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1 Niche models for British plants and lichens obtained using an

- 2 ensemble approach
- 3

4 <u>Summary</u>

5

6 Site-occupancy models that predict habitat suitability for plant species in relation to 7 measurable environmental factors can be useful for conservation planning. Such models can 8 be derived from large-scale presence-absence datasets on the basis of environmental 9 observations or, where only floristic data are available, using plant trait values averaged 10 across a plot. However, the estimated modelled relationship between species presence and 11 environmental variables depends on the type of statistical model adopted and hence can 12 introduce additional uncertainty. We used an ensemble-modelling approach to constrain 13 and quantify the uncertainty due to the choice of statistical model, applying generalised 14 linear models (GLM), generalised additive models (GAM), and multivariate adaptive 15 regression splines (MARS). Niche models were derived for over 1000 species of vascular 16 plants, bryophytes and lichens, representing a large proportion of the British flora and many 17 species occurring in continental Europe. Each model predicts habitat suitability for a species 18 in response to climate variables and trait-based scores (evaluated excluding the species 19 being modelled) for soil pH, fertility, wetness and canopy height. An R package containing 20 the fitted models for each species is presented which allows the user to predict the habitat 21 suitability of a given set of conditions for a particular species. Further functions within the 22 package are included so that these habitat suitability scores can be plotted in relation to 23 individual explanatory variables. A simple case study shows how the R package (MultiMOVE) 24 can be used to quickly and efficiently answer questions of scientific interests, specifically 25 whether climate change will counteract any benefits of sheep-grazing for a particular plant 26 community. The package itself is freely available via <u>http://doi.org/10.5285/94ae1a5a-2a28-</u> 27 4315-8d4b-35ae964fc3b9. 28

29 Keywords: biodiversity; climate change; envelope; niche occupancy; pollution; R package

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- 31

32 <u>1. Introduction</u>

33

34 Biodiversity loss is a pressing global concern, and can be seen as largely driven by declines in

- 35 habitat suitability and availability for individual species (Hirzel and Le Lay, 2008). In seeking
- 36 to understand habitat suitability for any given species, it is important that two key
- 37 components are considered: the spatial distribution of the species (e.g. as in Groom, 2013;
- 38 Preston et al., 2013); and the relationship between the species and other influencing factors
- 39 (e.g. Geddes and Miller, 2012). To aid decision making in conservation practice, the current
- 40 distribution of a species and likely changes to this under different management, pollutant or
- 41 climate pressures should be understood. It is therefore useful to be able to define the
- 42 ecological range of a given species in relation to different environmental factors, i.e. its
- 43 realised niche, to enable efficient and timely decision making.
- 44

45 Niche models can be developed from species records when these are sufficiently numerous,

- 46 by relating presence or absence to environmental data where these are recorded alongside
- 47 occurrence, and/or to proxy variables based on the traits of co-occurring species (Latour,
- 48 1993; Smart et al., 2010b). Niche modelling approaches based on correlative analyses
- 49 contrast with population dynamics models, which require detailed parameterisation to
- 50 represent processes of reproduction and mortality, emigration and immigration (Crone et al
- 51 2011). While correlative-based niche models cannot simulate impacts of dynamic population
- 52 processes, they provide useful indications of how the availability of favourable niche space
- 53 will change (Dormann et al., 2012; de Vries et al., 2013, Thuiller et al., 2008). Ideally, any
- 54 model attempting to describe the spatial distribution of a species should be based on a large
- 55 number of observation records representing equilibrium conditions, to ensure that the
- 56 predictions are robust (Elith et al 2010).
- 57

58 Many taxa have been the focus of species niche modelling (Elith & Leathwick 2009). It is

59 particularly useful to predict habitat suitability for plant species, since they deliver

- 60 supporting ecosystem services such as primary production, nectar provision for pollinating
- 61 insects, genetic variation for crop breeding and cultural significance for wildlife conservation
- 62 (Alexander et al., 1997; Costanza et al., 2007; Kremen et al., 2007; UK National Ecosystem

63 Assessment, 2011). Plants also underpin the diversity of other taxa by providing habitat

64 structure and a diverse range of food substrates. Here we present an ensemble of empirical

- niche models for a large number of higher and lower plants in the British flora comprising all
- 66 major community dominants and a range of subordinates.
- 67

