and *B. pascuorum*). Only females (shown in red) were sampled; hollow female symbol denotes queens, and female symbols containing a star denote workers; males (shown in blue) were not sampled.

Figure 2. Effects of habitat quality and land-use variables on bumblebee family lineage survival from the summer worker to spring queen stage (parameter ϕ_2 in Fig. 1b). Solid line shows model-fitted logistic regression; dashed lines show 95% confidence intervals. Predictor variables are proportions of: a) mixed semi-natural vegetation cover (including sown field margins) within 1,000 m of colony locations ($\chi_1^2 = 11.34$; $P < 0.001$); b) queen-visited spring flower cover within colony-specific foraging distance ($\chi_1^2 = 9.52$; $P < 0.01$); c) summed spring and summer flower cover within colony-specific foraging distance ($\chi_1^2 = 7.2$; $P < 0.01$); and d) arable field cover within 1,000 m of colony locations ($\chi_1^2 = 4.3$; $P < 0.05$). See Extended Data Table 5 for full descriptions of predictor variables. Relationships at radii showing best model fit for each variable are presented; see Extended Data Table 4 for model results for all variables at all radii. Data generated from 456 wild colonies of *Bombus terrestris* ($n = 69$), *B. lapidarius* ($n = 267$) and *B. pascuorum* ($n = 120$). Model comparisons showed no significant differences between species in apparent survival at this stage, hence data from all species were combined (see Methods).

Figure 1

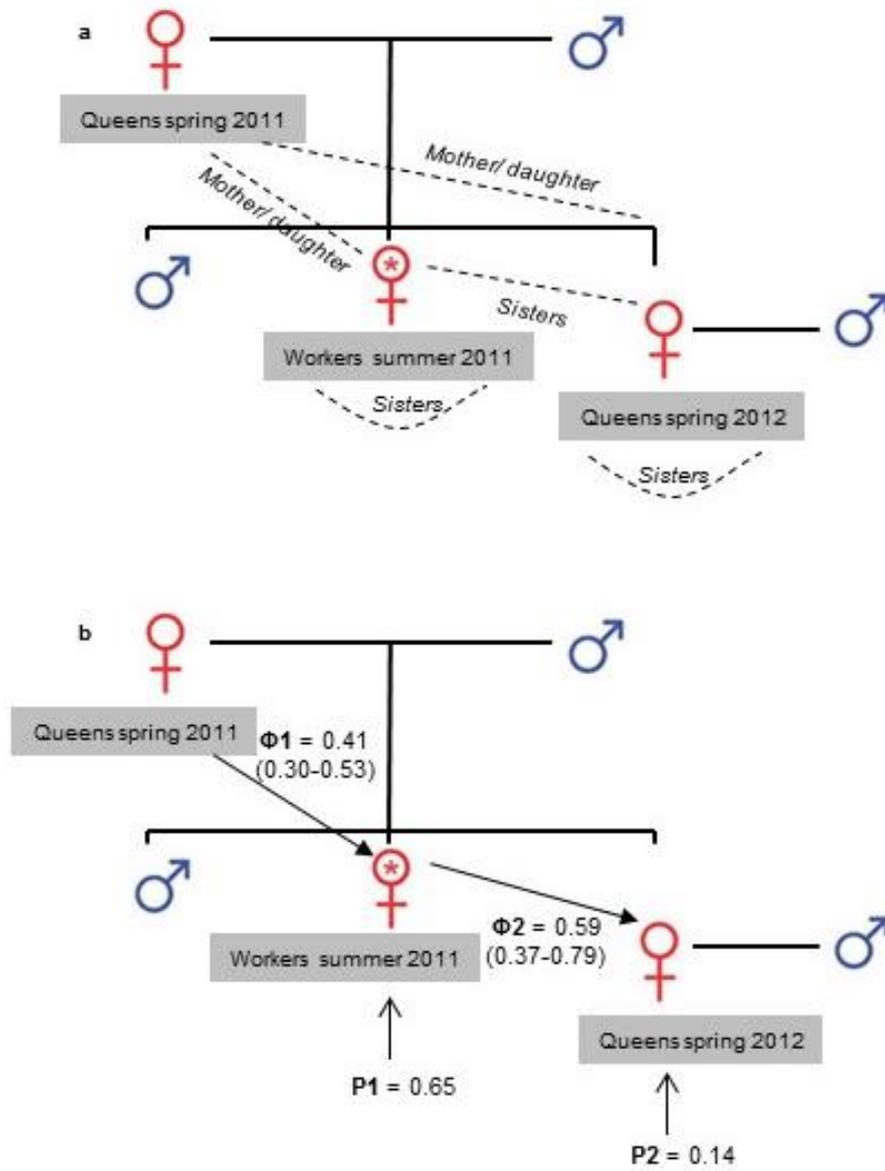
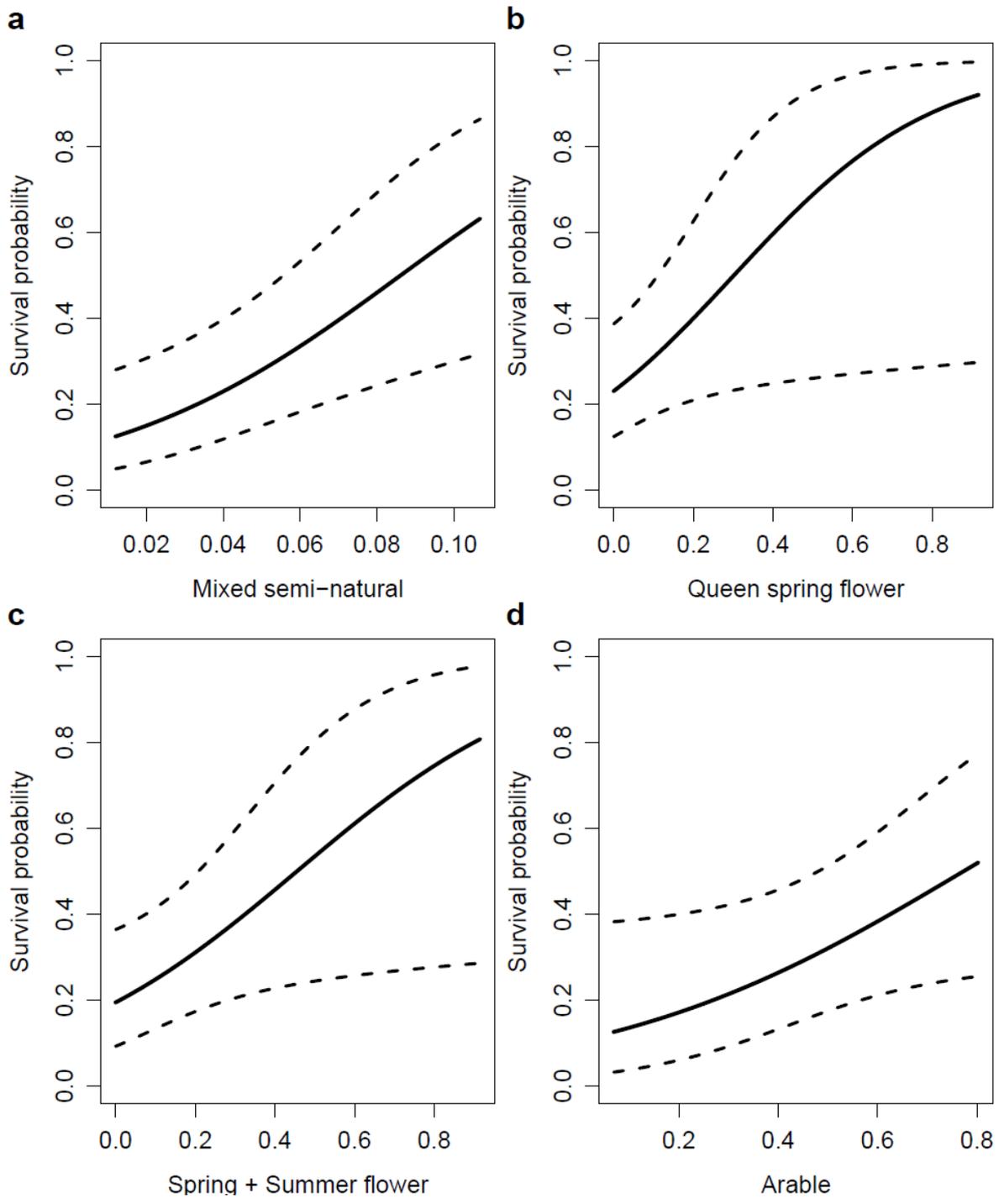