68 Previous work developed niche models for a similar group of plant species based on multiple

- 69 logistic regression (Smart et al 2010b); hereafter referred to in the more generic framework
- of Generalised Linear Models, GLMs. In recent years the diversity of techniques applied to
- niche modelling has expanded due, in large part, to the need to overcome issues related to
- 72 model constraints and interpretability (Elith & Leathwick 2009). However, different

- 73 modelling approaches can result in different representations of the variation in the observed
- data, leading to differences in model transferability (Munoz and Felicisimo, 2004; Leathwick
 et al., 2006; Smart et al 2010a; Wenger & Olden 2012). This has led to the increase in
- et al., 2006; Smart et al 2010a; Wenger & Olden 2012). This has led to the increase in
 popularity of an ensemble approach (eg Araújo and New, 2007; Thuiller et al., 2009) where
- 77 the explainable variation and uncertainty relating specifically to model selection are more
- 78 robustly conveyed based on output across different model types. The work presented here is
- 79 based on the addition of a further two techniques to the GLM approach applied in Smart et
- 80 al. 2010b: generalised additive modelling (GAM); and multivariate adaptive regression
- 81 splines (MARS). The models produced allow spatial and temporal prediction of change in the
- 82 favourability of niche space for each species based on outputs from the three modelling
- 83 techniques, conditional upon measured or predicted environmental conditions.
- 84
- 85 The models themselves have been bundled into a publicly available R package to allow the
- 86 wider community of scientists, land managers and conservation policy makers to query,
- 87 scrutinise and exploit the fitted models for scientific and decision making purposes. The
- 88 package facilitates understanding and explanation of species' distributions by allowing clear
- 89 inspection of species responses along environmental gradients. By applying projected
- 90 changes in input variables, the user can also explore future scenarios of environmental
- 91 change (eg. Smart et al 2010a). The aim of this paper is to present a brief overview of the
- 92 fitted models before introducing the R package containing all the model fits and a clear
- example of how this can be used to provide speedy and efficient answers to policy-relevantquestions.
- 95

96 2 Methods and Materials

- 97 <u>2.1 Data</u>
- 98
- 99 Fine-grained data on the presence/absence of plant species were available at a large number
- 100 of locations throughout the UK from four studies: the Countryside Survey (CS) (Smart et al.
- 101 2003), GB Woodland Survey (Kirby et al., 2005; Corney et al., 2006), the surveys that
- 102 provided data for Key Habitat Types (Hornung, 1996) and the National Vegetation
- 103 Classification (NVC) (Rodwell, 1991 et seq.). We pooled the data from the four different
- 104 surveys, giving a total of 32272 vegetation plots. The NVC surveys represent the largest
- 105 source of species data and were designed to sample from the full range of UK plant
- 106 assemblages, so they include more records for scarce species than would a random survey.
- 107 The NVC design may therefore have resulted in over-sampling under optimal conditions, so
- 108 data from surveys with a stratified randomized design (or which at least included an element
- 109 of random plot location) were also included in order to try and provide an unbiased
- 110 representation of the entire ecological range of a species, importantly including the tails of
- 111 the distribution moving away from the optimum. Information on plot size was unavailable
- 112 for NVC quadrats and therefore no standardisation across plot sizes was possible. To
- 113 overcome this models were, for species with sufficient data (typically n=30 records), re-fitted

114 using CS data only (for which the plots are all of a standard size and the location follows a 115 stratified, random design).

116

117 The plant species modelled exclude the rarest species in our flora and mainly comprise 118 habitat dominants and a large range of subordinates (sensu Grime 1998). Dominant species 119 are responsible for the majority of the primary production at a site and strongly underpin 120 other ecosystem functions (Smith & Knapp 2003; Laughlin 2011). Thus the species modelled 121 comprise a disproportionately large fraction of the biomass and cover in British habitats. The 122 list of species modelled includes 97% of the Common Standards Monitoring indicator 123 species (JNCC, 2004) used to judge conservation value of semi-natural habitats. The 124 suitability of conditions for rare species' populations often depends upon the varying cover 125 and persistence of the more abundant species, so even where not directly modelled, 126 information about the prospects for rare species can be inferred from modelling the 127 responses of their more common associates (Smart 2000, Gogol-Prokurat 2011). Non-native 128 species were also excluded from analysis as they are more likely to be undergoing increases 129 in range and colonisation of suitable niche space. As such species are not in equilibrium, the 130 estimated environmental effects from spatially derived models may be confounded with the 131 effects of incomplete dispersal (Svenning & Skov, 2004).

132

133 In choosing environmental characteristics to define the niche, we selected a set of variables

134 representing abiotic and climatic influences. Climate variables included in the models,

135 chosen due to their relationship with plant physiology and growth (Thuiller et al., 2005),

136 were long-term (1961-1990) annual average: rainfall; minimum January temperature; and

137 maximum July temperature (all of which are available from

138 http://www.metoffice.gov.uk/climatechange/science/monitoring/ukcp09/). Abiotic variables

139 selected were based on mean values for trait-based indicators. These were: mean un-

140 weighted Ellenberg scores for soil wetness (F), substrate fertility (N) and soil pH (R) for each

141 vegetation plot. Ellenberg scores are ordinal values that are assigned to each plant species.

142 They were originally designed to reflect habitat preferences in central Europe (Ellenberg *et*

143 *al.* 1991), but have subsequently been adapted for use with British higher and lower plant

species (Hill *et al.*, 1999, 2004, 2007). Mean Ellenberg scores provide a quantitative proxy for

145 the abiotic characteristics of a particular plot (Diekmann, 2003). Similarly, we used species-

146 specific scores for typical canopy height following the ordinal categorisation of Grime et al.

147 (1988) to derive a mean cover-weighted canopy height trait score. This provides a measure

148 of the successional status of the vegetation, which also reflects management intensity or the

- 149 frequency and severity of biomass removal.
- 150

151 Using trait scores allowed for models to be constructed with floristic and climate data only,

- 152 without the additional need for coincident soil measurements which are rarely available.
- 153 Consistent with Smart et al (2010b), if the species being modelled was recorded in a plot its
- 154 trait values were removed prior to calculating the trait-based explanatory variables for the

155 plot, to avoid circularity in the model. Where measurements or model estimates of soil

- 156 variables are available, translation functions can be used to predict the mean Ellenberg
- 157 scores required to solve the niche models alongside climate and cover-weighted canopy
- height (Rowe et al., 2011a; Smart et al., 2010b; Rowe et al., 2014b). Mean Ellenberg scores
- 159 were not weighted by observed species cover, since little information is gained when doing
- 160 so relative to the noise in the cover estimates (Valentin et al 2012; Kafer & Witte 2004).
- 161 Canopy height values were cover-weighted, however, to convey the influence of canopy
- 162 extent as well as height on shade at ground level.
- 163

164 <u>2.2 Models</u>

165

166 All surveys recorded the presence or absence of individual species within individual plots, 167 and models with a binomial distribution for the response variable were fitted to this data 168 using three modelling techniques – GLMs (McCullagh et al., 1989), GAMs (Hastie and 169 Tibshirani, 1990) and MARS (Friedman, 1991). Previously, GLMs were fitted to the same 170 vegetation data by Smart et al (2010b) and are useful for modelling simple linear or 171 polynomial responses. However, GLMs cannot fit more complex surfaces which may 172 characterise species' niches. GAMs are a much more flexible class of models, allowing the 173 relationship between the response and any individual predictor to have a smoothly varying 174 form. However, estimating such complex relationships is more difficult, so if the response is 175 in fact simple the models may have predictive power for new data (Smart et al., 2010a). The 176 MARS models are similar to GAMs but instead of fitting smoothed terms they fit "hinges" 177 (Friedman, 1991). They are similarly more flexible than GLMs, but their use of piecewise 178 linearity accommodates different types of responses to GAMs. An ensemble of all three 179 models was fitted for each species.

180

181 For each species the full model contained all seven covariates (long term averages for 182 maximum July temperature, minimum January temperature and annual rainfall, canopy 183 height and Ellenberg F, N and R scores) and all of the 21 possible two-way interactions. This 184 is in contrast to Thuiller et al. (2003) who considered only additive effects without 185 interaction. Correlation across all combinations of variables was assessed and evidence of 186 some relationships was found: out of the 21 pairwise correlations, 5 were either greater 187 than 0.5 or less than -0.5 and the largest correlation (between Ellenberg N and Ellenberg R) 188 was 0.75. However, due to the modelling techniques chosen and the implementation of a 189 suitable model selection routine for each species modelled individually, this was not 190 considered a problem. For the GLM models stepwise backwards selection based on 191 minimum AIC was used to define the final set of covariates and two-way interactions to be 192 considered for each species. Main effects were only considered for removal if all interactions 193 containing these variables had already been removed. The variables and interactions that were used within the GAM models were those chosen in the final selected GLM models, 194 195 because it was found to be computationally infeasible to perform a separate model selection