Figure 2



Extended Data Table 1. Numbers (and percentages) of bumblebee (*Bombus* spp.) colonies and lineages detected within each family relationship category. * includes from sampled queens and workers. *B. lapidarius* queens were not sampled in 2011.

Extended Data Table 2. Initial model results showing estimated survival and detection parameters for the three study bumblebee (*Bombus*) species. Column headings are: estimated colony survival rate ϕ_1 (mother queen to daughter worker stage); recapture probability p_1 (worker stage); and combined survival and recapture probability θ (summer worker to spring queen stage). Estimated via a Cormack-Jolly-Seber type mark-recapture model fitted in MARK⁴¹ except *where estimated as a simple binomial GLM from the numbers of colonies identified at the worker stage that produced records of queens the following year. Estimated asymptotic 95% confidence limits (based on the Hessian matrix and back-transformed) given in parentheses.

Extended Data Table 3. Estimated probabilities of survival and detection of bumblebee (*Bombus* spp.) family lineages using the modified mark-recapture model (see Methods and Supplementary Information). Column headings are: apparent survival rates ϕ_1 , ϕ_2 ; expected numbers of workers (λ_1) and spring queens (λ_2) caught per surviving family lineage; probabilities of at least one individual worker (p_1) or queen (p_2) per surviving family lineage being captured in the study landscape (recapture probabilities are calculated as $(1 - \text{the probability of no individuals being caught})$, based upon the estimates of λ); combined probabilities of survival and recapture, calculated as $\phi_1 p_1$ and $\phi_2 p_2$, respectively; and $-\log(L) =$ maximised log-likelihood values for each model. Estimated asymptotic 95% confidence limits (based on the Hessian matrix and back-transformed) given in parentheses. Parameter values

were estimated from all sampled families ($n = 1,665$), including those represented by single workers.

Extended Data Table 4. Model results for logistic regression of apparent survival (probability of bumblebee family lineage survival from the summer colony to spring queen stage (ϕ_2)) against habitat quality and land-use variables (Extended Data Table 5) at four spatial scales. $n = 456$ family lineages, excluding those represented by only queens or by single workers for which colony locations could not be estimated. Significant variables and P values are shown in bold. Foraging distance is the mean of the straight-line distances of each worker from its capture location to its estimated colony location.

Extended Data Table 5. Habitat quality and land-use variables for which effects on bumblebee family lineage survival and queen dispersal distance were tested. All variables were calculated as proportions of cover represented by the given category out of the total land area within a given radius (see Methods).

Extended Data Table 6. Plant groups used for field survey of habitats across the study landscape.

*Y/N denotes plant groups visited/not visited by foraging queens or workers during bee sampling.