- 196 procedure for the GAMS. Within the MARS framework model selection is performed
- 197 automatically as part of the model fitting process using forward selection, i.e. starting from a
- 198 null model and adding in those terms that lead to the greatest improvement in fit.
- 199
- 200 All models were fitted in the R statistical environment (R Development Core Team, 2009)
- 201 using the mgcv (Wood, 2006), earth (Milborrow, 2014) and leaps (Lumley, 2009) packages.
- 202 Nomenclature follows Stace (2010) for vascular plants, the Atlas of British and Irish
- 203 Bryophytes (Blockeel et al. 2014) and the British Lichen Society Taxon Dictionary (2015;
- 204 http://www.britishlichensociety.org.uk/resources/lichen-taxon-database).
- 205

206 2.3 Model Checking

207

208 Model performance was assessed by comparing the observed presence absence data (0 or 209 1) to predicted values using AUC (Fielding and Bell, 1997), where predicted values in this 210 case were defined as the estimated probability of presence for each of the three model 211 approaches (GLM, GAM, MARS). Values of AUC close to one indicate good levels of 212 predictive performance whereas a model with predictive power that is no better than 213 chance will return an AUC of 0.5. The AUC values should be interpreted with caution 214 because they effectively treat the cost associated with a false positive as being identical to 215 the cost associated with a false negative, and this may not always be appropriate. Other 216 measures are possible, though the low frequency of presences for the majority of species at 217 observed sites makes measures like false omission rate and negative predictive value 218 unhelpful.

219

220 Since an adequate independent test data set was not available, a cross-validation approach 221 is required to investigate the out-of-sample performance of the fitted models. However, 222 computational costs for checking all species models would be prohibitive, so the full set of 223 models was compared and examined using within-sample AUC diagnostics as described 224 above. Out of sample performance was assessed on a much smaller set of 30 species, listed 225 in Table 1. This set was designed to represent a range of distributions and taxa, and explicitly 226 included a number of species which show evidence of a non-linear relationship between 227 abiotic conditions and prevalence. For each of the 30 selected species, the observed data 228 were subset at random into two components: 75% of the data for training, 25% for testing. 229 Each of the three models were then built on the 75% dataset and AUC values were 230 calculated based on the remaining, independent 25% of the data. We repeated this process 231 10 times for each species.

232

233 The predicted values across the range of the training data were also mapped across GB in

- 234 order to provide an indication of the spatial extent of each species according to the models.
- 235 This enabled us to check against expert knowledge and previously produced maps, such as

- those in Preston et al., 2013 and those readily available on the BSBI and BRC websites, how
- 237 well the fitted models did in characterising the range and extent of species occurrence.

238 <u>3. Results</u>

239

240 Within-sample AUC values for each of the fitted models across the three methods built using

all data and only CS data respectively are shown as histograms in Figure 1. For the majority

- of species the overall performance of all methods in fitting the observed data appears to be
- very good with AUC values in excess of 0.8. AUC values for models built using CS data only
- were generally slightly lower (Figure 1, Table 1), but still showed good performance with a
- high percentage of AUC value in excess of 0.8 across all models (Table 1). The lower AUC
- values resulting from the CS models is likely to be due to the smaller proportion of absences
- 247 in the CS dataset than in the NVC dataset.

248



Figure 1: Histograms of AUC values for GLM (a, b), MARS (c, d) and GAM (e, f) models built using all survey data
 (a, c, e) and CS data only (b, d, f). AUC values > 0.8 coloured green; AUC values > 0.5 and < 0.8 coloured blue;
 AUC values < 0.5 coloured red.

253

254 Differences between methods were relatively small for many species, but there were 255 examples where the MARS and GAMS approaches clearly provided a substantially better fit 256 to the data (at least in terms of AUC) than the GLM approach. Across the full set of species 257 modelled using all data, only 1% had the highest AUC value for the GLM models compared 258 with 17% for the MARS and 82% for the GAMs (Table 1). For the models built using CS data 259 these percentages equate to 12%, 14 % and 73% respectively, though far fewer species 260 models were fitted using the CS data only due to the availability of sufficient data records 261 (Table 1).

Table 1: Number of species models fitted using each method and data source together with a summary of AUC
 statistics across each model type showing the proportion of cases where each model type was the "best"
 (highest AUC)

		All Veg data	a	CS only			
	GLM	MARS	GAM	GLM	MARS	GAM	
Number of Species Models	1017	1178	1017	387	388	387	
% Maximum AUC	0.85	17.23	81.92	12.37	14.18	73.45	
% AUC values > 0.8	99.31	94.74	99.80	80.62	72.42	91.47	
% AUC values > 0.9	85.84	73.43	95.87	37.47	31.44	57.62	

266

267

268 Spatial predictions assessed against the corresponding maps published in Preston et al. 2013

and the individual species distribution maps on the BSBI website, showed that the models

270 broadly captured the spatial drivers dictating the distribution of individual species.

271 Altitudinal, coastal and latitudinal preferences were clearly captured adequately by the

272 niche models. Some examples of predicted distributions for four species are shown in Figure

273 2. *Clematis vitalba* is shown to be a species of calcareous soils, with maximum habitat-

suitability on the chalk hills of southern England. Epilobium montanum is a species with a

wide range in terms of altitude, precipitation and soil. Lemna trisulca is shown to be a

276 species of wet lowland sites, and *Selaginella selaginoides* is restricted to wet mountains.



Figure 2: Probability distribution maps produced using the average fitted probability across the three modelling approaches based on preciting the full GB extend of input data (i.e the environmental values from the training datasets) for 4 species: a) *Clematis vitalba*; b) *Epilobium montanum*; c) *Lemna trisulca*; d) *Selaginella selaginoides*. White squares represent areas missing from the training data set and hence no prediction could be made there.

 285 AUC values obtained from the cross validation exercise showed that within-sample

286 predictive performance was matched by out-of-sample performance (Table 2). In only two

287 cases out of a total of 90 (3 models for each of 30 species) did the AUC value drop from

above 0.8 using within-sample prediction to less than 0.8 using the cross-validation. These

were both MARS models for *Senecio erucifolius* and *Teesdalia nudicaulis*. Given the range of

- 290 species tested, both rare and abundant species, we conclude that our models provide a
- useful way of assessing the favourability of species to environmental conditions.
- 292

295

293 **Table 2:** Model fit diagnostics (in the form of AUC statistics) for all three modelling approaches and an estimate

based on the average of the fitted values across these models from a subset of models for 30 test species.

	Within Sample AUC						Out of sample AUC		
Species Name	Models built using all survey data			Models built using CS data only			Cross Validation		
	MARS	GAM	GLM	MARS	GAM	GLM	MARS	GAM	GLM
Agrostis capillaris	0.881	NA	NA	0.858	0.818	0.707	0.872	0.852	0.749
Campylopus flexuosus	0.894	0.911	0.905	NA	NA	NA	0.891	0.906	0.904
Campylopus introflexus	0.867	0.965	0.948	NA	NA	NA	0.831	0.952	0.937
Campylopus pyriformis	0.881	0.918	0.903	NA	NA	NA	0.896	0.922	0.91
Carex caryophyllea	0.949	NA	NA	0.915	0.948	0.951	0.95	0.963	0.943
Carex limosa	0.995	0.997	0.991	NA	NA	NA	0.924	0.98	0.99
Carex viridula subsp. oedocarpa	0.947	NA	NA	0.916	0.944	0.893	0.935	0.949	0.923
Cochlearia pyrenaica	0.811	0.989	0.978	NA	NA	NA	0.9	0.978	0.972
Conopodium majus	0.914	NA	NA	0.808	0.862	0.787	0.91	0.911	0.85
Cynosurus cristatus	0.899	NA	NA	0.87	0.882	0.865	0.897	0.932	0.894
Dryas octopetala	0.995	0.999	0.992	NA	NA	NA	0.969	0.99	0.991
Helianthemum nummularium	0.964	0.977	0.97	0.893	0.901	0.947	0.964	0.976	0.963
Juncus articulatus	0.928	0.955	0.911	0.893	0.773	0.734	0.912	0.936	0.898
Lolium perenne	0.951	NA	NA	0.942	0.947	0.91	0.939	0.944	0.932
Neottia ovata	0.852	0.937	0.871	NA	NA	NA	0.875	0.937	0.872
Plantago major	0.892	NA	NA	0.861	0.869	0.822	0.872	0.879	0.859
Plantago media	0.954	0.973	0.963	0.827	0.944	0.948	0.944	0.971	0.959
Polytrichastrum alpinum	0.96	0.979	0.974	NA	NA	NA	0.957	0.976	0.974
Ranunculus repens	0.889	NA	NA	0.847	0.879	0.801	0.873	0.899	0.851
Sanguisorba officinalis	0.933	0.974	0.901	NA	NA	NA	0.925	0.942	0.9
Senecio erucifolius	0.816	0.946	0.897	0.838	0.941	0.909	0.765	0.924	0.892
Sphagnum cuspidatum	0.941	0.976	0.972	NA	NA	NA	0.938	0.971	0.969
Sphagnum denticulatum s.l.	0.944	0.95	0.93	NA	NA	NA	0.833	0.917	0.926
Sphagnum fimbriatum	0.95	0.969	0.919	NA	NA	NA	0.919	0.956	0.914
Sphagnum squarrosum	0.967	0.934	0.901	NA	NA	NA	0.957	0.916	0.894
Teesdalia nudicaulis	0.968	0.993	0.982	NA	NA	NA	0.458	0.932	0.971
Trifolium repens	0.909	NA	NA	0.894	0.906	0.861	0.901	0.917	0.876
Vaccinium myrtillus	0.956	NA	NA	0.944	0.95	0.941	0.952	0.962	0.951
Viola hirta	0.922	NA	NA	0.377	0.936	0.951	0.925	0.955	0.937
Viola riviniana	0.893	0.925	0.875	NA	NA	NA	0.901	0.934	0.878