#Y/N denotes preferred/not preferred forage plant groups of workers during sampling, preferred plants groups being identified as the five plant groups with the highest mean number of worker visits (across all three *Bombus* species) per plant species within that group.

Extended Data Figure 1. Variation and correlations between habitat and land-use variables across the study landscape. Plots show i) histograms to demonstrate variation within each habitat/land-use variable along the diagonal; ii) scatter plots showing correlations between variables (top right) with a linear model trend line fitted to the correlation data (in red, only extended to the limits of the data) and a 1:1 line (in pale grey) and iii) correlation coefficients with their significance (bottom left) where $P < 0.05 = *$, $P < 0.01 = **$ and $P < 0.001 = ***$. Axis values are standardised and represent proportional cover of the different habitat variables within 1,000 m of estimated colony locations, with variable names following the same order and shortened format as presented in Extended Data Table 5. Each point on the scatter plots represents one family lineage ($n = 456$).

Extended Data Figure 2. Simulation-based assessment of robustness of the modified CJS model. This shows that the estimated parameter values aggregate around the true values. Frequency distributions of parameter estimates are shown, from 1,000 simulated data sets, each of 2,000 families. Parameters plotted are (a) φ_1 : true value = 0.6 (b) φ_2 : true value = 0.5 (c) λ_1 : true value = 3 (d) λ_2 : true value = 2. To align with the real data in which some families were not detected at the founding queen (Q1) stage, if at all, data were simulated assuming a detection probability of 0.4 at the Q1 stage.

Extended Data Figure 3. Goodness of fit for the model of Extended Data Table 3. Frequency distributions across all species of a) observed counts of workers (W_{1i}); b) expected counts of workers (W_{1i}); c) observed counts of second-generation queens (Q_{2i}) and d) expected counts of second-generation queens (Q_{2i}).

Extended Data Table 1.

| Relationship category | <i>B. terrestris</i> | <i>B. lapidarius</i> | <i>B. pascuorum</i> |
|--|-----------------------------|-----------------------------|----------------------------|
| Queens 2011(total colonies detected) | 80 | | 37 |
| Queens 2011to daughter workers 2011 | 6 (7.5%) | | 20 (54.1%) |
| Queens 2011to daughter queens 2012 | 7 (8.8%) | | 3 (8.1%) |
| Workers 2011(total colonies detected) | 264 | 668 | 360 |
| Workers 2011to sister queens 2012 | 31(11.7%) | 33 (4.9%) | 35 (9.7%) |
| Queens 2012 (total colonies detected) | 216 | 75 | 87 |
| Queens 2012 from retained 2011families* | 38 (18%) | 33 (44%) | 38 (44%) |
| Queens 2012 from previously unknown families | 178 (82%) | 42 (56%) | 49 (56%) |

Extended Data Table 2.

| Species | φ_1 | ρ_1 | $\theta (= \varphi_2\rho_2)$ |
|------------------------|-------------------------------|----------------------------|--|
| All | 100 (-) | 0.21(0.15-0.29) | 0.08 (0.07-0.09) |
| <i>B. terrestris</i> | 0.77 (0.13 – 0.99) | 0.09 (0.03 – 0.23) | 0.12 (0.08 – 0.16) |
| <i>B. lapidarius</i> * | - | - | 0.05 (0.04 – 0.07) |
| <i>B. pascuorum</i> | 100 (-) | 0.51(0.36 – 0.66) | 0.10 (0.07-0.13) |

Extended Data Table 3

| Species | φ_1 | φ_2 | λ_1 | λ_2 | ρ_1 | ρ_2 | $\varphi_1\rho_1$ | $\varphi_2\rho_2$ | $-\log(L)$ |
|----------------------|-------------------------------|-------------------------------|-------------------------------|-------------------------------|----------------------------|----------------------------|-------------------------------------|-------------------------------------|------------------------------|
| All | 0.41 (0.30-0.53) | 0.59 (0.37-0.79) | 104 (0.98-111) | 0.15 (0.10-0.21) | 0.65 | 0.14 | 0.27 | 0.08 | 2045.6 |
| <i>B. terrestris</i> | 0.24 (0.14-0.39) | 0.86 (0.14-100) | 0.72 (0.62-0.88) | 0.17 (0.11-0.27) | 0.51 | 0.16 | 0.12 | 0.14 | 467.6 |
| <i>B. lapidarius</i> | - | 100 (-) | 124 (-) | 0.05 (-) | 0.71 | 0.05 | - | 0.05 | 992.6 |
| <i>B. pascuorum</i> | 0.94 (0.16-100) | 0.56 (0.22-0.85) | 0.89 (0.77-101) | 0.20 (0.11-0.38) | 0.59 | 0.18 | 0.55 | 0.1 | 538.6 |