296

297 <u>4. Model exposition and example application</u>

- 298 Any user, applied scientist or policy maker may wish to explore each of the fitted models in
- detail. For example, the models may be used to investigate relationships and make
- 300 predictions under different scenarios. However, there are too many individual models to
- 301 describe in sufficient detail. In order to provide full functionally and interrogation of all
- 302 models, they were packaged into a user friendly R library created specifically for this
- 303 purpose MultiMOVE. As well as containing the actual model fits themselves, the
- 304 MultiMOVE package created by the authors consists of four key functions to access and
- 305 query the fitted niche models. These functions enable predictions to be made, covariate
- 306 relationship to be plotted, raw probabilities to be converted to a rescaled habitat suitability
- 307 score and the final model formula to be exposed. Here we describe this R package and308 provide an example of its use in a scenario exploration for a single species.
- 309 As an example of using the fitted models and the MultiMOVE package to explore
- 310 relationships and answer scientific questions of interest relating to a species' niche, we
- 311 consider the question posed by Geddes and Miller (2012): will climate change counteract
- 312 the benefits of sheep-grazing in conserving a rare alpine dwarf-herb community? We focus
- 313 on *Festuca ovina agg.,* one of the key species in the study.
- 314

315 **4.1 Making predictions**

316

The prediction function allows the user to access the fitted ensemble of niche models and to make predictions for a given species with specified environmental data located within the modelled covariate space. The *MM_pred* function takes a data frame of the environmental covariates and returns an aligned data frame with model predictions of species occurrence probabilities. Upper and lower confidence limits for the predictions are also returned, though as the models did not account for any residual spatial autocorrelation, extreme caution is advised when using these.

324

Running the *MM_pred* function on *Festuca ovina agg.* (BRC number 920821) across the whole of the training data set and mapping the estimated probabilities allows visualisation of the predicted species range. This is useful not only for prediction, but also as a check to see if the models have adequately captured the distributional range and preferences for a given species before further investigation or use of the model. The predicted distribution map for *Festuca ovina agg.* (Figure 3) shows good agreement with the online atlas map.



332	
333 334	Figure 2: Producted distribution map of <i>Fastures oving and</i> using fitted probabilities from the MMA prod
334	command on the full training data set. Colour scale: Vellow = 0.1 corobability occurrence < 0.25: Orange =
336	0.25 < probability occurrence < 0.5: Red= 0.5 < probability occurrence < 1. White areas represent either returned
337	probabilities of 0 or areas entirely absent from the training data.
338	
339	4.2 Visualising covariate effects
340	
341	The MM_plot command feature enables the user to observe the marginal or joint effects of
342	covariates on the species response. When viewing the marginal or joint effects of specified
343	covariates all other covariates are held at their median values. Either one or two covariates
344	can be specified at a time to visualise their effect on species' occurrence probabilities.
345	
346	Geddes and Miller, 2012 were specifically interested in the effects of grazing and climate. In
347	the MultiMOVE models these variables are represented by cover-weighted canopy height
348	and maximum July temperature respectively. Therefore, to answer the question as to
349	whether climate change would counteract any advantages grazing would have, we can use
350	the MultiMOVE package to visualise the joint effects of canopy height and July temperature
351	and draw our inference from that. Running the MultiMOVE commands in R as follows,
352	
353	> MM_plot(input_data, BRC=920821, view_term=c("cov4"), display = "raw") ; and
354	> MM_plot(input_data, BRC=920821, view_term=c("mju","cov4"), display = "raw"),
355	
356	returns plots of the fitted marginal effects of canopy height (our proxy for grazing) and a plot
357	of the fitted joint effects between canopy height and July temperature for Festuca ovina
358	agg.
359	



Figure 4: Output from the *MM_plot* command run on *Festuca ovina agg*. The figure on the left shows the
 marginal effect of cover weighted canopy height on the occurrence of *Festuca ovina agg*. with each of the
 model predictions plotted. The plot on the right shows the joint effect of canopy height and maximum July
 temperature plotted using the average fitted values across the models.

- Fitted models within MultiMOVE show that *Festuca ovina agg.* favours lower canopy heights
 and hence that grazing has positive benefits for this species (Figure 4). The interaction
- 370 between canopy height and temperature indicates that when canopy height is low, if
- 371 maximum July temperature rises to above 19°C, the likelihood of occurrence decreases. Our
- 372 conclusion therefore would be that there is evidence to suggest that the potential effects of
- 373 climate change (increasing July temperature) could potentially counteract the benefits of
- grazing for *Festuca ovina agg.* This is in keeping with the results from the Geddes and Miller,2012 study.
- 376

4.3 Converting raw probabilities into habitat suitability scores

- 378
- MultiMOVE also contains a function (*HS_convert*) to transform the fitted probabilities, the values of which are dependent on the tolerance or niche breadth of the species and its prevalence, into a habitat suitability score on a standardised scale which is then directly comparable between species. There are two options for doing this within the function: based on the cumulative distribution of fitted probabilities to the training data; and based on the prevalence of the species within the training data.
- In the first case, the cumulative distribution of probabilities fitted to the training data set isobtained for all model types across all species. Rescaled probabilities are then estimated to

be the percentile of the cumulative distribution corresponding to the un-scaled raw
 occurrence probability. The rescaled "habitat suitability" scores therefore provide an
 indication of how extreme the raw occurrence probability is across the full sample in the

- 391 training data i.e. does the un-scaled occurrence probability correspond the edges of a392 species' range.
- 393

394 Using the species prevalence to re-scale the fitted occurrence probabilities follows the 395 approach and formula suggested by Real et al., 2006. The rescaling in this case ensures that 396 a habitat suitability score of 0.5 corresponds to what would be obtained by taking an 397 average of the prevalence across the whole training data set and hence corresponds to 398 equal favourability everywhere.

399

400 **4.4 Obtaining the fitted model**

401

The final function contained within the MultiMOVE package, *extract_MM_model*, enables the user to extract the fitted model for a given species, in particular to see which covariates

- 404 were included in the final model after model selection.
- 406 **<u>5. Discussion</u>**
- 407

405

408 Species niche models encapsulate the relationships that drive the spatial distribution of

409 individual species and have many potential uses in assessing the effects of environmental

- 410 change on habitat suitability for individual species, the consequences this has for
- 411 biodiversity value and the provision of ecosystem services. However, despite considerable
- 412 efforts in constructing such distribution models, the application of such models in
- 413 developing policy and management recommendations has been patchy (Guisan et al., 2013).
- 414 This can often be due to the complex nature and time involved in re-running models under
- 415 different scenarios. The large number of models developed and user friendly R package
- 416 described here make it highly relevant for assessing the likely impact of management and 417 policy decisions and for facilitating efficient and responsive outcomes for conservation
- 418 management.
- 419

420 The assessment of the model fits showed good performance based on within sample AUC

- 421 across all species and out of sample AUC for a subset of test species. AUC statistics
- 422 suggested that use of an ensemble approach was beneficial as no one model type was
- 423 optimal across all species. Greater model flexibility requires more data to estimate
- 424 accurately, so there can be cases with limited data where the more free-form models
- 425 underperform while a linear (or log-linear) model can adequately capture a pattern of
- 426 responses (as shown in Smart et al., 2010a). This suggests that there are benefits in using
- 427 multiple modelling techniques (Araújo & New, 2007).