Extended Data Table 4

| Habitat / land-use variable | Radius from colony location (metres) | Minimum negative log-likelihood | LRT statistic | P-value (χ^2 , 1 df) | Slope | Lower 95% Confidence Interval | Upper 95% Confidence Interval |
|--|--------------------------------------|---------------------------------|---------------|----------------------------|--------|-------------------------------|-------------------------------|
| All spring flower cover | 250 | 736.94 | 1209 | 0.272 | 0.942 | -0.80 | 2.68 |
| | 500 | 737.38 | 0.317 | 0.573 | 0.574 | -1.45 | 2.60 |
| | 1000 | 737.54 | 0.006 | 0.941 | 0.139 | -3.53 | 3.81 |
| | Foraging distance | 736.97 | 1.145 | 0.284 | 1.087 | -0.98 | 3.16 |
| All summer flower cover | 250 | 736.31 | 2.450 | 0.117 | 1.254 | -0.40 | 2.91 |
| | 500 | 737.10 | 0.871 | 0.351 | 0.922 | -1.05 | 2.89 |
| | 1000 | 737.42 | 0.249 | 0.618 | 0.961 | -2.84 | 4.77 |
| | Foraging distance | 736.61 | 1.850 | 0.174 | 1.345 | -0.77 | 3.46 |
| Queen-visited spring flower cover | 250 | 734.16 | 6.767 | 0.009 | 2.641 | -0.22 | 5.51 |
| | 500 | 734.17 | 6.730 | 0.009 | 3.221 | -0.60 | 7.04 |
| | 1000 | 733.07 | 8.943 | 0.003 | 9.024 | -2.20 | 20.25 |
| | Foraging distance | 732.78 | 9.523 | 0.002 | 3.986 | 0.27 | 7.70 |
| Spring + Summer flower cover | 250 | 734.70 | 5.685 | 0.017 | 2.260 | -0.02 | 4.54 |
| | 500 | 734.57 | 5.950 | 0.015 | 2.870 | -0.16 | 5.90 |
| | 1000 | 734.01 | 7.050 | 0.008 | 7.337 | -2.61 | 17.29 |
| | Foraging distance | 733.94 | 7.205 | 0.007 | 3.115 | 0.22 | 6.01 |
| Worker-preferred summer flower cover | 250 | 737.20 | 0.687 | 0.407 | -2.307 | -7.85 | 3.23 |
| | 500 | 737.21 | 0.652 | 0.420 | -2.651 | -9.17 | 3.87 |
| | 1000 | 736.11 | 2.869 | 0.090 | -9.514 | -21.10 | 2.07 |
| | Foraging distance | 737.22 | 0.642 | 0.423 | -1.682 | -5.88 | 2.51 |
| Arable field cover | 250 | 737.16 | 0.762 | 0.383 | 0.588 | -0.75 | 1.93 |
| | 500 | 737.19 | 0.692 | 0.406 | 0.723 | 0.00 | 0.00 |
| | 1000 | 735.39 | 4.307 | 0.038 | 2.748 | -0.04 | 5.53 |
| | Foraging distance | 737.09 | 0.893 | 0.345 | 0.715 | -0.80 | 2.23 |
| Mixed semi-natural vegetation cover | 250 | 733.89 | 7.300 | 0.007 | 6.871 | -4.68 | 18.42 |
| | 500 | 733.32 | 8.446 | 0.004 | 15.226 | -10.10 | 40.56 |
| | 1000 | 731.87 | 11.337 | 0.001 | 26.169 | 8.60 | 43.74 |
| | Foraging distance | 735.85 | 3.373 | 0.066 | 1.866 | -0.58 | 4.31 |
| Nesting habitat cover | 250 | 737.36 | 0.363 | 0.547 | -0.500 | -2.16 | 1.16 |
| | 500 | 737.08 | 0.927 | 0.336 | -0.998 | -3.13 | 1.13 |
| | 1000 | 736.48 | 2.128 | 0.145 | -2.630 | -6.83 | 1.57 |
| | Foraging distance | 737.52 | 0.036 | 0.850 | -0.149 | -1.69 | 1.40 |