428 It is important to note that potential residual spatial autocorrelation was not accounted for 429 within any of our models. This should not introduce systematic bias into our parameter 430 estimates (and, therefore, predictions), but it is likely to mean that we substantially 431 underestimate the uncertainty associated with the outputs from our models. This can lead 432 to overfitting, especially with the more complex models, compounding the need for caution. 433 On our small subset of species, we investigated any potential evidence of this by using the 434 pairwise distance sampling technique of Hijmans (2012) to reduce the impact of spatial 435 effects on cross-validated AUC measures. The results suggest that our methods still perform 436 guite well despite the above limitation. It is possible that the methods considered here could 437 be extended to explicitly account for residual spatial autocorrelation and recent software 438 developments (such as the R-INLA package; http://www.r-inla.org/) mean that it is now 439 computationally feasible to apply such extensions to relatively large datasets. However, the 440 large number of species across which each model is run means that the inclusion of these 441 additional models will inevitably require a substantial amount of computational effort. Due 442 to the complex nature of spatial statistical models and the large number of sites and species 443 in our data set, fitting and optimising simpler models, as done here, is a sensible first step.

444

445 The MultiMOVE package exploits the extensive datasets describing the occurrence of a large 446 number of vascular plant and bryophyte species across GB to build a picture of how 447 individual species favour different environments. Some care should be taken in interpreting 448 the outputs, since the fitted models will reflect any bias in the training data. A key advantage 449 of the current work has been the use of standardized guadrat data with which to train the 450 models, primarily from the Countryside Survey and the NVC, rather than opportunistic 451 presence-only data. This should greatly reduce the influence of gross spatial heterogeneities 452 in recording effort that often arise in opportunistic datasets. As always, the development of 453 niche models depends heavily on reliable data on which to train the models. For niche 454 models and their outputs to be representative of conditions and niche preferences across 455 Britain, the training data itself should be unbiased and representative of this same 456 population. As such, the fitted models do not apply to environments not included in the 457 training data - those outside of GB for example - and MultiMOVE accounts for this by 458 warning when predictions are attempted using input data outside the covariate space 459 observed in the training data.

460

461 The MultiMOVE package allows exploration of the impact of environmental change on plant 462 species and assemblages across terrestrial ecosystems in Britain, by making use of the large 463 datasets that are available and an ensemble approach to modelling niches. Site managers 464 can quickly assess current suitability for individual species, using freely available climate data 465 (http://www.metoffice.gov.uk/climatechange/science/monitoring/ukcp09/) and a list of 466 currently-present species. This makes use of the fact that the species that are present give a 467 considerable amount of information about site conditions. The package also allows the 468 effects of environmental change to be expressed in terms of likely changes in species

- 469 composition and derived biodiversity indicators (Rowe et al., 2014a). Observed
- 470 environmental changes can be interpreted in terms of how they are likely to have affected
- habitat suitability for species. Cautious ecological judgement is needed. The models are built
- on spatial patterns. When used to project change through time the assumption of space for
- time substitution is strong and likely to be highly questionable especially when confronting
- past patterns with potentially novel environments (Williams & Jackson 2007). In addition the
 uncertainty around the relationship between abiotic conditions and the Ellenberg values,
- 476 whose means are used to convey these conditions, is propagated through the application of
- 477 MultiMOVE. The impact of this uncertainty is likely to be especially critical when scenario
- 478 testing since any lagged changes are not simulated. Changes in habitat suitability can
- 479 however be driven by the outputs of dynamic models of climate or biogeochemical
- 480 processes (Rowe et al. 2011b). The opportunity to explore the ecological responses of the
- 481 UK flora that MultiMOVE provides gives it a large number of potential applications.
- 482

483 <u>Access</u>

484

MultiMOVE was built using R 3.1.2 and depends on the following R packages: mgcv, fields,
leaps, earth, stringr, gsubfn, randomForest and nnet. In some circumstances, when not
automatically installed by MultiMOVE itself, these may need to be installed prior to installing
MultiMOVE. The package binary - MultiMOVE_2.0.1.zip – is available via eidc.ceh.ac.uk (doi:
<u>http://doi.org/10.5285/94ae1a5a-2a28-4315-8d4b-35ae964fc3b9</u>).

490

491 Supplementary Information

492

493 A manual for the latest version of the package (v2.0.1) is available in the supplementary

- 494 material. Please note that this latest version of the MultiMOVE package contains an
- additional two methods to describe a species' niche: Random Forests; and Neural Networks.
- 496

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498

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