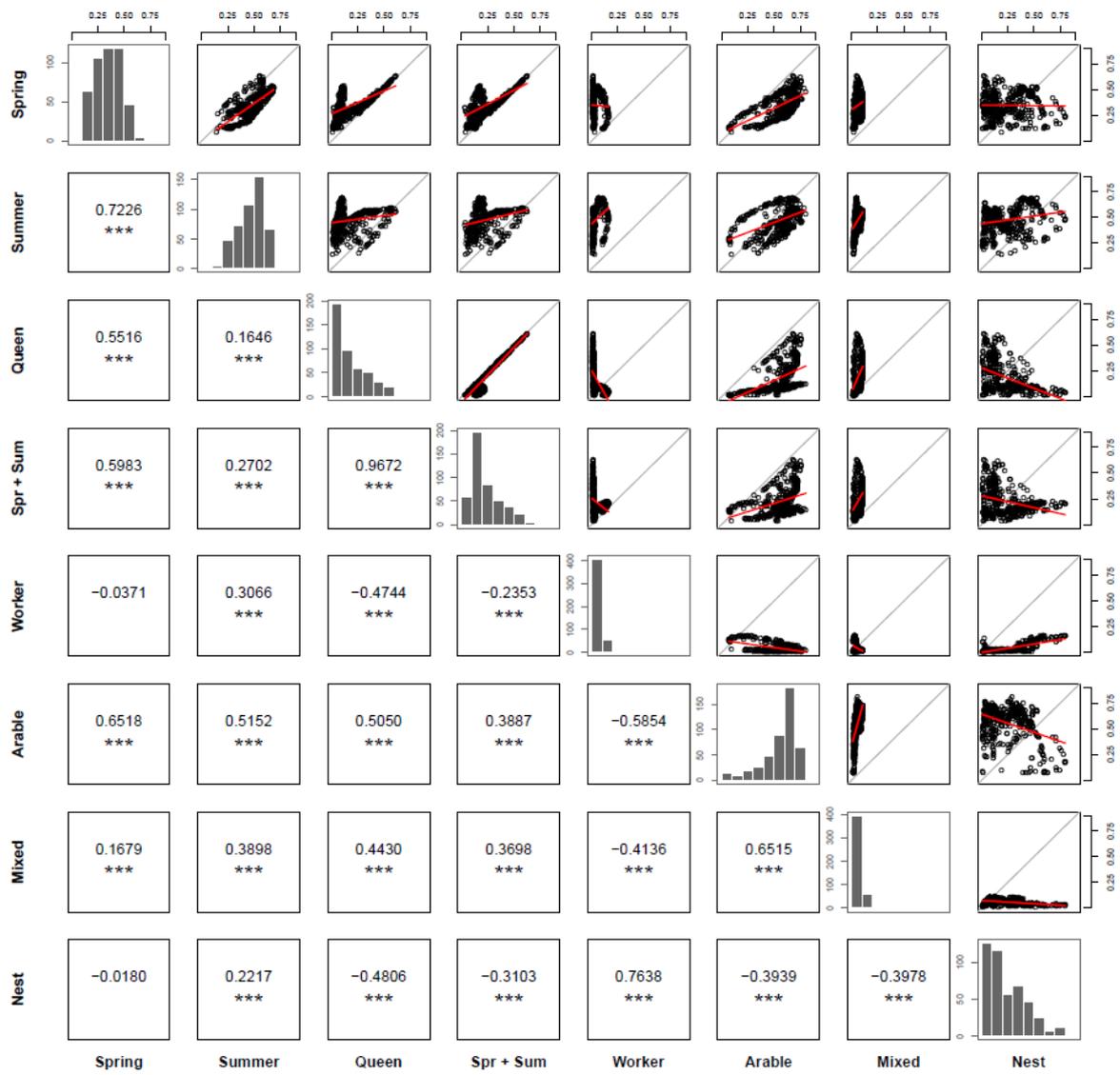
Extended Data Table 5

| Habitat / land-use variable | Description |
|---|--|
| (shortened name refers to Extended Data Figure 3 axis titles) | |
| All spring flower cover (Spring) | Flower cover of all surveyed plant groups in spring (April) |
| All summer flower cover (Summer) | Flower cover of all surveyed plant groups in summer (July and August) |
| Queen-visited spring flower cover (Queen) | Flower cover of plant groups visited by foraging queens during spring sampling (see Extended Data Table 6) |
| Spring + Summer flower cover (Spr + Sum) | Summed flower cover of plant groups visited by foraging queens during spring and preferred by foraging workers during summer (see Extended Data Table 6) |
| Worker-preferred summer flower cover (Worker) | Flower cover of plant groups visited by foraging workers during summer sampling and classified as 'preferred' based on number of visits by the three study bumblebee species (see Extended Data Table 6) |
| Arable field cover (Arable) | Fields growing arable crops including winter wheat (<i>Triticum aestivum</i>), oilseed rape (<i>Brassica napus</i>) and field bean (<i>Vicia faba</i>) |
| Mixed semi-natural vegetation cover (Mixed) | Mixed, non-woody, semi-natural vegetation including field margins sown with annual or perennial flower mixtures under the English agri-environment scheme, road verges and other non-woody linear habitats |
| Nesting habitat cover (Nest) | Cover of habitat classified as being of high nesting habitat suitability for at least one of the three study bumblebee species |

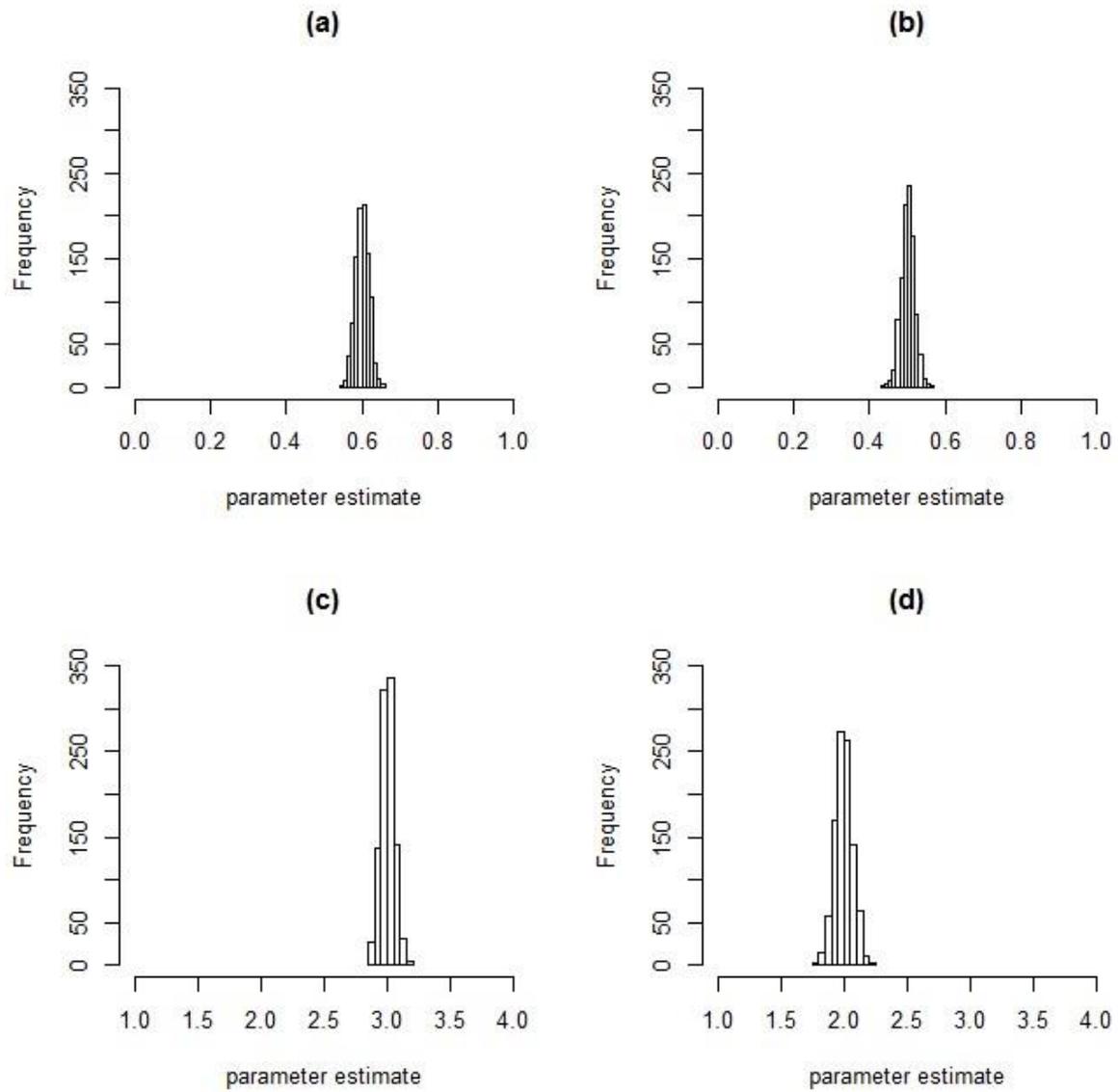
Extended Data Table 6

| Plant group | Example species | Spring queens – visited* | Workers - visited* | Workers – preferred# |
|-------------------------------|--|---------------------------------|---------------------------|-----------------------------|
| Lamiaceae, Scrophulariaceae | <i>Ajuga reptans</i> , <i>Ballota nigra</i> | Y | Y | N |
| Others, non-woody | <i>Apiaceae</i> , <i>Violaceae</i> | N | Y | N |
| Boraginaceae | <i>Borago officinalis</i> | N | Y | N |
| Oilseed rape | <i>Brassica napus</i> | Y | N | N |
| Other woody species | <i>Buddleja davidii</i> | N | Y | N |
| Convolvulaceae | <i>Calystegia sepium</i> | N | N | N |
| Knapweeds, Scabious, Teasels | <i>Centaurea spp.</i> , <i>Dipsacus fullonum</i> | N | Y | Y |
| Blue composites | <i>Cichorium intybus</i> | N | Y | N |
| Thistles | <i>Cirsium arvense</i> , <i>Carduus crispus</i> | N | Y | N |
| Hawthorn | <i>Crataegus monogyna</i> | N | N | N |
| Ericaceae, <i>Lavendula</i> | <i>Erica spp.</i> , <i>Lavendula spp.</i> | N | Y | N |
| White composites | <i>Leucanthemum vulgare</i> | N | Y | N |
| <i>Lotus spp.</i> | <i>Lotus corniculatus</i> | N | Y | Y |
| Other fruiting/flowering tree | <i>Malus spp.</i> | N | Y | N |
| Other clovers | <i>Melilotus officinalis</i> | N | Y | Y |
| Poppies | <i>Papaver spp.</i> | N | Y | N |
| Blackthorn, Cherry | <i>Prunus spinosa</i> | Y | N | N |
| Other Cruciferae | <i>Raphanus sativus</i> | N | Y | N |
| Rosaceae, non-woody | <i>Rosa spp.</i> | N | Y | N |
| Rosaceae, woody | <i>Rubus spp.</i> | N | Y | N |
| Willows, <i>Salix spp.</i> | <i>Salix caprea</i> | Y | N | N |
| Yellow composites | <i>Taraxacum agg.</i> , <i>Picris echioides</i> | Y | Y | N |
| Red clover | <i>Trifolium pratense</i> | N | Y | Y |
| White/Alsike clover | <i>Trifolium repens / hybridum</i> | N | Y | Y |
| Cereals | <i>Triticum aestivum</i> , <i>Zea mays</i> | N | N | N |
| Gorse | <i>Ulex europeaus</i> | N | N | N |
| Field bean | <i>Vicia faba</i> | N | N | N |
| Vetches | <i>Vicia spp.</i> , <i>Lathyrus spp.</i> | N | Y | N |

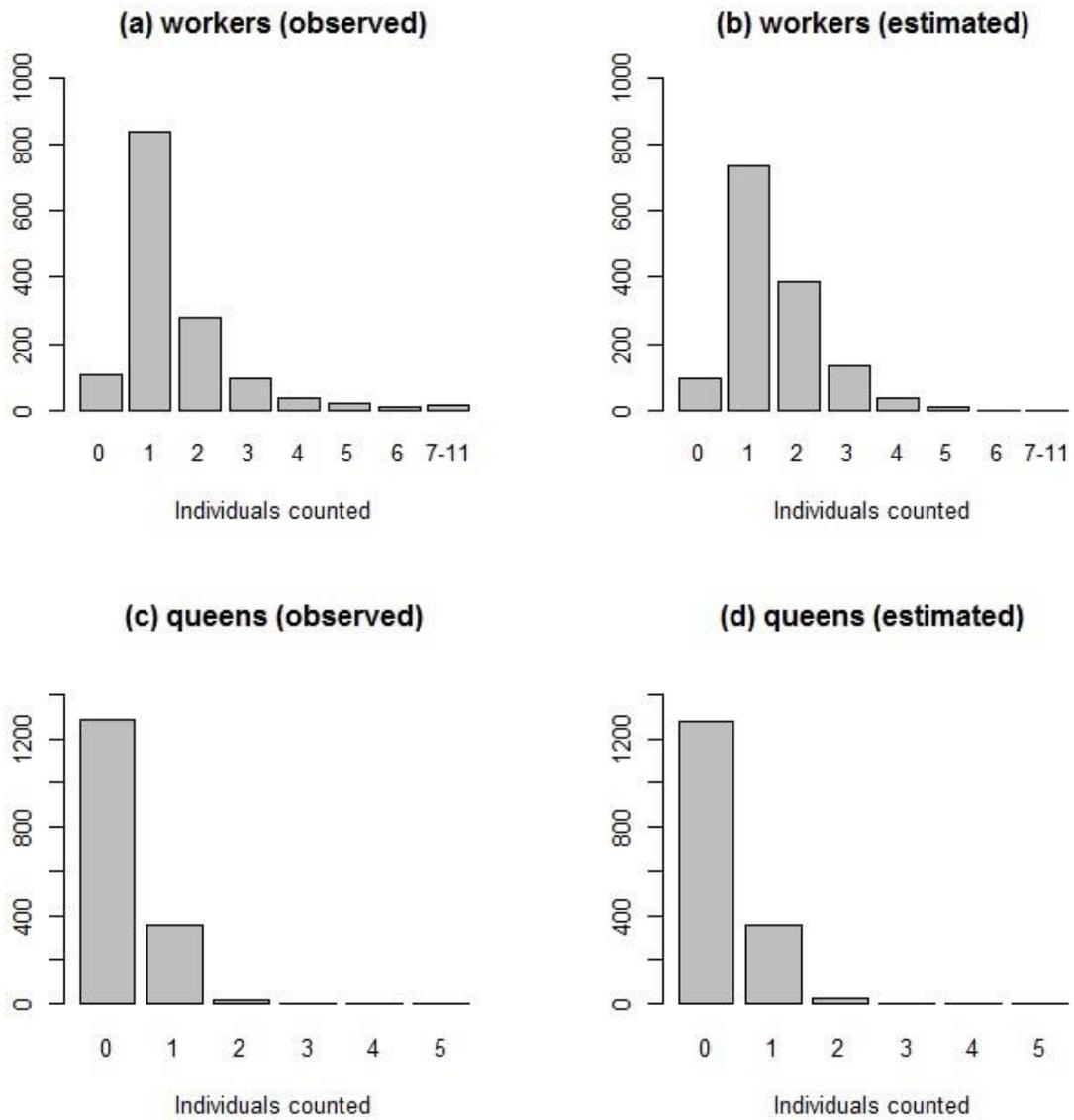
Extended Data Figure 1



Extended Data Figure 2



Extended Data Figure 3



Supplementary Information

A novel mark-recapture model for colonial species

Here we give a step-by-step account of model construction to explain how the parameters were derived, and follow with a description of its application to our dataset. Let the number of workers caught from colony i in 2011 be W_{1i} , and the numbers of queens caught be Q_{1i} in 2011 and Q_{2i} in 2012. As both the workers W_{1i} and queens Q_{2i} are the offspring of Q_{1i} , and existence of a colony can be ascertained without necessarily capturing the (single) founding queen, we have $Q_{1i} \in (0,1)$ but Q_{2i} and $W_{1i} \geq 0$, with $Q_{1i} + Q_{2i} + W_{1i} > 0$. Data from a single family lineage can thus be denoted by the vector $QWQ_i = (Q_{1i}, W_{1i}, Q_{2i})$. In formulating the likelihood, we define corresponding indicator variables $F_i = (F_{1i}, F_{2i}, F_{3i})$ where $F_{1i} = 1$ if and only if $Q_{1i} > 0$, and $F_{1i} = 0$ otherwise. F_{2i} and F_{3i} are defined similarly from W_{1i} and Q_{2i} . For colony i we further define the first and last encounters f_i and l_i as respectively the minimum and maximum values of j for which $F_{j,i} = 1$.

The data (F_{1i}, F_{2i}, F_{3i}) are of a form routinely analysed via mark-recapture models⁴¹ for individual animals, although here the ‘recaptures’ relate not to recognisable individuals, but to different individuals from the same bumblebee family lineage. In this respect, long-distance dispersal of individuals from a given colony simply lowers detection probability of the family, as opposed to confounding survival which would be the case with a standard individual CJS model. We define ϕ_1 as the probability that a founding queen produces at least one worker, and ϕ_2 as the probability that a colony producing workers (during 2011 in our dataset) subsequently produces new queens later in the season that survive hibernation to the following spring (2012 in our dataset). Completing the model, p_1 and p_2 are the probabilities that at least one worker, and at least one new queen, are captured within the study landscape given that the colony has survived to produce the caste in question. As we cannot differentiate between queens or

workers that have dispersed or emigrated from the study landscape and those that have died, φ_1 and φ_2 are termed ‘apparent survival’, because they represent the probability that the family lineage survives and remains available for capture within the study landscape.

The parameters φ and p form the standard Cormack-Jolly-Seber mark-recapture model and this is readily fitted to the data (F_{1i}, F_{2i}, F_{3i}) , for which the likelihood function LF for all n located colonies is given by:

$$(1) LF = \prod_{i=1}^n H_i \chi_{l_i}$$

Where, conditional upon the initial detection:

$$H_i = \prod_{j=f_i}^{l_i-1} \varphi_j p_j^{F_{j+1,i}} (1 - p_j)^{(1-F_{j+1,i})} \quad f_i < l_i$$

$$\chi_j = (1 - \varphi_j) + \varphi_j (1 - p_j) \chi_{j+1}, \quad j = 1, 2.$$

thus H_i is the probability of the detection history up to the final observation, with $H_i = 1$ for $f_i = l_i$ and χ_j is the probability of the colony subsequently remaining undetected, with $\chi_3 = 1$.

Because the parameters p_2 and φ_2 are confounded²⁵, only φ_1 and p_1 are uniquely estimable. It follows from the definitions that each observed history F can take one of only seven values: (1,1,1), (1,1,0), (1,0,1), (1,0,0), (0,1,1), (0,1,0) or (0,0,1). In contrast, encounters of multiple individuals from the same family lineage mean that possible values for a set of counts QWQ number many more, providing the means to model these counts and avoid the confounding present in the standard CJS model, in which the counts are not included. We therefore extended the model to accommodate the extra information contained in the full colony-specific counts (Q_{1i}, W_{1i}, Q_{2i}) . We assumed that W_{1i} and Q_{2i} , the counts of 2011 workers and 2012 queens from

existing families captured, were Poisson distributed with expected values λ_1 and λ_2 . For comparison with the CJS model above this gives:

$$(2) p_1 = 1 - e^{-\lambda_1} ; \quad p_2 = 1 - e^{-\lambda_2}$$

Formally, the joint probability of counts from a family $\Pr(QWQ) = \Pr(F) \times \Pr(QWQ/F)$ and we now have:

$$(3) \chi_j = (1 - \varphi_j) + \varphi_j(e^{-\lambda_j}) \chi_{j+1}, \quad j = 1, 2.$$

The full likelihood L_{QW} , again conditional upon first encounter, is presented in Supplementary Information Table 1. Now φ_2 and λ_2 are not confounded; all parameters are estimable. The means λ_t ($t = 1, 2$) are a measure of the combination of worker or queen abundance and the probability of individuals from a lineage being captured. A simulation-based assessment of model robustness in practice is given in Extended Data Figure 2. This shows that the estimated parameter values aggregate around the true values (i.e. those from which the data were simulated), demonstrating good performance of the model.

Using this basic model structure, it is possible to estimate parameters of interest either separately for each species, or to simplify the model by equating across species. Furthermore, it is possible to regress these within the likelihood on colony-specific covariates (X_i), which in this study were measures of habitat quality and land-use surrounding each estimated colony location (Extended Data Table 5). If colony locations of all founding queens were available, regressions with habitat variables could be carried out for both survival parameters (φ_1 and φ_2). However, these regressions were only carried out when colony locations could be reliably estimated; i.e. for colonies from which more than one worker was captured ($W_{1i} > 1$) and for the family lineage relationship between the summer worker and spring queen stage (φ_2). From this sample, survival between the spring queen and daughter worker stage (φ_1) could not be estimated since, by definition, all colonies considered survived to produce workers. This

requires a further minor modification to the model of SI Table 1, as the appropriate model for W_{1i} is now a truncated Poisson (such that the minimum value of W_{1i} is 2); that for Q_{2i} remains a standard Poisson. The model collapses to a product of a truncated and a zero-inflated Poisson, with the remaining parameters λ_1 , λ_2 and φ_2 all separately estimable. Model fitting requires a numerical optimisation of the likelihood LQW , which was carried out using the package ‘optim’ in R⁴². Logarithmic (λ) and logit (φ) link functions were used to keep estimated counts positive and probabilities within the range [0,1].

Code availability. The full R code for simulating a dataset and fitting the modified CJS model is available in Supplementary Information file: Carvell_Nature_Supplementary Information_Msimulation.R.

References

- 1 White, G. C. & Burnham, K. P. Program MARK: survival estimation from populations of marked animals. *Bird Study* 46, 120-139 (1999).
- 2 Lebreton, J. D., Burnham, K. P., Clobert, J. & Anderson, D. R. Modeling survival and testing biological hypotheses using marked animals: a unified approach with case studies. *Ecological Monographs* 62, 67-118 (1992).
- 3 *R: A language and environment for statistical computing*, <<http://www.R-project.org/>> (2015).

SI Table 1. Construction of the likelihood for the modified CJS model. Probabilities shown are probabilities associated with caste-specific colony counts, $\Pr(Q_{1i}, W_{1i}, Q_{2i})$, tabulated by first (f_i) and last (l_i) encounters. $Po(\cdot)$ indicates the Poisson probability mass function, and $Po'(\cdot)$ the zero-truncated form. The likelihood function LQW is then given by:

$$LQW = \prod_{i=1}^n \Pr(Q_{1i}, W_{1i}, Q_{2i})$$

Where χ_j is as defined in SI Equation (3). For the regression analyses, only data/colonies for which the first encounter $f_i = 1$ or 2 and $W_{1i} > 1$ can be used, and $Po(W_{1i}; \lambda_1)$ is truncated accordingly throughout.

| | <i>last encounter</i> | | | |
|------------------------|-----------------------|----------|--|--|
| | $l_i = 1$ | 2 | 3 | |
| <i>first encounter</i> | $f_i=1$ | χ_1 | $\varphi_1 Po(W_{1i}; \lambda_1) \chi_2$ | $\varphi_1 \varphi_2 Po(W_{1i}; \lambda_1) Po(Q_{2i}; \lambda_2) \chi_3$ |
| | 2 | | $Po'(W_{1i}; \lambda_1) \chi_2$ | $\varphi_2 Po'(W_{1i}; \lambda_1) Po(Q_{2i}; \lambda_2) \chi_3$ |
| | 3 | | | $Po'(Q_{2i}; \lambda_2) \chi_3$ |

SI Table 2. Sampling correlation matrix for (transformed) parameters of the modified CJS model for all species.

| | <i>Logit</i> (φ_1) | <i>Logit</i> (φ_2) | <i>Log</i> (λ_1) | <i>Log</i> (λ_2) |
|------------------------------|------------------------------|------------------------------|----------------------------|----------------------------|
| <i>Logit</i> (φ_1) | 1 | -0.01 | -0.12 | 0 |
| <i>Logit</i> (φ_2) | -0.01 | 1 | 0 | -0.88 |
| <i>Log</i> (λ_1) | -0.12 | 0 | 1 | 0 |
| <i>Log</i> (λ_2) | 0 | -0.88 | 0 | 1 |